STAT 516: Discrete Random Variables Lecture 5a: Generating Functions

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Levine STAT 516: Discrete Random Variables

Generating Function (GF)

- ► For a discrete nonnegative integer-valued random variable X define $G(s) = G_X(s) = \mathbb{E}(s^X) = \sum_{x=0}^{\infty} s^x P(X = x)$
- ► G(s) is always finite when s ≤ 1 for any random variable X; it may be finite in other regions as well
- Property 1: Let G(s) be finite in some open interval around the origin. Then, P(X = k) = G^(k)(0)/k! for any k ≥ 0
- ► Property 2: If lim_{s→1} G^(k)(s) is finite, the kth factorial moment

$$\mathbb{E}\left[(X(X-1)\dots(X-k+1)] = G^{(k)}(1) = \lim_{s \to 1} G^{(k)}(s)
ight]$$

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GF of a sum of independent RV's; distribution determining property

▶ Property 1: X₁,..., X_n are independent RV's with GF's G₁(s) ..., G_n(s). Then,

$$G_{X_1+\cdots+X_n}(s)=\prod_{i=1}^n G_i(s)$$

- The proof is trivial
- Property 2: If two GF's G(s) and H(s) coincide in any nonempty open interval, then X and Y have the same distribution
- The proof is based on a standard property of the power series

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• Take X a discrete uniform on $\{1, 2, \ldots, n\}$. Then,

$$G(s) = \mathbb{E}[s^X] = \frac{1}{n} \sum_{x=1}^n s^x = \frac{s(s^n - 1)}{n(s - 1)}$$

 Differentiating at 1 and applying the L'Hôpital's rule twice, we find that

$$G^{'}(1)=\frac{n+1}{2}$$

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- Take X with a pmf $p(x) = e^{-1} \frac{1}{x!}$, x = 0, 1, 2, ...
- Clearly, $p(x) \ge 0$ and $\sum_{x=0}^{\infty} p(x) = 1$
- The generating function is

$$G(s) = \mathbb{E}[s^X] = \sum_{x=0}^{\infty} s^x e^{-1} \frac{1}{x!} = e^{s-1}$$

- ▶ Take the first derivative and conclude that $\mathbb{E}(X) = G^{'}(1) = 1$
- We discovered the Poisson distribution...

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- If M(t) = E(exp tX) exists in an open neighborhood of size h around zero, it is called a moment generating function or mgf
- M(t) always exists for t = 0; if there exists a radius h s.t. M(t) < ∞ for any t : |t| < h, many useful properties can be derived
- ► Also note that the generating function G(t) = M(log t) whenever G(t) < ∞</p>

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If the MGF M(t) is finite in some open interval |t| < h, then it is infinitely differentiable in that interval and, for any k ≥ 1,

$$\mathbb{E}X = M^{(k)}(0)$$

- ► (Distribution determining property)For any two random variables X and Y with existing M_X(t) and M_Y(t) the distributions coincide if and only M_X(t) = M_Y(t) for t ∈ (-h, h) and h > 0
- ► If X₁,..., X_n are independent random variables, and each X_i has a finite mgf in an open interval around 0, we have

$$M_{X_1+\ldots+X_n}(t)=\prod_{i=1}^n M_{X_i}(t)$$

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Let X have the pmf P(X = x) = ¹/_n, x = 1, 2, ..., n
 Its mgf is

$$M(t) = \frac{1}{n} \sum_{k=1}^{n} e^{tk} = \frac{e^t(e^{nt} - 1)}{n(e^t - 1)}$$

To obtain the first derivative at zero; apply L'Hôpital's rule twice to find

$$\mathbb{E}(X) = \frac{n+1}{2}$$

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• Thus,
$$\mathbb{E}X = \mathbb{E}X^2 = p$$

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Example III

- ► For a fair spinner with the numbers 1, 2 and 3 on it let X be the number of spins until the first 3 occurs
- If the spins are independent, we have the geometric distribution
 1 (2) X⁻¹

$$p(x) = \frac{1}{3} \left(\frac{2}{3}\right)^{x-1}$$

for
$$x = 1, 2, 3, \ldots$$

The mgf of X is

$$M(t) = \mathbb{E}(\exp tX) = \sum_{x=1}^{\infty} \exp(tx) \frac{1}{3} \left(\frac{2}{3}\right)^{x-1} = \frac{1}{3} \left(1 - \exp(t)\frac{2}{3}\right)^{-1}$$

as long as $rac{2}{3}\exp\left(t
ight) < 1$ t.i. $t < \log\left(rac{3}{2}
ight)$

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Define

$$p(x) = \frac{6}{\pi^2 x^2}$$

for $x = 1, 2, 3, \ldots$ and 0 otherwise

- Easy to see that $\mathbb{E}(tX) = \sum_{x=1}^{\infty} \frac{6 \exp(tx)}{\pi^2 x^2} = +\infty$
- Therefore, this distribution doesn't have an mgf

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A simple cure is to define a characteristic function as

$$\phi(t) = \mathbb{E}(itX)$$

- $\phi(t)$ exists for *all* distributions
- $\phi(t)$ also defines the distribution uniquely
- For any existing moment of X we have

$$i^k \mathbb{E} X^k = \phi^{(k)}(0)$$