

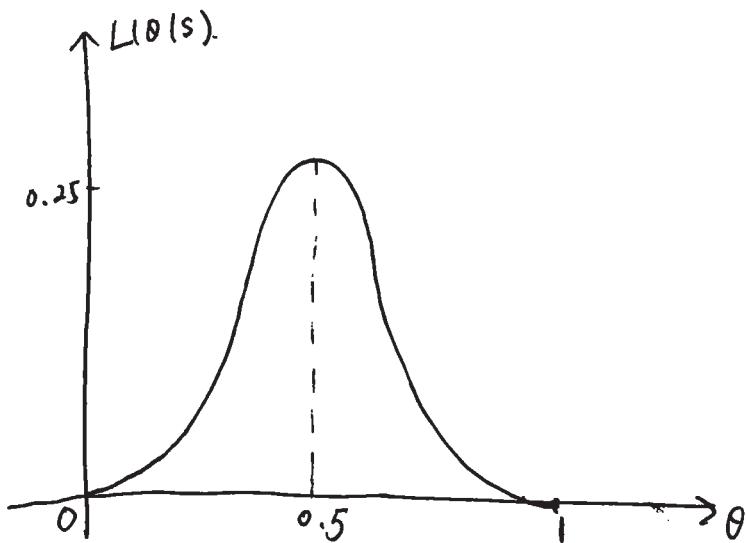
STAT 417 Lecture Note 7.

§ 6.2. Maximum likelihood principle.

In the eight-patient example. $n=8$ patients. $\theta = P(\text{death})$.
 (Individually, we estimate $\hat{\theta} = \frac{1}{2}$)
 $s=4$ is the number of dying patients. \downarrow since this is a binomial model, i.e., Binomial $(8, \theta)$. The likelihood function is

$$L(\theta|s) = \binom{8}{4} \theta^4 (1-\theta)^{8-4} = 70 \theta^4 (1-\theta)^4, \quad \theta \in [0, 1].$$

The plot of $L(\theta|s)$ is



For each θ , $L(\theta|s)$ is considered as the likelihood. A "good" θ should give us a "big" likelihood, the best θ should give us the maximum likelihood. To find $\hat{\theta}$, we only have to maximize $L(\theta|s)$. Since ~~the~~ $L(\theta|s)$ and $\log L(\theta|s)$ have the same maximizer, we maximize instead the log-likelihood, i.e.

$$\ell(\theta|s) = \log L(\theta|s).$$

In calculus, we know that maximizer of $\ell(\theta|s)$ is the solution of

$$\frac{\partial}{\partial \theta} \ell(\theta|s) = 0. \quad \dots \dots (**).$$

we define the solution to $(**)$ to be the maximum likelihood estimation, denoted as $\hat{\theta}_{ML}$.

Ex. 1. $L(\theta|s) = 70 \theta^4 (1-\theta)^4. \quad \ell(\theta|s) = \log 70 + 4 \log \theta + 4 \log (1-\theta).$

so. solving ~~$\frac{\partial}{\partial \theta} \ell(\theta|s)$~~ $\frac{\partial}{\partial \theta} \ell(\theta|s) = \frac{4}{\theta} - \frac{4}{1-\theta} = 0.$ we have

$$\hat{\theta}_{ML} = 0.5.$$

Ex. 1. $L(\theta|s) = \left(\frac{1}{4}\right) \theta^4 (1-\theta)^6. \quad$ find $\hat{\theta}_{ML}. (= 0.4).$

Ex. 2. (Normal model).

$$L(\theta|x_1, x_2) = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^2 \exp\left(-\frac{(x_1-\mu)^2 + (x_2-\mu)^2}{2\sigma^2}\right).$$

then $\ell(\theta|x_1, x_2) = -2\log(\sqrt{2\pi}\sigma) - \frac{(x_1-\mu)^2 + (x_2-\mu)^2}{2\sigma^2}.$

$$\left\{ \begin{array}{l} \frac{\partial}{\partial \mu} \ell(\theta|x_1, x_2) = -\frac{2(\mu-x_1) + 2(\mu-x_2)}{2\sigma^2} = 0 \quad \text{①} \\ \frac{\partial}{\partial \sigma} \ell(\theta|x_1, x_2) = -\frac{2}{\sigma} + \frac{(x_1-\mu)^2 + (x_2-\mu)^2}{\sigma^3} = 0. \quad \text{②} \end{array} \right.$$

From ①. $\mu = \frac{x_1+x_2}{2}, = \bar{x}.$

From ②. $\sigma^2 = \frac{(x_1-\bar{x})^2 + (x_2-\bar{x})^2}{2}, = \frac{(x_1-\bar{x})^2 + (x_2-\bar{x})^2}{2}.$

so. $\hat{\theta}_{ML} = \left(\bar{x}, \frac{(x_1-\bar{x})^2 + (x_2-\bar{x})^2}{2} \right).$

E.g. 3. (Normal model, more general).

$$L(\theta | x_1, x_2, \dots, x_n) = \left(\frac{1}{\sqrt{2\pi}\sigma} \right)^n \exp \left(- \frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2} \right).$$

$$\ell(\theta | x_1, \dots, x_n) = -n \log(\sqrt{2\pi}\sigma) - \frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2},$$

$$\text{so } \frac{\partial}{\partial \mu} \ell(\theta | x_1, \dots, x_n) = - \frac{2 \sum_{i=1}^n (x_i - \mu)}{2\sigma^2} = 0 \quad (1)$$

$$\frac{\partial}{\partial \sigma} \ell(\theta | x_1, \dots, x_n) = - \frac{n}{\sigma} + \frac{\sum_{i=1}^n (x_i - \mu)^2}{\sigma^3} = 0. \quad (2).$$

$$\text{From (1). } 0 = \sum_{i=1}^n (x_i - \mu) = n\mu - \sum_{i=1}^n x_i. \text{ so } \mu = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}.$$

$$\text{From (2). } \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{\sigma^3} = \frac{n}{\sigma}$$

$$\text{so. } \sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2.$$

Define $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ as the sample variance.

$$\text{then } \sigma^2 = \frac{n-1}{n} s^2. \text{ so } \hat{\theta}_{ML} = (\bar{x}, \frac{n-1}{n} s^2).$$

E.X. 2. ~~x_1, x_2 are i.i.d from $\text{Normal}(\alpha, \beta)$~~ .

~~$$(x_1, x_2) \xrightarrow{\alpha+1}$$~~

~~x_1, x_2, \dots, x_n~~ are i.i.d ~~F~~ samples from $\text{Exponential}(\theta)$. Find $\hat{\theta}_{ML}$.

$$\text{sol: } L(\theta | x_1, \dots, x_n) = f_\theta(x_1, \dots, x_n) = \theta^n e^{-\theta(x_1 + \dots + x_n)}.$$

$$\text{so. } \ell(\theta | x_1, \dots, x_n) = n \log \theta - \theta(x_1 + \dots + x_n).$$

$$\frac{\partial}{\partial \theta} \ell(\theta | x_1, \dots, x_n) = \frac{n}{\theta} - (x_1 + \dots + x_n) = 0 \text{ so } \hat{\theta}_{ML} = \frac{x_1 + \dots + x_n}{n} = \bar{x}.$$