

7.

- a. $\mu_{Y,2500} = 1800 + 1.3(2500) = 5050$
- b. expected change = slope = $\beta_1 = 1.3$
- c. expected change = $100\beta_1 = 130$
- d. expected change = $-100\beta_1 = -130$

8.

- a. $\mu_{Y,2000} = 1800 + 1.3(2000) = 4400$, and $\sigma = 350$, so $P(Y > 5000) = P\left(Z > \frac{5000 - 4400}{350}\right) = P(Z > 1.71) = .0436$.
- b. Now $E(Y) = 5050$, so $P(Y > 5000) = P(Z > -.14) = .5557$.
- c. $E(Y_2 - Y_1) = E(Y_2) - E(Y_1) = 5050 - 4400 = 650$, and $V(Y_2 - Y_1) = V(Y_2) + V(Y_1) = (350)^2 + (350)^2 = 245,000$, so the sd of $Y_2 - Y_1$ is 494.97. Thus $P(Y_2 - Y_1 > 0) = P\left(Z > \frac{1000 - 650}{494.97}\right) = P(Z > .71) = .2389$.
- d. The standard deviation of $Y_2 - Y_1$ is 494.97 (from c), and $E(Y_2 - Y_1) = 1800 + 1.3x_2 - (1800 + 1.3x_1) = 1.3(x_2 - x_1)$. Thus $P(Y_2 > Y_1) = P(Y_2 - Y_1 > 0) = P\left(z > \frac{-1.3(x_2 - x_1)}{494.97}\right) = .95$ implies that $-1.645 = \frac{-1.3(x_2 - x_1)}{494.97}$, so $x_2 - x_1 = 626.33$.

12.

- a. From the summary provided, $\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} = \frac{-341.959231}{1585.230769} = -0.21572$ and

$$\hat{\beta}_0 = \frac{\Sigma y - \hat{\beta}_1 \Sigma x}{n} = \frac{52.8 - (-0.21572)(303.7)}{13} = 9.1010. \text{ So the equation of the least squares regression}$$

line is $\hat{y} = 9.1010 - .21572x$.

Based on this equation, the predicted ammonium concentration (y) when transpiration (x) is 25 ml/h is $\hat{y}(25) = 9.1010 - .21572(25) = 3.708$ mg/L.

- b. If you plug $x = 45$ into the least squares regression line, you get $\hat{y}(45) = 9.1010 - .21572(45) = -0.606$. That's an impossible ammonium concentration level, since concentration can't be negative. But it doesn't make sense to predict y at $x = 45$ from this data set, because $x = 45$ is well outside of the scope of the data (this is an example of extrapolation and its potential adverse consequences).

- c. With the aid of software, $SSE = \sum_{i=1}^{13} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{13} (y_i - [9.1010 - .21572x_i])^2 = \dots = 3.505$. Or, using the available sum of squares and a derivation similar to the one described in the section, $SSE = S_{yy} - \hat{\beta}_1 S_{xy} = 77.270769 - (-0.21572)(-341.959231) = 3.505$. Either way, the residual standard

$$\text{deviation is } s = \sqrt{\frac{SSE}{n-2}} = \sqrt{\frac{3.505}{13-2}} = 0.564.$$

The typical difference between a sample's actual ammonium concentration and the concentration predicted by the least squares regression line is about ± 0.564 mg/L.

- d. With $SST = S_{yy} = 77.270769$, $r^2 = 1 - \frac{SSE}{SST} = 1 - \frac{3.505}{77.270769} = .955$. So, the least squares regression line

helps to explain 95.5% of the total variation in ammonium concentration. Given that high percentage and the linear relationship visible in the scatterplot (see Exercise 4), yes, the model does a good job of explaining observed variation in ammonium concentration.

30.

- a. $\Sigma(x_i - \bar{x})^2 = 7,000,000$, so $V(\hat{\beta}_1) = \frac{(350)^2}{7,000,000} = .0175$ and the standard deviation of $\hat{\beta}_1$ is $\sqrt{.0175} = .1323$.

- b. $P(1.0 \leq \hat{\beta}_1 \leq 1.5) = P\left(\frac{1.0-1.25}{.1323} \leq Z \leq \frac{1.5-1.25}{.1323}\right) = P(-1.89 \leq Z \leq 1.89) = .9412$.

- c. Although $n = 11$ here and $n = 7$ in **a**, $\Sigma(x_i - \bar{x})^2 = 1,100,000$ now, which is smaller than in **a**. Because this appears in the denominator of $V(\hat{\beta}_1)$, the variance is smaller for the choice of x values in **a**.

32. Let β_1 denote the true average change in runoff for each 1 m³ increase in rainfall. To test the hypotheses

$$H_0: \beta_1 = 0 \text{ versus } H_a: \beta_1 \neq 0, \text{ the calculated } t \text{ statistic is } t = \frac{\hat{\beta}_1 - 0}{s_{\hat{\beta}_1}} = \frac{.82697}{.03652} = 22.64 \text{ which (from the}$$

printout) has an associated P -value of ~ 0.000 . Therefore, since the P -value is so small, H_0 is rejected and we conclude that there is a useful linear relationship between runoff and rainfall.

A confidence interval for β_1 is based on $n - 2 = 15 - 2 = 13$ degrees of freedom. Since $t_{.025,13} = 2.160$, the interval estimate is $\hat{\beta}_1 \pm t_{.025,13} \cdot s_{\hat{\beta}_1} = .82697 \pm (2.160)(.03652) = (.748, .906)$. Therefore, we can be confident that the true average change in runoff, for each 1 m^3 increase in rainfall, is somewhere between $.748 \text{ m}^3$ and $.906 \text{ m}^3$.