## Answer Keys to Homework#7

## Problem 0

(a) I will answer this in terms of 4 trts rather than two sets of two treatments. A RCBD or BIBD could be used here. For the RCBD to be balanced, we utilize only four of the six possible runs each day and run the experiment over 6 days.

Source	$\mathrm{DF}$
Day	5
Trt	3
Error	15
Total	23

For the BIBD, there are several ways to do this. If only two combinations are run each day, a total of 6 days are needed for the design to be balanced. Since  $6 \times 2 = 12$ , the combinations each day could be replicated or you could run the design over 12 days. Likewise, one could run three of the four treatment combinations each day. A total of 4 days are needed for the design to be balanced. Given there are six possible runs per day, each of these combinations could be replicated or run over a total of 8 days. Since more days takes away degrees of freedom from error, the replication design will be used. The ANOVA table below is for a BIBD with k = 3.

Source	$\mathrm{DF}$
Day	3
Trt	3
Error	17
Total	23

(b) The std deviation of a treatment difference when using a RCBD is  $\sqrt{2\sigma^2/b} = \sqrt{2\sigma^2/6}$ . With 15 degrees of freedom, the t-statistic is 2.131. Thus the half-length of the confidence interval would be  $2.131\sqrt{\sigma^2/3}$ . For the BIBD above, the std deviation of a treatment difference is  $\sqrt{2k\sigma^2/(\lambda a)} = \sqrt{2(3)\sigma^2/(4(4))}$ . With 17 degrees of freedom the t statistic is 2.110 so the half length is  $2.110\sqrt{3\sigma^2/8}$ . In this situation,

$$\frac{2.131\sqrt{\sigma^2/3}}{2.110\sqrt{3\sigma^2/8}} < 1$$

so the RCBD is better. This is the best BIBD among the ones suggested so all other BIBDs are also worse. Other designs could be formed using only 4 days but they would not be balanced. If the experimenter was more interested in comparing certain treatment combinations then a RCBD could be combined with a PBIB over 4 days.

## Problem 1

Consider a BIBD where a = 4, b = 6, k = 2.

## Problem 2

(a) We have a = 5 treatments (gasoline additives) and b = 5 blocks (cars) with  $a \le b$ . Each block contains k = 4 treatments, each treatment appears in r = 4 blocks, and each pair of treatments

appears in the same blocks  $\lambda = 3$  times. The total number of runs N = ar = bk = 20, and  $\lambda(a-1) = r(k-1) = 12$ . Hence this is a balanced incomplete block design.

(b) The ANOVA table from SAS is as follows (line for Model SS replaced by Type I SS of the block and the treatment)

		Sum of			
Source	DF	Squares	Mean Square	F Value	Pr > F
car	4	31.20000000	7.8000000	8.57	0.0022
trt	4	35.73333333	8.93333333	9.81	0.0012
Error	11	10.01666667	0.91060606		
Corrected Total	19	76.95000000			

Since the *p*-value for treatment effect is very small (= 0.0012), I conclude that there is a difference between the five gasoline additives.

(c) First, the row and column sums are calculated based on the original data table.

	car1	$\operatorname{car2}$	car3	car4	$\operatorname{car5}$	$y_{i\cdot}$
add1		17	14	13	12	56
add2	14	14		13	10	51
add3	12		13	12	9	46
add4	13	11	11	12		47
add5	11	12	10		8	41
$y_{\cdot j}$	50	54	48	50	39	

And the overall mean

$$\bar{y}_{\cdot\cdot} = \frac{1}{N} \sum_{j} y_{\cdot j} = \frac{50 + 54 + 48 + 50 + 39}{20} = 12.05.$$

Second, let's compute

$$Q_i = y_{i\cdot} - \frac{1}{k} \sum_j n_{ij} y_{\cdot j}, \qquad i = 1, 2, 3, 4, 5,$$

where  $n_{ij}$  equals 1 if treatment *i* appears in block *j* and 0 otherwise.

$$Q_{1} = y_{1} - \frac{1}{4} \sum_{j=1}^{5} n_{1j} y_{\cdot j} = 56 - \frac{54 + 48 + 50 + 39}{4} = 8.25,$$

$$Q_{2} = y_{2} - \frac{1}{4} \sum_{j=1}^{5} n_{2j} y_{\cdot j} = 51 - \frac{50 + 54 + 50 + 39}{4} = 2.75,$$

$$Q_{3} = y_{3} - \frac{1}{4} \sum_{j=1}^{5} n_{3j} y_{\cdot j} = 46 - \frac{50 + 48 + 50 + 39}{4} = -0.75,$$

$$Q_{4} = y_{4} - \frac{1}{4} \sum_{j=1}^{5} n_{4j} y_{\cdot j} = 47 - \frac{50 + 54 + 48 + 50}{4} = -3.5,$$

$$Q_{5} = y_{5} - \frac{1}{4} \sum_{j=1}^{5} n_{5j} y_{\cdot j} = 41 - \frac{50 + 54 + 48 + 39}{4} = -6.75.$$

Finally, the adjusted means are calculated by

$$\hat{\mu}_i = \hat{\mu} + \hat{\tau}_i = \bar{y}_{\cdots} + \frac{kQ_i}{\lambda a}$$

Hence

$$\begin{aligned} \hat{\mu}_1 &= 12.05 + \frac{4 \cdot 8.25}{3 \cdot 5} = 14.25, \\ \hat{\mu}_2 &= 12.05 + \frac{4 \cdot 2.75}{3 \cdot 5} = 12.7833, \\ \hat{\mu}_3 &= 12.05 + \frac{4 \cdot (-0.75)}{3 \cdot 5} = 11.85, \\ \hat{\mu}_4 &= 12.05 + \frac{4 \cdot (-3.5)}{3 \cdot 5} = 11.1167, \\ \hat{\mu}_5 &= 12.05 + \frac{4 \cdot (-6.75)}{3 \cdot 5} = 10.25. \end{aligned}$$

(d) The standard error of the difference between two treatment estimates is

$$\sqrt{\hat{Var}(\hat{\tau}_i - \hat{\tau}_j)} = \sqrt{\frac{2k}{\lambda a}MS_E} = \sqrt{\frac{2\cdot 4}{3\cdot 5}\cdot 0.9106} = 0.6969.$$

(e) The critical difference for Tukey's pairwise comparisons is

$$CD = \frac{q_{\alpha,a,ar-a-b+1}}{\sqrt{2}} \sqrt{\frac{2k}{\lambda a}} MS_E = \frac{q_{0.05,5,11}}{\sqrt{2}} \cdot 0.6969 = \frac{4.58}{\sqrt{2}} \cdot 0.6969 = 2.2569$$

The five treatment mean estimates are ordered as 14.25 > 12.7833 > 11.85 > 11.1167 > 10.25 ( $\hat{\mu}_1 > \hat{\mu}_2 > \hat{\mu}_3 > \hat{\mu}_4 > \hat{\mu}_5$ ). After computing differences following this order and comparing them with the critical distance, I reach the following conclusion.

- The pairs of gasoline additives which have significantly different mileage performances are (1,3), (1,4), (1,5), (2,5).
- The mileage performances within any other pairs of gasoline additives are not significantly different.

The results from SAS with the options "lsmeans trt / tdiff adjust=tukey;" are given below.

Least Squares Means Adjustment for Multiple Comparisons: Tukey-Kramer

		LSMEAN
trt	y LSMEAN	Number
1	14.2500000	1
2	12.7833333	2
3	11.8500000	3
4	11.1166667	4
5	10.2500000	5

Least Squares Means for Effect trt t for HO: LSMean(i)=LSMean(j) / Pr > |t|

Dependent Variable: y

i/j	1	2	3	4	5
1		2.104587	3.443869	4.496162	5.739781
		0.2838	0.0355	0.0065	0.0010
2	-2.10459		1.339282	2.391576	3.635195
	0.2838		0.6746	0.1884	0.0259
3	-3.44387	-1.33928		1.052293	2.295913
	0.0355	0.6746		0.8262	0.2167
4	-4.49616	-2.39158	-1.05229		1.243619
	0.0065	0.1884	0.8262		0.7280
5	-5.73978	-3.63519	-2.29591	-1.24362	
	0.0010	0.0259	0.2167	0.7280	

In the SAS output, the pairs that are significant different have *p*-values (the bottom values for entries in the output table) less than 0.05. Hence, the significantly different pairs from SAS are (1,3), (1,4), (1,5), (2,5), which are consistent with my conclusion.

(f) The contrast I use is  $\mathbf{C} = (1, 1, 0, -1, -1)'$ . The significance test of it is shown below.

Contrast	DF	Contrast SS	Mean Square	F Value	Pr > F
C	1	30.10416667	30.10416667	33.06	0.0001

Since the *p*-value for the test is very small (= 0.0001), I conclude that the combination of additives 1 and 2 has significantly different characteristics from the combination of additives 4 and 5.