

posterior distribution under independent samples. STAT 417. Lecture Note 14

Eg 7. If $x_1, x_2, \dots, x_n$ are iid samples following the Bernoulli models:

$$P(x_i = 1) = 1 - P(x_i = 0) = \theta.$$

Assume that θ has the following prior distribution:

$$f(\theta) = \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}, \quad 0 \leq \theta \leq 1. \quad (\text{Beta distribution}).$$

Then the posterior distribution of θ is

$$f(\theta | x_1, \dots, x_n) \propto f_0(x_1, \dots, x_n) f(\theta)$$

$$= f_{\theta}(x_1) \dots f_{\theta}(x_n) f(\theta)$$

$$= \theta^{x_1} (1-\theta)^{1-x_1} \theta^{x_2} (1-\theta)^{1-x_2} \dots \theta^{x_n} (1-\theta)^{1-x_n} f(\theta)$$

$$= \theta^{n\bar{x}} (1-\theta)^{n(1-\bar{x})} \cdot \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}.$$

$$= \frac{1}{B(\alpha, \beta)} \theta^{n\bar{x} + \alpha - 1} (1-\theta)^{n(1-\bar{x}) + \beta - 1}.$$

so the posterior distribution of θ is $\text{Beta}(n\bar{x} + \alpha, n(1-\bar{x}) + \beta)$.

Ex 8. Let $X_1, \dots, X_n$ be iid samples following $N(\mu, \sigma_0^2)$.

where σ_0^2 is known. $\mu \in \mathbb{R}$ is unknown. The prior for μ is

$\mu \sim N(\mu_0, \tau_0^2)$, where both μ_0 and τ_0^2 are known.

Then the

$$f(x_1, x_2, \dots, x_n) \propto f(x_1, \dots, x_n | \mu) f(\mu)$$

$$= \left(\frac{1}{\sqrt{2\pi} \sigma_0} \right)^n e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma_0^2}} \cdot \frac{1}{\sqrt{2\pi} \tau_0} e^{-\frac{(\mu - \mu_0)^2}{2\tau_0^2}}$$

$$= \exp\left(-\frac{(n\mu^2 - 2n\bar{x}\mu)\tau_0^2 + \mu^2\sigma_0^2 - 2\mu_0\sigma_0^2\mu}{2\sigma_0^2\tau_0^2}\right) \cdot \left(\frac{1}{\sqrt{2\pi}\sigma_0}\right)^n \cdot \frac{1}{\sqrt{2\pi}\tau_0} \exp\left(-\frac{\sum x_i^2}{2\sigma_0^2}\right) \cdot \exp\left(-\frac{\mu_0^2}{2\tau_0^2}\right)$$

$$= \exp\left(-\frac{(n\tau_0^2 + \sigma_0^2)\mu^2 - 2(n\bar{x}\tau_0^2 + \mu_0\sigma_0^2)\mu}{2\sigma_0^2\tau_0^2}\right) \cdot \left(\frac{1}{\sqrt{2\pi}\sigma_0}\right)^n \frac{1}{\sqrt{2\pi}\tau_0} \exp\left(-\frac{\sum x_i^2}{2\sigma_0^2}\right) \cdot \exp\left(-\frac{\mu_0^2}{2\tau_0^2}\right)$$

$$= \exp\left(-\frac{(n\tau_0^2 + \sigma_0^2)\left(\mu - \frac{n\bar{x}\tau_0^2 + \mu_0\sigma_0^2}{n\tau_0^2 + \sigma_0^2}\right)^2}{2\sigma_0^2\tau_0^2}\right)$$

$$\cdot \left(\frac{1}{\sqrt{2\pi}\sigma_0}\right)^n \frac{1}{\sqrt{2\pi}\tau_0} \exp\left(-\frac{\sum x_i^2}{2\sigma_0^2}\right) \cdot \exp\left(-\frac{\mu_0^2}{2\tau_0^2}\right) \cdot \exp\left(-\frac{(n\bar{x}\tau_0^2 + \mu_0\sigma_0^2)}{2\sigma_0^2\tau_0^2(n\tau_0^2 + \sigma_0^2)}\right)$$

So μ has posterior distribution $N\left(\frac{n\bar{x}\tau_0^2 + \mu_0\sigma_0^2}{n\tau_0^2 + \sigma_0^2}, \frac{\sigma_0^2\tau_0^2}{n\tau_0^2 + \sigma_0^2}\right)$

an exercise: If $\bar{x} = 0$, $n = 100$, ~~$\tau_0^2 = 1$~~ , $\mu_0 = 0$, $\sigma_0^2 = 200$, $\tau_0^2 = 2$.

then the posterior distribution of μ is $N(0, 1)$. find ~~$P(\mu < 0.1 | S)$~~ mean and variance of μ given S .

Answer: 0 and 1, find $P(\mu < 1 | S) = 0.8413$.

EX1
~~99~~ Suppose $X_1, X_2, \dots, X_n$ follow Exponential(θ).

while, θ has prior Exponential(θ_0). $\theta_0 > 0$ is known.

Find the posterior distribution of θ .

Sol: $f(\theta | x_1, \dots, x_n) \propto f_\theta(x_1, \dots, x_n) f(\theta)$

$$\propto \theta e^{-\theta x_1} \dots \theta e^{-\theta x_n} \cdot \theta_0 e^{-\theta_0 \theta}$$

$$\propto \theta_0 \cdot \theta^n e^{-(n\bar{x} + \theta_0)\theta}$$

so the posterior of θ is $f(\theta | x_1, \dots, x_n)$ as above.
 ~~$\theta \sim \text{Exponential}$~~

clearly, $\theta | x_1, \dots, x_n \sim \text{Gamma}(n+1, n\bar{x} + \theta_0)$.

posterior mean: and variance:

suppose θ given the sample s has posterior distribution $f(\theta|s)$.

then the posterior mean of θ is

$$E(\theta|s) = \int \theta f(\theta|s) d\theta.$$

the posterior variance of θ is

$$\text{Var}(\theta|s) = \int (\theta - E(\theta|s))^2 f(\theta|s) d\theta.$$

$$= \int \theta^2 f(\theta|s) d\theta - [E(\theta|s)]^2.$$

Ex 9. In Ex 7. the posterior distribution of θ given $x_1, \dots, x_n$ is

$$f(\theta|x_1, \dots, x_n) \propto \theta^{n\bar{x} + \alpha - 1} (1 - \theta)^{n(1 - \bar{x}) + \beta - 1}, \quad 0 \leq \theta \leq 1.$$

clearly, $\theta \sim \text{Beta}(n\bar{x} + \alpha, n(1 - \bar{x}) + \beta)$.

Find the posterior mean and variance of θ .

sol: Beta(a, b) has mean $\frac{a}{a+b}$,

$$\text{variance} = \frac{ab}{(a+b+1)(a+b)^2}.$$

$$\text{so, posterior mean} = \frac{n\bar{x} + \alpha}{n + \alpha + \beta}.$$

$$\text{posterior variance} = \frac{(n\bar{x} + \alpha)(n(1 - \bar{x}) + \beta)}{(n + \alpha + \beta + 1)(n + \alpha + \beta)^2}.$$

Ex 10. In Ex 8. Let $x_1, \dots, x_n$ be iid samples from

$N(\mu, \sigma_0^2)$, the prior distribution on μ is $N(\mu_0, \tau_0^2)$.

Where $\sigma_0^2, \mu_0, \tau_0^2$ are known. Then we obtained the posterior

distribution of μ : $N\left(\frac{n\bar{x}\tau_0^2 + \mu_0\sigma_0^2}{n\tau_0^2 + \sigma_0^2}, \frac{\sigma_0^2\tau_0^2}{n\tau_0^2 + \sigma_0^2}\right)$

So the posterior mean of μ is $\frac{n\bar{x}\tau_0^2 + \mu_0\tau_0^2}{n\tau_0^2 + \sigma_0^2}$.

the posterior variance of μ is $\frac{\sigma_0^2\tau_0^2}{n\tau_0^2 + \sigma_0^2}$.

Ex 2. In Ex 1. $x_1, \dots, x_n$ are iid samples from $\text{Exponential}(\theta)$.

while θ follows prior $\text{Exponential}(\theta_0)$, where $\theta_0 > 0$ is known.

Find the posterior mean and variance of θ .

Clearly, $\theta | x_1, \dots, x_n \sim \text{Gamma}(n+1, n\bar{x} + \theta_0)$.

So the posterior mean of θ is $\frac{n+1}{n\bar{x} + \theta_0}$.

the posterior variance = $\frac{n+1}{(n\bar{x} + \theta_0)^2}$

Ex 11. Let $X_1, \dots, X_n$ be i.i.d samples from $\text{Unif}(0, \theta)$.

the prior distribution of θ be ~~$\text{Beta}(\alpha, \beta)$~~ $\text{Exponential}(\alpha)$.

Find the posterior distribution of θ .

$$\text{Sol: } f_{\theta}(x_1, \dots, x_n) = \frac{1}{\theta} \mathbb{I}(0 < x_1 < \theta) \cdots \frac{1}{\theta} \mathbb{I}(0 < x_n < \theta).$$

$$= \frac{1}{\theta^n} \mathbb{I}(0 < x_1 < \theta) \cdots \mathbb{I}(0 < x_n < \theta)$$

$$= \begin{cases} \theta^{-n}, & \text{if } \theta > \max\{x_1, \dots, x_n\}. \\ 0, & \text{o.w.} \end{cases}$$

$$= \theta^{-n} \mathbb{I}_{\theta > \max(x_1, \dots, x_n)}.$$

So, the posterior distribution of θ is

$$f_{\theta}(x_1, \dots, x_n) \propto f_{\theta}(x_1, \dots, x_n) f(\theta) = \theta^{-n} \alpha e^{-\alpha \theta} \mathbb{I}_{\theta > \max(x_1, \dots, x_n)}.$$

$$= \alpha \theta^{-n} e^{-\alpha \theta} \mathbb{I}_{\theta > \max(x_1, \dots, x_n)}.$$