## Lecture 11: Blocking and Confounding in $2^{k}$ design

Montgomery: Chapter 7

## Randomized Complete Block $2^{k}$ Design

- There are $n$ blocks
- Within each block, all treatments (levl combinations) are conducted.
- Run order in each block must be randomized
- Analysis follows general block factorial design
- When $k$ is large, cannot afford to conduct all the treatments within each block. Other blocking strategy should be considered.

Filtration Rate Experiment (revisited)

| factor |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| $A$ | $B$ | $C$ | $D$ | original response |
| - | - | - | - | 45 |
| + | - | - | - | 71 |
| - | + | - | - | 48 |
| + | + | - | - | 65 |
| - | - | + | - | 68 |
| + | - | + | - | 60 |
| - | + | + | - | 80 |
| + | + | + | - | 65 |
| - | - | - | + | 43 |
| + | - | - | + | 100 |
| - | + | - | + | 45 |
| + | + | - | + | 104 |
| - | - | + | + | 75 |
| + | - | + | + | 86 |
| - | + | + | + | 70 |
| + | + | + | + | 96 |

- Suppose there are two batches of raw material. Each batch can be used for only 8 runs. It is known these two batches are very different. Blocking should be employed to eliminate this variability.
- How to select 8 treatments (level combinations, or runs) for each block?


## $2^{2}$ Design with Two Blocks

Suppose there are two factors $(A, B)$ each with 2 levels, and two blocks $\left(b_{1}, b_{2}\right)$ each contiaining two runs (treatments). Since $b_{1}$ and $b_{2}$ are interchangeable, there are three possible blocking scheme:

|  |  |  | blocking scheme |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $A$ | $B$ | response | 1 | 2 | 3 |
| - | - | $y--$ | $b_{1}$ | $b_{1}$ | $b_{2}$ |
| + | - | $y_{+-}$ | $b_{1}$ | $b_{2}$ | $b_{1}$ |
| - | + | $y-+$ | $b_{2}$ | $b_{1}$ | $b_{1}$ |
| + | + | $y_{++}$ | $b_{2}$ | $b_{2}$ | $b_{2}$ |

Comparing blocking schemes:
Scheme 1:

- block effect: $b=\bar{y}_{b_{2}}-\bar{y}_{b_{1}}=\frac{1}{2}\left(-y_{--}-y_{+-}+y_{-+}+y_{++}\right)$
- main effect: $B=\frac{1}{2}\left(-y_{--}-y_{+-}+y_{-+}+y_{++}\right)$
- $B$ and $b$ are not distinguishable, or, confounded.


## Comparing Blocking Schemes (continued)

Scheme 2:

$$
\begin{gathered}
\text { block effect: } b=\bar{y}_{b_{2}}-\bar{y}_{b_{1}}=\frac{1}{2}\left(-y_{--}+y_{+-}-y_{-+}+y_{++}\right) \\
\text {main effect: } A=\frac{1}{2}\left(-y_{--}+y_{+-}-y_{-+}+y_{++}\right)
\end{gathered}
$$

$A$ and $b$ are not distinguishable, or confounded.

Scheme 3:

$$
\begin{gathered}
\text { block effect: } b=\bar{y}_{b_{2}}-\bar{y}_{b_{1}}=\frac{1}{2}\left(y_{--}-y_{+-}-y_{-+}+y_{++}\right) \\
\text {interaction: } A B=\frac{1}{2}\left(y_{--}-y_{+-}-y_{-+}+y_{++}\right)
\end{gathered}
$$

$A B$ and $b$ become indistinguishable, or confounded.

The reason for confounding: the block arrangement matches the contrast of some factorial effect.

Confounding makes the effect Inestimable.

Question: which scheme is the best (or causes the least damage)?

## $2^{k}$ Design with Two Blocks via Confounding

Confound blocks with the effect (contrast) of the highest order

Block 1 consists of all treatments with the contrast coefficient equal to -1 Block 2 consists of all treatments with the contrast coefficient equal to 1

Example 1. Block $2^{3}$ Design

| factorial effects (contrasts) |  |  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| । | A | B | C | AB | AC | BC | ABC |
| 1 | -1 | -1 | -1 | 1 | 1 | 1 | -1 |
| 1 | 1 | -1 | -1 | -1 | -1 | 1 | 1 |
| 1 | -1 | 1 | -1 | -1 | 1 | -1 | 1 |
| 1 | 1 | 1 | -1 | 1 | -1 | -1 | -1 |
| 1 | -1 | -1 | 1 | 1 | -1 | -1 | 1 |
| 1 | 1 | -1 | 1 | -1 | 1 | -1 | -1 |
| 1 | -1 | 1 | 1 | -1 | -1 | 1 | -1 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |

Defining relation: $b=A B C$ :
Block 1: ( --- ), (+ + -), ( +-+ ), ( -++ )
Block 2: (+ - -), ( -+- ), ( -++ ), $(+++)$

Example 2: For $2^{4}$ design with factors: $A, B, C, D$, the defining contrast
(optimal) for blocking factor (b) is

$$
b=A B C D
$$

In general, the optimal blocking scheme for $2^{k}$ design with two blocks is given by $b=A B \ldots K$, where $A, B, \ldots, K$ are the factors.

## Analyze Unreplicated Block $2^{k}$ Experiment

Filtration Experiment (four factors: $A, B, C, D$ ):

- Use defining relation: $b=A B C D$, i.e., if a treatment satisfies
$A B C D=-1$, it is allocated to block $1\left(b_{1}\right)$; if $A B C D=1$, it is allocated to block $2\left(b_{2}\right)$.
- (Assume that, all the observations in block 2 will be reduced by 20 because of the poor quality of the second batch of material, i.e. the true block effect=-20).

| factor |  |  |  |  | blocks |  |  |
| :--- | :--- | :--- | :--- | ---: | :--- | :---: | :---: |
| $A$ | $B$ | $C$ | $D$ | $b=A B C D$ | response |  |  |
| - | - | - | - | $1=b_{2}$ | $45-20=25$ |  |  |
| + | - | - | - | $-1=b_{1}$ | 71 |  |  |
| - | + | - | - | $-1=b_{1}$ | 48 |  |  |
| + | + | - | - | $1=b_{2}$ | $65-20=45$ |  |  |
| - | - | + | - | $-1=b_{1}$ | 68 |  |  |
| + | - | + | - | $1=b_{2}$ | $60-20=40$ |  |  |
| - | + | + | - | $1=b_{2}$ | $80-20=60$ |  |  |
| + | + | + | - | $-1=b_{1}$ | 65 |  |  |
| - | - | - | + | $-1=b_{1}$ | 43 |  |  |
| + | - | - | + | $1=b_{2}$ | $100-20=80$ |  |  |
| - | + | - | + | $1=b_{2}$ | $45-20=25$ |  |  |
| + | + | - | + | $-1=b_{1}$ | 104 |  |  |
| - | - | + | + | $1=b_{2}$ | $75-20=55$ |  |  |
| + | - | + | + | $-1=b_{1}$ | 86 |  |  |
| - | + | + | + | $-1=b_{1}$ | 70 |  |  |
| + | + | + | + | $1=b_{2}$ | $96-20=76$ |  |  |

## SAS File for Block Filtration Experiment

```
goption colors=(none);
data filter;
    do D = -1 to 1 by 2; do C = -1 to 1 by 2;
    do B = -1 to 1 by 2; do A = -1 to 1 by 2;
    input y @@; output;
    end; end; end; end;
cards;
25
;
data inter;
set filter; AB=A*B; AC=A*C; AD=A*D; BC=B*C; BD=B*D; CD=C*D; ABC=AB*C;
ABD=AB*D; ACD=AC*D; BCD=BC*D; block=ABC*D;
proc glm data=inter;
class A B C D AB AC AD BC BD CD ABC ABD ACD BCD block;
model y=block A B C D AB AC AD BC BD CD ABC ABD ACD BCD; run;
proc reg outest=effects data=inter;
```

```
model y=A B C D AB AC AD BC BD CD ABC ABD ACD BCD block;
data effect2; set effects; drop y intercept _RMSE_;
proc transpose data=effect2 out=effect3;
data effect4; set effect3; effect=col1*2;
proc sort data=effect4; by effect;
proc print data=effect4;
data effect5; set effect4; where __NAME__='block';
proc print data=effect5; run;
proc rank data=effect5 normal=blom;
var effect; ranks neff;
symbol1 v=circle;
proc gplot; plot effect*neff=_NAME_; run;
```


## SAS output: ANOVA Table

| Source | DF | Squares | Mean Square | F Value | Pr > F |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model | 15 | 7110.937500 | 474.062500 | - | - |
| Error |  | 0 | 0.000000 | - |  |
| Co Total | 15 | 7110.937500 |  |  |  |
| Source | DF | Type I SS | Mean Square | F Value | $\operatorname{Pr}>\mathrm{F}$ |
| block | 1 | 1387.562500 | 1387.562500 | - | - |
| A | 1 | 1870.562500 | 1870.562500 | - | - |
| B | 1 | 39.062500 | 39.062500 | - | - |
| C | 1 | 390.062500 | 390.062500 | - | . |
| D | 1 | 855.562500 | 855.562500 | - | - |
| AB | 1 | 0.062500 | 0.062500 | - | - |
| AC | 1 | 1314.062500 | 1314.062500 | - | - |
| AD | 1 | 1105.562500 | 1105.562500 | - | - |
| BC | 1 | 22.562500 | 22.562500 | - | - |
| BD | 1 | 0.562500 | 0.562500 | - | - |
| CD | 1 | 5.062500 | 5.062500 | - | - |
| ABC | 1 | 14.062500 | 14.062500 | . | - |


| ABD | 1 | 68.062500 | 68.062500 |
| :--- | :--- | :--- | :--- |
| ACD | 1 | 10.562500 | 10.562500 |
| BCD | 1 | 27.562500 | 27.562500 |

proportion of variance explained by blocks

$$
\frac{1387.5625}{7110.9375}=19.5 \%
$$

Similarly proportion of variance can be calculated for other effects.

## SAS output: factorial effects and block effect

| Obs | _NAME_ | COL1 | effect |
| ---: | :--- | ---: | ---: |
|  |  |  |  |
| 1 | block | -9.3125 | -18.625 |
| 2 | AC | -9.0625 | -18.125 |
| 3 | BCD | -1.3125 | -2.625 |
| 4 | ACD | -0.8125 | -1.625 |
| 5 | CD | -0.5625 | -1.125 |
| 6 | BD | -0.1875 | -0.375 |
| 7 | AB | 0.0625 | 0.125 |
| 8 | ABC | 0.9375 | 1.875 |
| 9 | BC | 1.1875 | 2.375 |
| 10 | B | 1.5625 | 3.125 |
| 11 | ABD | 2.0625 | 4.125 |
| 12 | C | 4.9375 | 9.875 |
| 13 | D | 7.3125 | 14.625 |
| 14 | AD | 8.3125 | 16.625 |
| 15 | A | 10.8125 | 21.625 |

Factorial effects are exactly the same as those from the original data (why?)
blocking effect: $-18.625=\bar{y}_{b_{2}}-\bar{y}_{b_{1}}$, is in fact
-20 (true blocking effect) +1.375 (some interaction of $A B C)$

This is caused by confounding between $b$ and $A B C$.

## SAS output: QQ plot without Blocking Effect


significant effects are:

$$
A, C, D, A C, A D
$$

## $2^{k}$ Design with Four Blocks

Need two 2-level blocking factors to generate 4 different blocks.
Confound each blocking factors with a high order factorial effect.
The interaction between these two blocking factors matters.
The interaction will be confounded with another factorial effect.

Optimal blocking scheme has least confounding severity.
$2^{4}$ design with four blocks: factors are $A, B, C, D$ and the blocking factors are $b 1$ and $b 2$

possible blocking schemes:
Scheme 1:
defining relations: $b 1=A B C, b 2=A C D$; induce confounding

$$
b 1 b 2=A B C * A C D=A^{2} B C^{2} D=B D
$$

Scheme 2:
Defining relations: $b 1=A B C D, b 2=A B C$, induce confounding

$$
b 1 b 2=A B C D * A B C=D
$$

Which is better?

## $2^{k}$ Design with $2^{p}$ Blocks

- $k$ factors: $A, B, \ldots K$, and $p$ is usually much less than $k$.
- $p$ blocking factors: $b 1, b 2, \ldots b p$ with levels -1 and 1
- confound blocking factors with $k$ chosen high-order factorial effects, i.e., $b 1=$ effect1, $b 2=$ effect 2 , etc. ( $p$ defining relations)
- These $p$ defining relations induce another $2^{p}-p-1$ confoundings.
- treatment combinations with the same values of $b 1, \ldots b p$ are allocated to the same block. Within each block.
- each block consists of $2^{k-p}$ treatment combinations (runs)
- Given $k$ and $p$, optimal schemes are tabulated, e.g., Montgomery Table 7.8, or Wu\&Hamada Appendix 3A

