

STAT 413 Lecture Note 19

Suppose the ~~the~~ posterior distribution of θ given the sample s is

$f(\theta|s)$. For any ~~and~~ set of real numbers A , we have

$$P(\theta \in A | s) = \int_A f(\theta|s) d\theta. \quad (\text{This is the STAT 416 material})$$

The posterior odds ratio in favor of $H_0: \theta \in A$ is

$$\frac{P(H_0|s)}{P(H_1|s)} = \frac{P(\theta \in A|s)}{P(\theta \notin A|s)} = \frac{P(\theta \in A|s)}{1 - P(\theta \in A|s)}$$

If we change the posterior distribution to be prior distribution, $f(\theta)$.

We can similarly define the so-called prior odds ratio in favor of H_0 .

$$\frac{P(H_0)}{P(H_1)} = \frac{P(\theta \in A)}{P(\theta \notin A)} = \frac{P(\theta \in A)}{1 - P(\theta \in A)}.$$

where $P(\theta \in A)$ is the prior probability that $\theta \in A$, or, H_0 is true.

We use $f(\theta)$ to calculate $P(\theta \in A)$, i.e.,

$$P(\theta \in A) = \int_A f(\theta) d\theta, \quad \text{where } f(\theta) \text{ is the prior of } \theta.$$

We could use prior distribution to test H_0 vs H_1 by the following rule:

If $\frac{P(H_0)}{P(H_1)} < 1$, then reject H_0 :

If $\frac{P(H_0)}{P(H_1)} > 1$, then do not reject H_0 :

Eg 3. Suppose we want to test $H_0: \theta \leq 1$ vs $H_1: \theta > 1$.

The prior distribution on θ is $N(0,1)$. ① Find the prior odds ratio in favor of H_0 , and ② Based on ①, make decision on whether or not you will reject H_0 ?

Sol: $p(H_0) = \int p(\theta \leq 1) = \Phi(1) = 1 - \Phi(-1) = 1 - 0.1587 = 0.8413$

so, the prior odds ratio is

$$\frac{p(H_0)}{1 - p(H_0)} = \frac{0.8413}{1 - 0.8413} = 5.30 > 1.$$

so, we do not reject H_0 .

Eg 4. Suppose we want to ~~test~~ test $H_0: 0 \leq \theta \leq 1$ vs $H_1: \theta > 1$.

The prior on θ is $\theta \sim \text{Exponential}(1)$. ① Find the prior odds ratio in favor of H_0 ②, do you reject H_0 ?

Sol: $p(H_0) = \int_0^1 e^{-\theta} d\theta = 0.63$

so the prior odds ratio in favor of H_0 is

$$\frac{p(H_0)}{1 - p(H_0)} = \frac{0.63}{1 - 0.63} = 1.72 > 1.$$

so, we don't reject H_0 .

Ex 1. Let $\theta =$ lifetime of a Garmin GPS. Our prior distribution

$$\text{on } \theta \text{ is } f(\theta) = \begin{cases} \frac{1}{(1+\theta)^2} & , \text{ if } \theta \geq 0 \\ 0 & \text{otherwise.} \end{cases}$$

Test $H_0: 1 \leq \theta \leq 2$ ~~vs~~ by prior odds ratio in favor of H_0 .

$$\text{Sol: } p(H_0) = \int_1^2 \frac{1}{(1+\theta)^2} d\theta = -\frac{1}{1+\theta} \Big|_1^2 = \frac{1}{6}.$$

$$\text{So } \frac{p(H_0)}{1 - p(H_0)} = \frac{\frac{1}{6}}{1 - \frac{1}{6}} = 0.2 < 1.$$

so we reject H_0 .

Ex 2: Let $\theta =$ lifetime of Garmin GPS, the prior on θ is

Exponential(α), for $\alpha > 0$. Test $H_0: \theta \geq 1$ ~~$\theta \leq 1$~~ by prior odds ratio, in favor of H_0 .

For what α do you reject H_0 ?

$$\text{Sol: } \cancel{p(H_0) = \int_1^2 \alpha e^{-\alpha\theta} d\theta = -e^{-2\alpha} + e^{-\alpha}}$$

$$p(H_0) = \int_1^{\infty} \alpha e^{-\alpha\theta} d\theta = e^{-\alpha}.$$

The prior odds ratio = $\frac{e^{-\alpha}}{1 - e^{-\alpha}}$. reject H_0 if $\frac{e^{-\alpha}}{1 - e^{-\alpha}} < 1$.

so $\alpha > \log 2 \approx 0.693$.

Bayes Factor.

Bayes factor is the third way for Bayesian hypothesis testing.

To understand the basic principal, look at a simple example.

Eg 1. Suppose we toss a coin $n=10$ times. ~~each~~ each time,

$P(\text{get } H) = p$, $P(\text{get } T) = 1 - p$. but we do not know the value of p . If p has two candidate values, $p=0.5$ or $p=0.9$.

($p=0.5$ corresponds to fair coin, $p=0.9$ corresponds to unfair coin).

We want to decide which value is correct, i.e. we test

$$H_0: p=0.5 \quad \text{vs} \quad H_1: p=0.9.$$

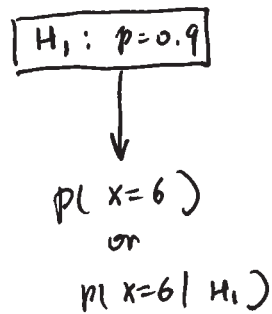
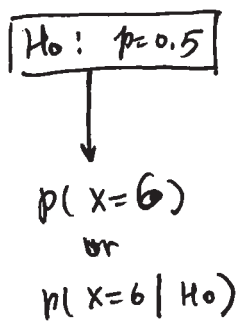
After the 10 tosses, ^{let $X = \#$ of H s among the 10 tosses} we observe $X = 6$ Heads ~~and~~

~~What~~ Question: what is our decision?

Bayes solves the problem based on the following thinking:

If H_0 is correct, i.e., $p=0.5$, then we calculate the probability of getting $X=6$ heads. Suppose the probability is very small, which will contradict our hypothesis H_0 , then we believe that H_0 is incorrect. But if the probability of getting $X=6$ heads is big, then we believe that H_0 is plausible.

The flow chart shows the idea:



If $p(X=6 | H_0) > p(X=6 | H_1)$, then prefer H_0 .

If $p(X=6 | H_0) < p(X=6 | H_1)$, then prefer H_1 .

Next we exactly calculate the two probabilities.

If H_0 is correct. $p=0.5$, then $X \sim \text{Binomial}(10, 0.5)$.

$$\Rightarrow p(X=6 | H_0) = \binom{10}{6} 0.5^6 (1-0.5)^{10-6} = (210) (0.5)^{10} = 0.21.$$

If H_1 is correct. $p=0.9$. then $X \sim \text{Binomial}(10, 0.9)$.

$$\Rightarrow p(X=6 | H_1) = \binom{10}{6} 0.9^6 (1-0.9)^{10-6} = (210) (0.9^6) (0.1)^4 = 0.01$$

So, $p(X=6 | H_0) > p(X=6 | H_1)$, which means that H_0 is NOT rejected.

The Bayes factor is defined to be

$$\frac{p(X=6 | H_0)}{p(X=6 | H_1)} = \frac{0.21}{0.01} = 21. > 1, \text{ so } H_0 \text{ is rejected.}$$

In general, Suppose we want to test H_0 vs H_1 , and we have observation S , then the Bayes factor ~~is defined~~ in favor of H_0 is defined as

$$BF_{H_0} = \frac{p(S | H_0)}{p(S | H_1)}$$

If $BF_{H_0} > 1$, then do not reject H_0 .

$BF_{H_0} < 1$, then reject H_0 .

Eg 2. Suppose ~~the number of~~ suicides occur in a population at a rate p per person per year. p is unknown. If we model the number of suicides in a population of total N persons as $\text{poisson}(Np)$, and we observed 22 suicides with $N = 30345$.

Test, by Bayes factor, $H_0: p = 0.001$ vs $H_1: p = 0.0001$

Sol: let $X = \#$ suicides in the N people.

$$X \sim \text{poisson}(Np)$$

$$\text{So, } p(X=22 | H_0) = \frac{[(30345)(0.001)]^{22}}{22!} e^{-(30345)(0.001)} = 0.024$$

$$p(X=22 | H_1) = \frac{[(30345)(\del{0.001}{0.0001})]^{22}}{22!} e^{-(30345)(0.0001)} = 1.73 \times 10^{-12}$$

$$\text{So, } BF_{H_0} = \frac{0.024}{1.73 \times 10^{-12}} \doteq 1.38 \times 10^8, > 1. \text{ So, Do not reject } H_0.$$

Relationship between Bayes factor and prior/posterior odds ratio.

Theorem: For testing H_0 vs H_1 .

$$BF_{H_0} = \frac{\text{posterior odds ratio in favor of } H_0}{\text{prior odds ratio in favor of } H_0}$$

proof: $BF_{H_0} = \frac{p(s | H_0)}{p(s | H_1)} = \frac{p(s, H_0) / p(H_0)}{p(s, H_1) / p(H_1)}$

$$= \frac{p(s, H_0)}{p(s, H_1)} \bigg/ \frac{p(H_0)}{p(H_1)}$$

$$= \left(\frac{p(s, H_0) / p(s)}{p(s, H_1) / p(s)} \right) \bigg/ \left(\frac{p(H_0)}{p(H_1)} \right)$$

$$= \frac{p(H_0 | s)}{p(H_1 | s)} \bigg/ \frac{p(H_0)}{p(H_1)}$$

$$= \frac{\text{posterior odds ratio in favor of } H_0}{\text{prior odds ratio in favor of } H_0}.$$

Ex1. If X is uniform on $\{1, 2, \dots, N\}$, where N is unknown integer-valued parameter. Want to test $H_0: N=10$ vs $H_1: N=20$. Suppose we observe a sample of X , which is $X=9$. Find BF_{H_0} and make decision.

Sol: $p(X=9 | H_0) = \frac{1}{10}$, $p(X=9 | H_1) = \frac{1}{20}$.

So, $BF_{H_0} = \frac{\frac{1}{10}}{\frac{1}{20}} = 2 > 1$. So do not reject H_0 .

Ex3. ~~If~~ A measurement X follows $N(\mu, 4)$. Want to test $H_0: \mu \leq 1$, vs $H_1: \mu > 1$. If we have observed a sample of X , which is $X=0.9$. ~~Find the posterior for~~ If the prior on μ is $\mu \sim N(0, 1)$ Find the BF_{H_0} and make decision.

Sol: $\hat{\mu} = \frac{n\bar{X}\tau_0^2 + \mu_0\sigma_0^2}{n\tau_0^2 + \sigma_0^2} = \frac{0.36}{5} = 0.072$.

$\sigma^2 = \frac{\tau_0^2\sigma_0^2}{n\tau_0^2 + \sigma_0^2} = \frac{(1)(4)}{(1)(1)+4} = \frac{4}{5} = 0.8$

So, the posterior of μ is $N(0.072, 0.8)$.

The posterior odds ratio in favor of H_0 is $\frac{p(\mu \leq 1 | X)}{p(\mu > 1 | X)} = \frac{\Phi(1.038)}{1 - \Phi(1.038)} = \frac{0.85}{1 - 0.85} = 5.7$

The prior odds ratio in favor of H_0 is $\frac{p(\mu \leq 1)}{p(\mu > 1)} = \frac{\Phi(1)}{1 - \Phi(1)} = \frac{0.84}{1 - 0.84} = 5.3$.

So $BF_{H_0} = \frac{5.7}{5.3} = 1.08 > 1$. So do not reject H_0 .

Ex2. If $x_1, x_2, \dots, x_{99}$ are i.i.d normal samples from $N(\mu, 1)$.

The prior on μ is standard normal. If it is observed that $\bar{x} = 0.1$

Test $H_0: \mu \leq 0$ vs $H_1: \mu > 0$

$$\text{Sol: } \hat{\mu} = \frac{(99)(0.1)(1) + (0)(1)}{(99)(1) + 1} = 0.099.$$

$$\sigma^2 = \frac{(1)(1)}{(99)(1) + 1} = 0.01 \quad . \text{ The posterior of } \mu \text{ is } \mu \sim N(0.099, 0.01)$$

So. the posterior OR in favor of H_0 is

$$\frac{p(\mu \leq 0 | x_1, \dots, x_{99})}{p(\mu > 0 | x_1, \dots, x_{99})} = \frac{\Phi(-0.99)}{1 - \Phi(-0.99)} = \frac{0.16}{1 - 0.16} = 0.19$$

the prior OR in favor of H_0 is

$$\frac{p(\mu \leq 0)}{p(\mu > 0)} = \frac{\Phi(0)}{1 - \Phi(0)} = 1.$$

$$\text{So } BF_{H_0} = \frac{0.19}{1} = 0.19 < 1. \quad \text{So reject } H_0.$$

STAT 417 Lecture Note 21.

Eg 4* If $x_1, x_2, \dots, x_{10}$ are iid samples from uniform $[0, \theta]$

for some unknown parameter $\theta > 0$. Assume a uniform prior on θ as

$\theta \sim \text{uniform } [0, 1]$. Suppose in practice we observed the samples as

0.1, 0.2, 0.3, 0.4, 0.2, 0.7, 0.85, 0.17, 0.29, 0.31.

Test $H_0: \theta \leq 0.9$ vs $H_1: \theta > 0.9$.

Sol: First of all, find the posterior of θ as

$$\begin{aligned} f(\theta | x_1, \dots, x_{10}) &\propto f(x_1, \dots, x_{10} | \theta) f(\theta) \\ &= \prod_{i=1}^{10} \theta^{-1} \mathbb{1}_{[0, \theta]}(x_i) \cdot f(\theta) \\ &= \begin{cases} \theta^{-10} & \text{if } x_{(10)} \leq \theta \leq 1, \\ 0 & \text{o.w.} \end{cases} \end{aligned}$$

where $x_{(10)} = \max \{x_1, \dots, x_{10}\} = 0.85$. That is,

$$f(\theta | x_1, \dots, x_{10}) \propto \theta^{-10} \quad \text{if } 0.85 \leq \theta \leq 1.$$

Assume that $f(\theta | x_1, \dots, x_{10}) = c \theta^{-10}$, $0.85 \leq \theta \leq 1$, where c is constant

such that $\int_{0.85}^1 c \theta^{-10} d\theta = 1$, so $c = 2.71$ (not necessarily to find c).

Second, the prior odds ratio in favor of H_0 is

$$\frac{p(H_0)}{p(H_1)} = \frac{\int_{0.1}^{0.9} d\theta}{\int_{0.1}^1 d\theta} = 9.$$

Third, the posterior odds ratio in favor of H_0 is

$$\frac{p(H_0 | x_1, \dots, x_{10})}{p(H_1 | x_1, \dots, x_{10})} = \frac{\int_{0.85}^{0.9} 2.71 \theta^{-10} d\theta}{\int_{0.1}^1 2.71 \theta^{-10} d\theta} = 1.0' > 1.$$

So, the Bayes factor in favor of H_0 is

$$BF_{H_0} = \frac{1.1}{9} = 0.122 < 1. \quad \text{So } H_0 \text{ is rejected.}$$

Ex3. Find the range of Bayes Factor.

Sol: $BF_{H_0} = \frac{\text{posterior OR}}{\text{prior OR}}.$

If posterior OR = 0. then $BF_{H_0} = 0$. otherwise $BF_{H_0} > 0$.

So BF_{H_0} ranges from 0 to ∞ . i.e. $BF_{H_0} \in [0, \infty)$.

Ex4. If $x_1, x_2, \dots, x_n$ are i.i.d samples from $\text{Exponential}(\theta)$, $\theta > 0$ is unknown. let θ have prior as $\text{Exponential}(1)$. $n=10, \bar{X}=2$.

Use Bayes Factor to test $H_0: \theta \geq 1$ vs $H_1: \theta < 1$.

Hint:
$$\frac{\int_1^\infty \theta^{10} e^{-21\theta} d\theta}{\int_0^1 \theta^{10} e^{-21\theta} d\theta} \approx 0.006$$

Sol: The posterior for θ is

$$f(\theta | x_1, \dots, x_n) \propto \frac{n}{i=1} \theta e^{-\theta x_i} \cdot e^{-\theta} = \theta^n e^{-(n\bar{x}+1)\theta}, \quad \theta \geq 0$$

then the prior odds ratio is

$$\frac{P(\theta \geq 1)}{P(\theta < 1)} = \frac{\int_1^\infty e^{-\theta} d\theta}{\int_0^1 e^{-\theta} d\theta} = \frac{e^{-1}}{1 - e^{-1}} = \frac{1}{e - 1} = 0.582.$$

the posterior odds ratio is

$$\frac{P(\theta \geq 1 | x_1, \dots, x_n)}{1 - P(\theta \geq 1 | x_1, \dots, x_n)} = \frac{\int_1^\infty \theta^{10} e^{-21\theta} d\theta}{\int_0^1 \theta^{10} e^{-21\theta} d\theta} = 0.006 \text{ (approx.)}, \text{ so } BF_{H_0} = \frac{0.006}{0.582} = 0.01 < 1.$$

So H_0 is rejected.