

Purdue-NCKU program

2^k **Lecture 6**
Factorial Design

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2^k Factorial Design

- Involving k factors
- Each factor has two levels (often labeled + and –)
- Factor screening experiment (preliminary study)
- Factors need not be on numeric scale
- Identify important factors and their interactions
- Interaction (of any order) has **ONE** degree of freedom

2² Factorial Design

Example:

factor			replicate			
A	B	treatment	1	2	3	mean
-	-	(1)	28	25	27	80/3
+	-	a	36	32	32	100/3
-	+	b	18	19	23	60/3
+	+	ab	31	30	29	90/3

- $y_{ijk} = \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \epsilon_{ijk}$
- Let $\bar{y}(A_+)$, $\bar{y}(A_-)$, $\bar{y}(B_+)$ and $\bar{y}(B_-)$ be the level means of A and B.
- Let $\bar{y}(A_-B_-)$, $\bar{y}(A_+B_-)$, $\bar{y}(A_-B_+)$ and $\bar{y}(A_+B_+)$ be the treatment means

Main Effect

Define main effects of A (denoted again by A) as follows:

$$\begin{aligned} A &= m.e.(A) = \bar{y}(A_+) - \bar{y}(A_-) \\ &= \frac{1}{2}(\bar{y}(A_+B_+) + \bar{y}(A_+B_-)) - \frac{1}{2}(\bar{y}(A_-B_+) + \bar{y}(A_-B_-)) \\ &= \frac{1}{2}(\bar{y}(A_+B_+) + \bar{y}(A_+B_-) - \bar{y}(A_-B_+) - \bar{y}(A_-B_-)) \\ &= \frac{1}{2}(-\bar{y}(A_-B_-) + \bar{y}(A_+B_-) - \bar{y}(A_-B_+) + \bar{y}(A_+B_+)) \\ &= 8.33 \end{aligned}$$

- Let $C_A = (-1, 1, -1, 1)$, a contrast on treatment mean responses, then

$$m.e.(A) = \frac{1}{2}\hat{C}_A$$

- Notice that

$$A = m.e.(A) = (\bar{y}(A_+) - \bar{y}..) - (\bar{y}(A_-) - \bar{y}..) = \hat{\tau}_2 - \hat{\tau}_1 = 2\hat{\tau}_2$$

Main effect is defined in a different way than the factorial modeling. But they are connected and equivalent.

Interaction

- Interaction between A and B: does the effect of A depend on the level of B?
- Define interaction between A and B

$$AB = \text{Int}(AB) = \frac{1}{2}(m.e.(A | B_+) - m.e.(A | B_-))$$

$$= \frac{1}{2}(\bar{y}(A_+ | B_+) - \bar{y}(A_- | B_+)) - \frac{1}{2}(\bar{y}(A_+ | B_-) - \bar{y}(A_- | B_-))$$

$$= \frac{1}{2}(\bar{y}(A_- B_-) - \bar{y}(A_+ B_-) - \bar{y}(A_- B_+) + \bar{y}(A_+ B_+))$$

Let $C_{AB} = (1, -1, -1, 1)$, a contrast on treatment means, then

$$AB = \text{Int}(AB) = \frac{1}{2}\hat{C}_{AB}$$

- Notice that $\text{Int}(AB) = \hat{\tau}\beta_{22} - \hat{\tau}\beta_{21} =$ similar interaction factorial effects difference $= 2\hat{\tau}\beta_{22}$

Effects and Contrasts

factor				effect (contrast)			
A	B	total	mean	I	A	B	AB
-	-	80	80/3	1	-1	-1	1
+	-	100	100/3	1	1	-1	-1
-	+	60	60/3	1	-1	1	-1
+	+	90	90/3	1	1	1	1

- There is a one-to-one correspondence between effects and contrasts, and contrasts can be directly used to estimate the effects.
- For a effect corresponding to contrast $c = (c_1, c_2, \dots)$ in 2^2 design

$$\text{effect} = \frac{1}{2} \sum_i c_i \bar{y}_i$$

where i is an index for treatments and the summation is over all treatments.

- Pay attention to the column of the contrast matrix

2³ Factorial Design

factor			treatment	response		
A	B	C		1	2	total
-	-	-	(1)	-3	-1	-4
+	-	-	a	0	1	1
-	+	-	b	-1	0	-1
+	+	-	ab	2	3	5
-	-	+	c	-1	0	-1
+	-	+	ac	2	1	3
-	+	+	bc	1	1	2
+	+	+	abc	6	5	11

$$y_{ijkl} = \mu + \tau_i + \beta_j + \gamma_k + (\tau\beta)_{ij} + (\beta\gamma)_{jk} + (\tau\gamma)_{ik} + (\tau\beta\gamma)_{ijk} + \epsilon_{ijkl}$$

factorial effects and contrasts

Main effects:

$$\begin{aligned}A &= m.e.(A) = \bar{y}(A_+) - \bar{y}(A_-) \\&= \frac{1}{4}(\bar{y}(- - -) + \bar{y}(+ - -) - \bar{y}(- + -) + \bar{y}(+ + -) - \bar{y}(- - +) \\&\quad + \bar{y}(+ - +) - \bar{y}(- + +) + \bar{y}(+ + +)) \\&= 3.00 = 2\hat{\tau}_2\end{aligned}$$

The contrast is $(-1, 1, -1, 1, -1, 1, -1, 1)$

$$B : (-1, -1, 1, 1, -1, -1, 1, 1), B = 2.25$$

$$C : (-1, -1, -1, -1, 1, 1, 1, 1), C = 1.75$$

2-factor interactions:

$$AB: A \times B \text{ componentwise, } AB = .75 = \tau\hat{\beta}_{22}$$

$$AC: A \times C \text{ componentwise, } AC = .25$$

$$BC: B \times C \text{ componentwise, } BC = .50$$

High order interaction

k -th order interaction means: does the $(k - 1)$ -th interaction depend on level of the k -th factor

3-factor interaction:

$$\begin{aligned} ABC = int(ABC) &= \frac{1}{2}(int(AB | C+) - int(AB | C-)) \\ &= \frac{1}{4}(-\bar{y}(- - -) + \bar{y}(+ - -) + \bar{y}(- + -) - \bar{y}(+ + -) \\ &\quad + \bar{y}(- - +) - \bar{y}(+ - +) - \bar{y}(- + +) + \bar{y}(+ + +)) \\ &= .50 = 2\tau\hat{\beta}\gamma_{222} \end{aligned}$$

The contrast is $(-1, 1, 1, -1, 1, -1, -1, 1) = A \times B \times C$.

Contrasts for Calculating Effects in 2^3 Design

			factorial effects								
A	B	C	treatment	<i>I</i>	<i>A</i>	<i>B</i>	<i>AB</i>	<i>C</i>	<i>AC</i>	<i>BC</i>	<i>ABC</i>
-	-	-	(1)	1	-1	-1	1	-1	1	1	-1
+	-	-	a	1	1	-1	-1	-1	-1	1	1
-	+	-	b	1	-1	1	-1	-1	1	-1	1
+	+	-	ab	1	1	1	1	-1	-1	-1	-1
-	-	+	c	1	-1	-1	1	1	-1	-1	1
+	-	+	ac	1	1	-1	-1	1	1	-1	-1
-	+	+	bc	1	-1	1	-1	1	-1	1	-1
+	+	+	abc	1	1	1	1	1	1	1	1

Estimates:

$$\text{grand mean: } \frac{\sum \bar{y}_i}{2^3}$$

$$\text{effect : } \frac{\sum c_i \bar{y}_i}{2^{3-1}}$$

General 2^k Design

- k factors: A, B, \dots, K each with 2 levels (+, -)
- consists of all possible level combinations (2^k treatments) each with n replicates
- Classify factorial effects:

type of effect	label	the number of effects
main effects (of order 1)	A, B, C, \dots, K	k
2-factor interactions (of order 2)	AB, AC, \dots, JK	$\binom{k}{2}$
3-factor interactions (of order 3)	ABC, ABD, \dots, IJK	$\binom{k}{3}$
...
k-factor interaction (of order k)	$ABC \dots K$	$\binom{k}{k}$

- Each effect (main or interaction) has 1 degree of freedom
full model (i.e. model consisting of all the effects) has $2^k - 1$ degrees of freedom.
- Error component has $2^k(n - 1)$ degrees of freedom.
- One-to-one correspondence between effects and contrasts:
 - For main effect: convert the level column of a factor using
– $\Rightarrow -1$ and $\dagger \Rightarrow 1$
 - For interactions: multiply the contrasts of the main effects of the involved factors, componentwisely.

General 2^k Design: Analysis

- Estimates:

$$\text{grand mean} : \frac{\sum \bar{y}_i}{2^k}$$

For effect with contrast $C = (c_1, c_2, \dots, c_{2^k})$, its estimate is

$$\text{effect} = \frac{\sum c_i \bar{y}_i}{2^{(k-1)}}$$

- Variance

$$\text{Var}(\text{effect}) = \frac{\sigma^2}{n2^{k-2}}$$

$$\text{S.E.}(\text{effect}) = \frac{\text{MSE}}{n2^{k-2}}$$

- C.I. for every factorial effect

$$\text{effect} \pm t_{\alpha/2, 2^k(n-1)} \text{S.E.}(\text{effect})$$

Unreplicated 2^k Design

- $n = 1$
- No degree of freedom left for error component if full model is fitted.
- Same estimation method
- No error sum of squares available, cannot estimate σ^2 and test effects in both the ANOVA and Regression approaches.
- **Approach 1:** pooling high-order interactions
 - Often assume 3 or higher interactions do not occur
 - Pool estimates together for error
 - Warning: may pool significant interaction

- Approach 2: Using the normal probability plot (QQ plot) to identify significant effects.

- Recall

$$\text{Var}(\text{effect}) = \frac{\sigma^2}{2^{(k-2)}}$$

If the effect is not significant ($=0$), then the effect estimate follows $N(0, \frac{\sigma^2}{2^{(k-2)}})$

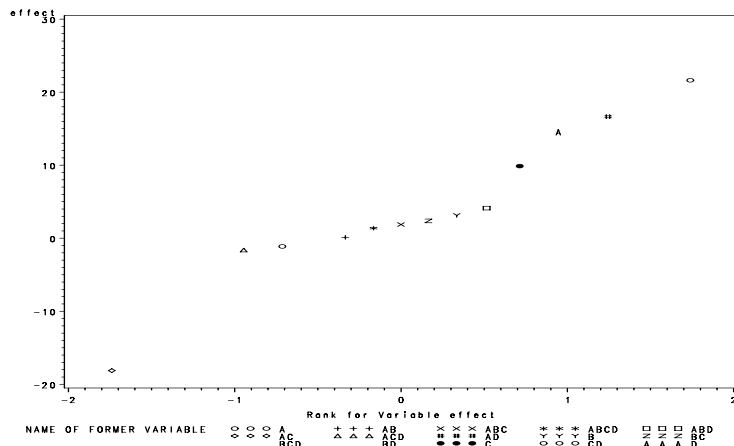
- Assume all effects not significant, their estimates can be considered as a random sample from $N(0, \frac{\sigma^2}{2^{(k-2)}})$
- QQ plot of the estimates is expected to be a linear line
- Deviation from a linear line indicates significant effects

A case study

factor				filtration
<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	
-	-	-	-	45
+	-	-	-	71
-	+	-	-	48
+	+	-	-	65
-	-	+	-	68
+	-	+	-	60
-	+	+	-	80
+	+	+	-	65
-	-	-	+	43
+	-	-	+	100
-	+	-	+	45
+	+	-	+	104
-	-	+	+	75
+	-	+	+	86
-	+	+	+	70
+	+	+	+	96

ALL Ranked Effects

Obs	_NAME_	COL1	effect	neff
1	AC	-9.0625	-18.125	-1.73938
2	BCD	-1.3125	-2.625	-1.24505
3	ACD	-0.8125	-1.625	-0.94578
4	CD	-0.5625	-1.125	-0.71370
5	BD	-0.1875	-0.375	-0.51499
6	AB	0.0625	0.125	-0.33489
7	ABCD	0.6875	1.375	-0.16512
8	ABC	0.9375	1.875	-0.00000
9	BC	1.1875	2.375	0.16512
10	B	1.5625	3.125	0.33489
11	ABD	2.0625	4.125	0.51499
12	C	4.9375	9.875	0.71370
13	D	7.3125	14.625	0.94578
14	AD	8.3125	16.625	1.24505
15	A	10.8125	21.625	1.73938



Effect Selection and Analysis

- Potentially significant effects: A, AD, C, D, AC .
- ANOVA model involving only A, C, D and their interactions (projecting the original unreplicated 2^4 experiment onto a replicated 2^3 experiment)
- Make inferences with non-ZERO MSE
- Diagnostics using residuals.

2^{k-p} Fractional Factorial Design

Fundamental Principles Regarding Factorial Effects

Suppose there are k factors (A, B, \dots, J, K) in an experiment. All possible factorial effects include

effects of order 1: A, B, \dots, K (main effects)

effects of order 2: AB, AC, \dots, JK (2-factor interactions)

.....

- Hierarchical Ordering principle
 - Lower order effects are more likely to be important than higher order effects.
 - Effects of the same order are equally likely to be important
- Effect Sparsity Principle (Pareto principle)
 - The number of relatively important effects in a factorial experiment is small
- Effect Heredity Principle
 - In order for an interaction to be significant, at least one of its parent factors should be significant.

Fractional Factorials

- May not have sources (time, money, etc) for full factorial design
- Number of runs required for full factorial grows quickly
 - Consider 2^k design
 - If $k = 7 \rightarrow 128$ runs required
 - Can estimate 127 effects
 - Only 7 df for main effects, 21 for 2-factor interactions
 - the remaining 99 df are for interactions of order ≥ 3
- Often only lower order effects are important
- Full factorial design may not be necessary according to
 - Hierarchical ordering principle
 - Effect Sparsity Principle
- A fraction of the full factorial design (i.e. a subset of all possible level combinations) is sufficient.

There are four factors in the experiment (A , B , C and D), each of 2 levels. Suppose the available resource is enough for conducting 8 runs. 2^4 full factorial design consists of all the 16 level combinations of the four factors. We need to choose half of them.

- If you drop one factors for a 2^3 full factorial design, this factor and their interactions with other factors cannot be investigated.
- Want investigate all 4 factors in the experiment
- A fraction of 2^4 full factorial design will be used.
- Confounding (aliasing) will happen because using a subset

The chosen half is called 2^{4-1} fractional factorial design.

2^{4-1} Fractional Factorial Design

- the number of factors: $k = 4$
- the fraction index: $p = 1$
- the number of runs (level combinations): $N = \frac{2^4}{2^1} = 8$
- Construct 2^{4-1} designs via “confounding” (**aliasing**)
 - select 3 factors (e.g. A, B, C) to form a 2^3 full factorial (basic design)
 - confound (**alias**) D with a high order interaction of A, B and C . For example,

$$D = ABC$$

factorial effects (contrasts)							
I	A	B	C	AB	AC	BC	ABC=D
1	-1	-1	-1	1	1	1	-1
1	1	-1	-1	-1	-1	1	1
1	-1	1	-1	-1	1	-1	1
1	1	1	-1	1	-1	-1	-1
1	-1	-1	1	1	-1	-1	1
1	1	-1	1	-1	1	-1	-1
1	-1	1	1	-1	-1	1	-1
1	1	1	1	1	1	1	1

- Note: 1 corresponds to $+$ and -1 corresponds to $-$.

Verify:

1. the chosen level combinations form a half of the 2^4 design.
2. the product of columns A , B , C and D equals 1, i.e.,

$$I = ABCD$$

which is called the **defining relation**, or $ABCD$ is called a **defining word** (contrast).

Aliasing in 2^{4-1} Design

For four factors A , B , C and D , there are $2^4 - 1$ effects: A , B , C , D , AB , AC , AD , BC , BD , CD , ABC , ABD , ACD , BCD , $ABCD$

Response	I	A	B	C	D	AB	..	CD	ABC	BCD	...	ABCD
y_1	1	-1	-1	-1	-1	1	..	1	-1	-1	...	1
y_2	1	1	-1	-1	1	-1	..	-1	1	1	...	1
y_3	1	-1	1	-1	1	-1	..	-1	1	-1	...	1
y_4	1	1	1	-1	-1	1	..	1	-1	1	...	1
y_5	1	-1	-1	1	1	1	..	1	1	-1	...	1
y_6	1	1	-1	1	-1	-1	..	-1	-1	1	...	1
y_7	1	-1	1	1	-1	-1	..	-1	-1	-1	...	1
y_8	1	1	1	1	1	1	..	1	1	1	...	1

Contrasts for main effects by converting $-$ to -1 and $+$ to 1; contrasts for other effects obtained by multiplication.

$$A = \bar{y}_{A+} - \bar{y}_{A-} = \frac{1}{4}(-y_1 + y_2 - y_3 + y_4 - y_5 + y_6 - y_7 + y_8)$$

$$BCD = \frac{1}{4}(-y_1 + y_2 - y_3 + y_4 - y_5 + y_6 - y_7 + y_8)$$

A , BCD are aliases or aliased. The contrast is for $A+BCD$. A and BCD are not distinguishable.

$$AB = \bar{y}_{AB+} - \bar{y}_{AB-} = \frac{1}{4}(y_1 - y_2 - y_3 + y_4 + y_5 - y_6 - y_7 + y_8) \quad CD = \bar{y}_{CD+} - \bar{y}_{CD-} = \frac{1}{4}(y_1 - y_2 - y_3 + y_4 + y_5 - y_6 - y_7 + y_8)$$

AB , CD are aliases or aliased. The contrast is for $AB+CD$. AB and CD are not distinguishable.

There are other 5 pairs. They are caused by the defining relation

$$I = ABCD,$$

that is, I (the intercept) and 4-factor interaction $ABCD$ are aliased.

Alias Structure for 2^{4-1} with $I = ABCD$

- Alias Structure:

$$I = ABCD$$

$$A = A * I = A * ABCD = BCD$$

$$B = \dots\dots\dots = ACD$$

$$C = \dots\dots\dots = ABD$$

$$D = \dots\dots\dots = ABC$$

$$AB = AB * I = AB * ABCD = CD$$

$$AC = \dots\dots\dots = BD$$

$$AD = \dots\dots\dots = BC$$

- all 16 factorial effects for A , B , C and D are partitioned into 8 groups each with 2 aliased effects.
- When a low order effect is aliased with a high order effect, by Hierarchical Order principle, we tend to believe that the effect is mostly contributed by the low order effect

A Different 2^{4-1} Design

- the defining relation $I = ABD$ generates another 2^{4-1} fractional factorial design, denoted by d_2 . Its alias structure is given below.

$$I = ABD$$

$$A = BD$$

$$B = AD$$

$$C = ABCD$$

$$D = AB$$

$$ABC = CD$$

$$ACD = BC$$

$$BCD = AC$$

- Recall d_1 is defined by $I = ABCD$. Comparing d_1 and d_2 , which one we should choose or which one is better?
 1. **Length** of a defining word is defined to be the number of the involved factors.
 2. **Resolution** of a fractional factorial design is defined to be the minimum length of the defining words, usually denoted by Roman numbers, III, IV, V, etc...

Resolution and Maximum Resolution Criterion

- $d_1: I = ABCD$ is a resolution IV design denoted by 2_{IV}^{4-1} .
- $d_2: I = ABD$ is a resolution III design denoted by 2_{III}^{4-1} .
- If a design is of resolution R , then none of the i -factor interactions is aliased with any other interaction of order less than $R - i$.

d_1 : main effects are not aliased with other main effects and 2-factor interactions

d_2 : main effects are not aliased with main effects

- d_1 is better, because d_1 has higher resolution than d_2 . In fact, d_1 is optimal among all the possible fractional factorial 2^{4-1} designs

- **Maximum Resolution Criterion**

fractional factorial design with maximum resolution is optimal

How to Analyze 2^{4-1} design

- Compute all effects
- Use QQ plot to determine which ones are significant
- Resolve the ambiguities in aliased effects via the fundamental principles beneficial
- Project the design to a replicated factorial design
- Example
 - $I=ABCD$
 - QQ plot determine A, B, CD are significant
 - By HO principle C, D are not significant
 - By EH principle, CD are not significant
 - All significant effects are A, B and AB
 - View the data as a 2^2 experiment with 2 replications

General 2^{k-1} Design

- k factors: A, B, \dots, K
- can only afford half of all the combinations (2^{k-1})
- Basic design: a 2^{k-1} full factorial for $k - 1$ factors: A, B, \dots, J .
- The setting of k th factor is determined by aliasing K with the $ABC\dots J$, i.e., $K = ABC \dots J$
- Defining relation: $I = ABCD\dots\tilde{I}JK$. Resolution= k
- 2^k factorial effects are partitioned into 2^{k-1} groups each with two aliased effects.
- Use fundamental principles, domain knowledge, follow-up experiment to de-alias.

One Quarter Fraction: 2^{k-2} Design

Parts manufactured in an injection molding process are showing excessive shrinkage. A quality improvement team has decided to use a designed experiment to study the injection molding process so that shrinkage can be reduced.

The team decides to investigate six factors

A: mold temperature

B: screw speed

C: holding time

D: cycle time

E: gate size

F: holding pressure

each at two levels, with the objective of learning about main effects and interactions.

They decide to use 16-run fractional factorial design.

- a full factorial has $2^6=64$ runs.
- 16-run is one quarter of the full factorial
- How to construct the fraction?

Injection Molding Experiment: 2^{6-2} Design

basic design						
<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	$E = ABC$	$F = BCD$	shrinkage
-	-	-	-	-	-	6
+	-	-	-	+	-	10
-	+	-	-	+	+	32
+	+	-	-	-	+	60
-	-	+	-	+	+	4
+	-	+	-	-	+	15
-	+	+	-	-	-	26
+	+	+	-	+	-	60
-	-	-	+	-	+	8
+	-	-	+	+	+	12
-	+	-	+	+	-	34
+	+	-	+	-	-	60
-	-	+	+	+	-	16
+	-	+	+	-	-	5
-	+	+	+	-	+	37
+	+	+	+	+	+	52

Two defining relations are used to generate the columns for E and F .

$$I = ABCE, \text{ and } I = BCDF$$

They induce another defining relation:

$$I = ABCE * BCDF = AB^2C^2DEF = ADEF$$

The complete defining relation:

$$I = ABCE = BCDF = ADEF$$

Defining contrasts subgroup: $\{I, ABCE, BCDF, ADEF\}$

Alias Structure

$I = ABCE = BCDF = ADEF$ implies

$$A = BCE = ABCDF = DEF$$

Similarly, we can derive the other groups of aliased effects.

$A = BCE = DEF = ABCDF$	$AB = CE = ACDF = BDEF$
$B = ACE = CDF = ABDEF$	$AC = BE = ABDF = CDEF$
$C = ABE = BDF = ACDEF$	$AD = EF = BCDE = ABCF$
$D = BCF = AEF = ABCDE$	$AE = BC = DF = ABCDEF$
$E = ABC = ADF = BCDEF$	$AF = DE = BCEF = ABCD$
$F = BCD = ADE = ABCEF$	$BD = CF = ACDE = ABEF$
	$BF = CD = ACEF = ABDE$

$ABD = CDE = ACF = BEF$
$ACD = BDE = ABF = CEF$

Wordlength pattern $W = (W_0, W_1, \dots, W_6)$, where W_i is the number of defining words of length i (i.e., involving i factors)

$$W = (1, 0, 0, 0, 3, 0, 0)$$

Resolution is the smallest i such that $i > 0$ and $W_i > 0$. Hence it is a 2_{IV}^{6-2} design

2^{6-2} Design: an Alternative

- Basic Design: A, B, C, D
- $E = ABCD, F = ABC$, i.e., $I = ABCDE$, and $I = ABCF$
- which induces: $I = DEF$
- complete defining relation: $I = ABCDE = ABCF = DEF$
- wordlength pattern: $W = (1, 0, 0, 1, 1, 1, 0)$

- Alias structure (ignore effects of order 3 or higher)

$A = ..$	$AB = CF = ..$
$B = ..$	$AC = BF = ..$
$C = ..$	$AD = ..$
$D = EF = ..$	$AE = ..$
$E = DF = ..$	$AF = BC = ..$
$F = DE = ..$	$BD = ..$
	$BE = ..$
	$CD = ..$
	$CE = ..$

- an effect is said to be **clearly estimable** if it is not aliased with main effect or two-factor interactions.
- Which design is better d_1 or d_2 ? d_1 has six clearly estimable main effects while d_2 has three clearly estimable main effects and six clearly estimable two-factor ints.

Minimum Aberration Criterion

Recall 2^{k-p} with maximum resolution should be preferred. But, it is possible that there are two designs that attain the maximum resolution. How should we further distinguish them?

For example, consider 2^{7-2} fractional factorial design

d_1 : basic design: A, B, C, D, E ; $F = ABC, G = ABDE$

complete defining relation: $I = ABCF = ABDEG = CDEFG$

wordlength pattern: $W = (1, 0, 0, 0, 1, 2, 0, 0)$

Resolution: IV

d_2 : basic design: A, B, C, D, E ; $F = ABC, G = ADE$

complete defining relation: $I = ABCF = ADEG = BCDEFG$

wordlength pattern: $W = (1, 0, 0, 0, 2, 0, 1, 0)$

Resolution: IV

d_1 and d_2 , which is better?

Minimum Aberration Criterion

Definition: Let d_1 and d_2 be two 2^{k-p} designs, let r be the smallest **positive** integer such that $W_r(d_1) \neq W_r(d_2)$.

If $W_r(d_1) < W_r(d_2)$, then d_1 is said to have less aberration than d_2 .

If there does not exist any other design that has less aberration than d_1 , then d_1 has minimum aberration.

Chapter Review

- 2^k design
- 2^k design without replication
- 2^{k-p} design