

Exponential Families.

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- Statistics: data $\rightarrow$ knowledge
- exponential families: classes of distributions
- $h: \mathbb{R}^n \rightarrow \mathbb{R}$ nonnegative function
- $T_1 \dots T_s$ — measurable functions: $\mathbb{R}^n \rightarrow \mathbb{R}$

$$A(\eta) = \log \int e^{\sum_{i=1}^s \eta_i T_i(x)} h(x) d\mu(x) \quad \begin{array}{l} \mu - \text{measure on } \mathbb{R}^n \\ \eta = (\eta_1, \dots, \eta_s) \in \mathbb{R}^s \end{array}$$

$$p_\eta(x) = e^{\sum_{i=1}^s \eta_i T_i(x) - A(\eta)} \cdot h(x) \quad x \in \mathbb{R}^n$$

$$\int p_\eta(x) d\mu = 1 \quad \rightarrow \quad \text{a family of probability densities indexed by } \eta$$

$$\Xi = \{ \eta : A(\eta) < \infty \} \quad \text{—— natural parameter space}$$

$$\{ p_\eta : \eta \in \Xi \} \quad \text{—— } s\text{-parameter exponential family in canonical form.}$$

ex. exponential density: $p_\eta(x) = \frac{1}{-\eta} e^{+\eta x} \quad \eta < 0 \quad x > 0$

$$p_\eta(x) = e^{+\eta x + \log(-\eta)} \quad h = 1_{(0, \infty)} \quad s = 1$$

$$T_1(x) = x \quad A(\eta) = \log \int e^{+\eta x} dx = \begin{cases} \log(-\frac{1}{\eta}) & \eta < 0 \\ \infty & \eta \geq 0 \end{cases}$$

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- other parameterization.

$$p_{\theta}(x) = e^{\sum_{i=1}^s \eta_i(\theta) T_i(x) - B(\theta)} \cdot h(x) \quad \begin{array}{l} \theta \in \Omega. \\ x \in \mathbb{R}^n. \end{array}$$

$$B(\theta) = A(\eta(\theta)).$$

$\{p_{\theta} : \theta \in \Omega\}$ — s-parameter exponential families.

ex: $N(\mu, \sigma^2)$.

$$p_{\theta}(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} = \frac{1}{\sqrt{2\pi}} e^{\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \left(\frac{\mu^2}{2\sigma^2} + \frac{1}{2}\log\sigma\right)}$$

$\theta = (\mu, \sigma^2)$. Two-parameter exponential family.

$$T_1(x) = x, \quad T_2(x) = x^2, \quad \eta_1(\theta) = \frac{\mu}{\sigma^2}, \quad \eta_2(\theta) = -\frac{1}{2\sigma^2}, \quad B(\theta) = \frac{\mu^2}{2\sigma^2} + \frac{1}{2}\log\sigma.$$

$$h(x) = \frac{1}{\sqrt{2\pi}}.$$

ex. $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$.

$$p(x_1, \dots, x_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i-\mu)^2}{2\sigma^2}} = \frac{1}{(2\pi)^{n/2}} e^{\frac{\mu}{\sigma^2} \sum_{i=1}^n x_i - \frac{1}{2\sigma^2} \sum_{i=1}^n x_i^2 - n\left(\frac{\mu^2}{2\sigma^2} + \frac{1}{2}\log\sigma\right)}$$

$$T_1(x) = \sum_{i=1}^n x_i, \quad T_2(x) = \sum_{i=1}^n x_i^2, \quad \eta_1(\theta) = \frac{\mu}{\sigma^2}, \quad \eta_2(\theta) = -\frac{1}{2\sigma^2}$$

$$B(\theta) = n\left[\frac{\mu^2}{2\sigma^2} + \frac{1}{2}\log\sigma\right], \quad h(x) = \frac{1}{(2\pi)^{n/2}}.$$

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$$X_1 \dots X_n \stackrel{\text{iid}}{\sim} \exp\left[\sum_{i=1}^s \eta_i(\theta) T_i(x) - B(\theta)\right] h(x)$$

$$\text{joint} \quad \exp\left[\sum_{i=1}^s \eta_i(\theta) \left(\sum_{j=1}^n T_i(x_j)\right) - n B(\theta)\right] \prod_{j=1}^n h(x_j)$$

$$s\text{-para. exp. family.} \quad \tilde{T}_i(x) = \sum_{j=1}^n T_i(x_j) \quad \tilde{B}(\theta) = n B(\theta) \\ \tilde{h}(x) = \prod_{j=1}^n h(x_j).$$

- Relate moments and cumulants for the statistic $T_1, \dots, T_s$ to derivatives of A .

$$\text{Thm.} \quad \text{if } \int |x|^k e^{\sum \eta_i T_i(x)} h(x) d\mu(x) < \infty$$

$$\Rightarrow g(\eta) = \int f(x) e^{\sum \eta_i T_i(x)} h(x) d\mu(x) \text{ is continuous and}$$

has continuous partial derivatives of all orders of η .

these derivatives can be computed by differentiation under the integral sign.

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$$g(\eta) = e^{A(\eta)} = \int e^{\sum_i \eta_i T_i(x)} h(x) d\mu(x).$$

$$e^{A(\eta)} \frac{\partial A(\eta)}{\partial \eta_j} = \int \frac{\partial}{\partial \eta_j} e^{\sum_i \eta_i T_i(x)} h(x) d\mu(x) = \int T_j(x) e^{\sum_i \eta_i T_i(x)} h(x) d\mu(x)$$

$$\frac{\partial A(\eta)}{\partial \eta_j} = \int T_j(x) P_\eta(x) d\mu(x) \Rightarrow E_\eta T_j(x) = \frac{\partial A(\eta)}{\partial \eta_j}.$$

$T = (T_1 \dots T_s)' \in \mathbb{R}^s$. a random vector.

- mgf $M_T(u) = E e^{u_1 T_1 + \dots + u_s T_s} = E e^{u^* T} \quad u \in \mathbb{R}^s.$

- cumulant generating function : $K_T(u) = \log M_T(u)$

lem 2.7. If $M_X(u)$ and $M_Y(u)$ are finite and agree for u in some set with a nonempty interior, then X and Y have the same distribution. $P_X = P_Y$.

- moments: $a_{r_1 \dots r_s} = E[T_1^{r_1} \dots T_s^{r_s}]$

Thm. $a_{r_1 \dots r_s} = \frac{\partial^{r_1}}{\partial u_1^{r_1}} \dots \frac{\partial^{r_s}}{\partial u_s^{r_s}} M_T(u) \Big|_{u=0}.$

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$$- K_{r_1 \dots r_s} = \frac{\partial^{r_1}}{\partial u_{r_1}^{r_1}} \dots \frac{\partial^{r_s}}{\partial u_{r_s}^{r_s}} K_T(u) \Big|_{u=0} \quad \text{Cumulants.}$$

$$s=1 \quad K_T' = \frac{M_T'}{M_T} \quad K_T'' = \frac{M_T M_T'' - M_T'^2}{M_T^2}$$

$$K_1 = ET \quad K_2 = ET^2 - (ET)^2 = \text{Var}(T)$$

$$\text{lem.} \quad X \perp Y \Rightarrow E(XY) = EX EY$$

$$\nabla T = Y_1 + \dots + Y_n \quad Y_1 \dots Y_n \text{ independent.}$$

$$M_T(u) = E e^{u'T} = M_{Y_1}(u) \times \dots \times M_{Y_n}(u)$$

$$K_T(u) = K_{Y_1}(u) + \dots + K_{Y_n}(u).$$

$$- X \sim e^{\eta T(x) - A(\eta)} h(x) \Rightarrow T \text{ has msf}$$

$$\begin{aligned} E_{\eta} e^{u \cdot T(x)} &= \int e^{u T(x)} e^{\eta T(x) - A(\eta)} h(x) d\mu(x) \\ &= e^{A(u+\eta) - A(\eta)} \int e^{(u+\eta) T(x) - A(u+\eta)} h(x) d\mu(x) \\ &= e^{A(u+\eta) - A(\eta)} \end{aligned}$$

$$K_T(u) = A(u+\eta) - A(\eta).$$

$$K_{r_1 \dots r_s} = \frac{\partial^{r_1}}{\partial \eta_{r_1}^{r_1}} \dots \frac{\partial^{r_s}}{\partial \eta_{r_s}^{r_s}} A(\eta)$$

ex. $X \sim \text{Poisson}(\lambda)$. $P(X=x) = \frac{\lambda^x e^{-\lambda}}{x!} = \frac{1}{x!} e^{x \log \lambda - \lambda}$ (6)
 $x=0,1,\dots$

not in canonical form. The canonical para. is $\eta = \log \lambda$.

$$P(X=x) = \frac{1}{x!} e^{x\eta - e^\eta} \quad x=0,1,\dots \quad A(\eta) = e^\eta$$

all of the cumulants of $T=X$ are $e^\eta = \lambda$.

ex. $N(\mu, \sigma^2)$. σ^2 is known.

$$\phi_{\mu}(x) = e^{\frac{\mu x}{\sigma^2} - \frac{\mu^2}{2\sigma^2}} \frac{e^{-\frac{x^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$

$$T(x) = x. \quad \eta = \frac{\mu}{\sigma^2}. \quad A(\eta) = \frac{\sigma^2 \eta^2}{2}.$$

$$K_1 = A'(\eta) = \sigma^2 \eta = \mu. \quad K_2 = A''(\eta) = \sigma^2. \quad \text{higher-order cumulants } K_3, K_4, \dots \text{ are zero}$$

- Calculate moments from cumulants, when $s=1$.

$$M = e^K. \quad M' = K' e^K. \quad M'' = (K'' + K'^2) e^K.$$

$$M''' = (K''' + 3K'K'' + K'^3) e^K. \quad M'''' = (K'''' + 3K''^2 + 4K'K''' + 6K'^2K'' + K'^4) e^K$$

At zero.

$$ET = K_1 \quad ET^2 = K_2 + K_1^2.$$

$$ET^3 = K_3 + 3K_1K_2 + K_1^3. \quad ET^4 = K_4 + 3K_2^2 + 4K_1K_3 + 6K_1^2K_2 + K_1^4.$$

ex. $X \sim \text{Pois}(\lambda)$. $EX = \lambda$. $EX^2 = \lambda + \lambda^2$. $EX^3 = \lambda + 3\lambda^2 + \lambda^3$.

$$EX^4 = \lambda + 7\lambda^2 + 6\lambda^3 + \lambda^4.$$

$$X \sim N(\mu, \sigma^2). \quad EX^3 = 3\mu\sigma^2 + \mu^3. \quad EX^4 = 3\sigma^4 + 4\mu^2\sigma^2 + \mu^4.$$