

Bayesian Estimation

①

- Bayesian approach: $\int R(\theta, s) d\Lambda(\theta)$.

Λ — a specific probability measure on Ω .

- $\theta \sim \Lambda$ prior distribution

$X | \theta = \theta \sim P_\theta$.

$$E[L(\theta, s(X)) | \theta = \theta] = \int L(\theta, s(x)) dP_\theta(x) = R(\theta, s)$$

$$E L(\theta, s) = E E[L(\theta, s) | \theta] = E R(\theta, s) = \int R(\theta, s) d\Lambda(\theta)$$

- The choice of the prior distribution Λ is critical.
subjective opinion.

- minimize $\int R(\theta, s) d\Lambda(\theta) \rightarrow$ called Bayes.

Thm: Suppose $\theta \sim \Lambda$. $X | \theta = \theta \sim P_\theta$ $L(\theta, d) \geq 0$. If

a) $E L(\theta, s_0) < \infty$ for some s_0

b) for a.e. x . $\exists \delta_\Lambda(x)$ minimize $E[L(\theta, d) | X=x]$
w.r.t. d .

$\Rightarrow \delta_\Lambda$ is a Bayes estimator.

$$E(L(\theta, s(x)) | X) \geq E(L(\theta, \delta_\Lambda(x)) | X)$$

ex. Weighted L_2 -loss.

$$L(\theta, d) = w(\theta)(d - g(\theta))^2.$$

minimize:

$$\begin{aligned} & E[\omega(\theta)(d - g(\theta))^2 | X=x] \\ &= d^2 E[\omega(\theta) | X=x] - 2d E[\omega(\theta)g(\theta) | X=x] \\ & \quad + E[\omega(\theta)g^2(\theta) | X=x] \end{aligned}$$

$$\delta_\Lambda(x) = \frac{E[\omega(\theta)g(\theta) | X=x]}{E[\omega(\theta) | X=x]}.$$

if $w=1$. $\delta_\Lambda(x) = E[g(\theta) | x]$. — posterior mean of $g(\theta)$.

Λ is absolutely continuous with Lebesgue density λ .

joint (X, θ) : $p_\theta(x) \lambda(\theta)$

marginal X : $g(x) = \int p_\theta(x) \lambda(\theta) d\theta$

$\theta | X=x$: $\lambda(\theta | x) = \frac{p_\theta(x) \lambda(\theta)}{g(x)}$

$$\delta_\Lambda(x) = \frac{\int \omega(\theta) g(\theta) p_\theta(x) \lambda(\theta) d\theta}{\int \omega(\theta) p_\theta(x) \lambda(\theta) d\theta}$$

Ex. Binomial.

$$X \sim \text{Bin}(n, \theta)$$

prior. $\theta \sim \text{Beta}(\alpha, \beta)$.

$$\lambda(\theta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1} \quad 0 < \theta < 1.$$

$$\int_0^1 \theta^{\alpha-1} (1-\theta)^{\beta-1} d\theta = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$E\theta = \frac{\alpha}{\alpha+\beta}$$

$$\begin{aligned} q(x) &= \int p_0(x) \lambda(\theta) d\theta = \int_0^1 \binom{n}{x} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1} d\theta \\ &= \binom{n}{x} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{\Gamma(x+\alpha)\Gamma(n-x+\beta)}{\Gamma(n+\alpha+\beta)} \end{aligned}$$

$$\lambda(\theta|x) = \frac{\Gamma(n+\alpha+\beta)}{\Gamma(\alpha+x)\Gamma(\beta+n-x)} \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1} \quad \theta \in (0,1)$$

$$\theta|x=x \sim \text{Beta}(x+\alpha, n-x+\beta)$$

$$f_n(x) = E(\theta|x) = \frac{x+\alpha}{n+\alpha+\beta} = \frac{n}{n+\alpha+\beta} \frac{x}{n} + \left[1 - \frac{n}{n+\alpha+\beta}\right] \frac{\alpha}{\alpha+\beta}.$$

Prior distributions that ensures a posterior from the same class are called conjugate.

ex. Negative Binomial.

X — number of failures before the second success.

$$P_{\theta}(x) = P_{\theta}(X=x) = (x+1)\theta^2(1-\theta)^x. \quad x=0,1,\dots$$

$$g(\theta) = \frac{1}{\theta}. \quad \theta \sim \text{unif}(0,1).$$

$$\lambda(\theta|x) \propto P_{\theta}(x)\lambda(\theta) \propto \theta^2(1-\theta)^x.$$

$$\theta|x=x \sim \text{Beta}(3, x+1)$$

$$\begin{aligned} \delta_0(x) &= E(\theta^{-1}|X=x) = \frac{\Gamma(x+4)}{\Gamma(3)\Gamma(x+1)} \int_0^1 \theta(1-\theta)^x d\theta \\ &= \frac{\Gamma(x+4)\Gamma(2)\Gamma(x+1)}{\Gamma(3)\Gamma(x+1)\Gamma(x+3)} = \frac{x+3}{2}. \end{aligned}$$

ex 5.3. UMVUE of $\frac{1}{\theta}$: $\delta_1(x) = \frac{x+2}{2}$.

$$\delta_0(x) = \delta_1(x) + \frac{1}{2}.$$

$$\text{bias: } b(\theta, \delta_0) = E_{\theta} \delta_0(x) - \frac{1}{\theta} = \frac{1}{2}.$$

$$R(\theta, \delta_0) = \text{Var}(\delta_0) + \frac{1}{4} = \text{Var}_{\theta}(\delta_1) + \frac{1}{4} = R(\theta, \delta_1) + \frac{1}{4}$$

UMVUE δ_1 has uniformly smaller risk than δ_0 !

An estimator is call inadmissible if a competing estimator has a better risk function.

Condition (a) of thm fails!