# Lecture 12: Blocking and Confounding in $2^k$ design

Montgomery: Chapter 7

# Randomized Complete Block $2^k$ Design

- There are *n* blocks
- Within each block, all treatments (level combinations) are conducted.
- Run order in each block must be randomized
- Analysis follows general block factorial design
- When k is large, cannot afford to conduct all the treatments within each block.
   Other blocking strategy should be considered.

	fac	tor		
A	B	C	D	original response
_	_	_	_	45
+	_	_	_	71
_	+	_	_	48
+	+	_	_	65
_	—	+	_	68
+	_	+	_	60
—	+	+	—	80
+	+	+	—	65
—	—	—	+	43
+	—	—	+	100
_	+	—	+	45
+	+	—	+	104
—	—	+	+	75
+	—	+	+	86
_	+	+	+	70
+	+	+	+	96

### Filtration Rate Experiment (revisited)

- Suppose there are two batches of raw material. Each batch can be used for only 8 runs. It is known these two batches are very different. Blocking should be employed to eliminate this variability.
- How to select 8 treatments (level combinations, or runs) for each block?

# $2^2\ {\rm Design}$ with Two Blocks

Suppose there are two factors (A, B) each with 2 levels, and two blocks  $(b_1, b_2)$  each containing two runs (treatments). Since  $b_1$  and  $b_2$  are interchangeable, there are three possible blocking scheme:

			blocking scheme				
A	В	response	1	2	3		
_	_	$y_{}$	$b_1$	$b_1$	$b_2$		
+	_	$y_{+-}$	$b_1$	$b_2$	$b_1$		
_	+	$y_{-+}$	$b_2$	$b_1$	$b_1$		
+	+	$y_{++}$	$b_2$	$b_2$	$b_2$		

Comparing blocking schemes:

Scheme 1:

- block effect:  $b = \bar{y}_{b_2} \bar{y}_{b_1} = \frac{1}{2}(-y_{--} y_{+-} + y_{-+} + y_{++})$
- main effect:  $B = \frac{1}{2}(-y_{--} y_{+-} + y_{-+} + y_{++})$
- B and b are not distinguishable, or, confounded.

### **Comparing Blocking Schemes (continued)**

Scheme 2:

block effect: 
$$b = \bar{y}_{b_2} - \bar{y}_{b_1} = \frac{1}{2}(-y_{--} + y_{+-} - y_{-+} + y_{++})$$

main effect: 
$$A = \frac{1}{2}(-y_{--} + y_{+-} - y_{-+} + y_{++})$$

 $\boldsymbol{A}$  and  $\boldsymbol{b}$  are not distinguishable, or confounded.

Scheme 3:

block effect: 
$$b = \bar{y}_{b_2} - \bar{y}_{b_1} = \frac{1}{2}(y_{--} - y_{+-} - y_{-+} + y_{++})$$

interaction: 
$$AB = \frac{1}{2}(y_{--} - y_{+-} - y_{-+} + y_{++})$$

AB and b become indistinguishable, or confounded.

The reason for confounding: the block arrangement matches the contrast of some factorial effect.

Confounding makes the effect **Inestimable**.

Question: which scheme is the best (or causes the least damage)?

# $2^k$ Design with Two Blocks via Confounding

Confound blocks with the effect (contrast) of the highest order

Block 1 consists of all treatments with the contrast coefficient equal to -1 Block 2 consists of all treatments with the contrast coefficient equal to 1

Example 1. Block  $2^3$  Design

	factorial effects (contrasts)												
I	А	B C AB AC BC A											
1	-1	-1	-1	1	1	1	-1						
1	1	-1	-1	-1	-1	1	1						
1	-1	1	-1	-1	1	-1	1						
1	1	1	-1	1	-1	-1	-1						
1	-1	-1	1	1	-1	-1	1						
1	1	-1	1	-1	1	-1	-1						
1	-1	1	1	-1	-1	1	-1						
1	1	1	1	1	1	1	1						

Defining relation: b = ABC:

Block 1: (---), (++-), (+-+), (-++)Block 2: (+--), (-+-), (-++), (+++)

Example 2: For  $2^4$  design with factors: A, B, C, D, the defining contrast

(optimal) for blocking factor (b) is

b = ABCD

In general, the optimal blocking scheme for  $2^k$  design with two blocks is given by  $b = AB \dots K$ , where  $A, B, \dots, K$  are the factors.

## Analyze Unreplicated Block $2^k$ Experiment

Filtration Experiment (four factors: A, B, C, D):

- Use defining relation: b = ABCD, i.e., if a treatment satisfies ABCD = -1, it is allocated to block 1( $b_1$ ); if ABCD = 1, it is allocated to block 2 ( $b_2$ ).
- (Assume that, all the observations in block 2 will be reduced by 20 because of the poor quality of the second batch of material, i.e. the true block effect=-20).

	fac	ctor		blocks	
A	B	C	D	b = ABCD	response
_	_	_	_	1=b <sub>2</sub>	45-20=25
+	_	_	_	-1=b <sub>1</sub>	71
—	+	_	_	-1=b <sub>1</sub>	48
+	+	—	_	1=b <sub>2</sub>	65-20=45
—	—	+	—	-1=b <sub>1</sub>	68
+	—	+	—	1=b <sub>2</sub>	60-20=40
_	+	+	—	1=b <sub>2</sub>	80-20=60
+	+	+	—	-1=b <sub>1</sub>	65
_	—	—	+	-1=b <sub>1</sub>	43
+	—	—	+	1=b <sub>2</sub>	100-20=80
_	+	—	+	1=b <sub>2</sub>	45-20=25
+	+	—	+	-1=b <sub>1</sub>	104
—	—	+	+	1=b <sub>2</sub>	75-20=55
+	—	+	+	-1=b <sub>1</sub>	86
_	+	+	+	-1=b <sub>1</sub>	70
+	+	+	+	1=b <sub>2</sub>	96-20=76

#### **SAS File for Block Filtration Experiment**

```
goption colors=(none);
data filter;
 do D = -1 to 1 by 2;do C = -1 to 1 by 2;
 do B = -1 to 1 by 2; do A = -1 to 1 by 2;
 input y @@; output;
 end; end; end; end;
cards;
25 71 48 45 68 40 60 65 43 80 25 104 55 86 70 76
;
data inter;
set filter; AB=A*B; AC=A*C; AD=A*D; BC=B*C; BD=B*D; CD=C*D; ABC=AB*C;
ABD=AB*D; ACD=AC*D; BCD=BC*D; block=ABC*D;
proc glm data=inter;
```

class A B C D AB AC AD BC BD CD ABC ABD ACD BCD block; model y=block A B C D AB AC AD BC BD CD ABC ABD ACD BCD; run;

proc reg outest=effects data=inter;

```
model y=A B C D AB AC AD BC BD CD ABC ABD ACD BCD block;
data effect2; set effects; drop y intercept _RMSE_;
proc transpose data=effect2 out=effect3;
data effect4; set effect3; effect=col1*2;
proc sort data=effect4; by effect;
proc print data=effect4;
```

data effect5; set effect4; where \_\_NAME\_^='block';
proc print data=effect5; run;

```
proc rank data=effect5 normal=blom;
var effect; ranks neff;
```

```
symbol1 v=circle;
proc gplot; plot effect*neff=_NAME_; run;
```

## SAS output: ANOVA Table

Source Model Error	DF 15	Squares 7110.937500 0	Mean Square 474.062500 0.000000	F Value	Pr > F
Co Total	15	7110.937500			
Source	DF	Type I SS	Mean Square	F Value	Pr > F
block	1	1387.562500	1387.562500		•
A	1	1870.562500	1870.562500	•	•
В	1	39.062500	39.062500	•	•
С	1	390.062500	390.062500	•	•
D	1	855.562500	855.562500	•	•
AB	1	0.062500	0.062500	•	•
AC	1	1314.062500	1314.062500	•	•
AD	1	1105.562500	1105.562500	•	•
BC	1	22.562500	22.562500	•	
BD	1	0.562500	0.562500	•	•
CD	1	5.062500	5.062500		•
ABC	1	14.062500	14.062500	•	•

ABD	1	68.062500	68.062500
ACD	1	10.562500	10.562500
BCD	1	27.562500	27.562500

proportion of variance explained by blocks

$$\frac{1387.5625}{7110.9375} = 19.5\%$$

Similarly proportion of variance can be calculated for other effects.

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### SAS output: factorial effects and block effect

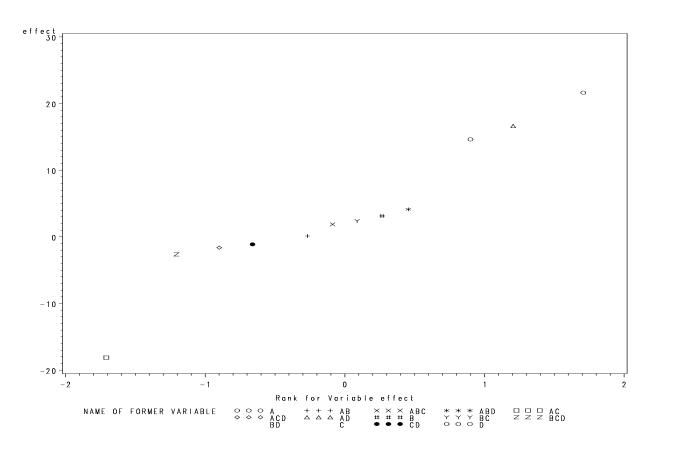
Obs	_NAME_	COL1	effect
1	block	-9.3125	-18.625
2	AC	-9.0625	-18.125
3	BCD	-1.3125	-2.625
4	ACD	-0.8125	-1.625
5	CD	-0.5625	-1.125
б	BD	-0.1875	-0.375
7	AB	0.0625	0.125
8	ABC	0.9375	1.875
9	BC	1.1875	2.375
10	В	1.5625	3.125
11	ABD	2.0625	4.125
12	C	4.9375	9.875
13	D	7.3125	14.625
14	AD	8.3125	16.625
15	A	10.8125	21.625

Factorial effects are exactly the same as those from the original data (why?)

blocking effect: -18.625= $\bar{y}_{b_2}$  –  $\bar{y}_{b_1}$ , is in fact

-20(true blocking effect) + 1.375(some interaction of ABC)

This is caused by confounding between b and ABC.



### SAS output: QQ plot without Blocking Effect

significant effects are:

A, C, D, AC, AD

## $2^k$ Design with Four Blocks

Need two 2-level blocking factors to generate 4 different blocks. Confound each blocking factors with a high order factorial effect. The interaction between these two blocking factors matters. The interaction will be confounded with another factorial effect.

Optimal blocking scheme has least confounding severity.

 $2^4$  design with four blocks: factors are A, B, C, D and the blocking factors are b1 and b2

А	В	С	D	AB	AC	 'D	ABC	ABD	ACD	BCD	ABCD			
-1	-1	-1	-1	1	1	1	-1	-1	-1	-1	1			
1	-1	-1	-1	-1	-1	1	1	1	1	-1	-1	b1	b2	blocks
-1	1	-1	-1	-1	1	1	1	1	-1	1	-1	-1	-1	1
1	1	-1	-1	1	-1	1	-1	-1	1	1	1	1	-1	2
												-1	1	3
• • •		• • •	• • •	• • •	• • • •	 • • •						1	1	4
-1	-1	1	1	1	-1	1	1	1	-1	-1	1			
1	1	1	1	-1	1	1	-1	-1	1	-1	-1			
-1	-1	1	1	-1	-1	1	-1	-1	-1	1	-1			
1	1	1	1	1	1	1	1	1	1	1	1			

possible blocking schemes:

Scheme 1:

defining relations: b1 = ABC, b2 = ACD; induce confounding

$$b1b2 = ABC * ACD = A^2BC^2D = BD$$

Scheme 2:

Defining relations: b1 = ABCD, b2 = ABC, induce confounding

$$b1b2 = ABCD * ABC = D$$

Which is better?

# $2^k \ \mathrm{Design} \ \mathrm{with} \ 2^p \ \mathrm{Blocks}$

- k factors:  $A, B, \dots K$ , and p is usually much less than k.
- p blocking factors: b1, b2,...bp with levels -1 and 1
- confound blocking factors with k chosen high-order factorial effects, i.e., b1=effect1, b2=effect2, etc.(p defining relations)
- These p defining relations induce another  $2^p p 1$  confounding.
- treatment combinations with the same values of b1,...bp are allocated to the same block. Within each block.
- each block consists of  $2^{k-p}$  treatment combinations (runs)
- Given k and p, optimal schemes are tabulated, e.g., Montgomery Table 7.8, or Wu&Hamada Appendix 3A