

## Minimum Mean Square Error Forecasts

Consider the forecasting, or prediction, of  $z_{t+l}$  given  $z_t, z_{t-1}, \dots$ , assuming  $\{z_t\}$  follows some *known* ARIMA( $p, d, q$ ) model.

Let  $\tilde{z}_t(l)$  be any function of  $Z_t = (z_t, z_{t-1}, \dots)$ . The mean square error of forecasting  $z_{t+l}$  by  $\tilde{z}_t(l)$  is seen to satisfy

$$\begin{aligned} E[z_{t+l} - \tilde{z}_t(l)]^2 &= E[z_{t+l} - E[z_{t+l}|Z_t]]^2 + E[E[z_{t+l}|Z_t] - \tilde{z}_t(l)]^2 \\ &\geq E[z_{t+l} - E[z_{t+l}|Z_t]]^2. \end{aligned}$$

Hence,  $E[z_{t+l}|Z_t]$  is the minimum mean square error forecast of  $z_{t+l}$  given  $Z_t$ , to be denoted by  $\hat{z}_t(l)$ .

Recall the truncated MA form  $z_{t+l} = I_t(l) + C_t(l)$ , where  $I_t(l) = \sum_{j=0}^{l-1} \psi_j a_{t+l-j}$  and  $C_t(l)$  is the complimentary function at origin  $t$ . It is easily seen that  $\hat{z}_t(l) = C_t(l)$ .  $I_t(l)$  is the forecasting error, to be denoted by  $e_t(l)$ .

## Other Forms of Forecasts, Updating

Consider the difference equation form of the model

$$z_{t+l} = \sum_{j=1}^{p+d} \varphi_j z_{t+l-j} + a_{t+l} - \sum_{j=1}^q \theta_j a_{t+l-j}.$$

Taking conditional expectations at time  $t$ , one has

$$\hat{z}_t(l) = \sum_{j=1}^{p+d} \varphi_j \hat{z}_t(l-j) - \sum_{j=l}^q \theta_j a_{t+l-j},$$

where  $\hat{z}_t(k) = z_{t+k}$  for  $k \leq 0$ .

Based on the AR form of the model,  $z_{t+l} = \sum_{j=1}^{\infty} \pi_j z_{t+l-j} + a_{t+l}$ , one has  $\hat{z}_t(l) = \sum_{j=1}^{\infty} \pi_j \hat{z}_t(l-j)$ , where  $\hat{z}_t(k) = z_{t+k}$  for  $k \leq 0$ .

From  $z_{t+l} = \sum_{j=0}^{l-1} \psi_j a_{t+l-j} + \hat{z}_t(l)$ , it is easy to show that

$$\hat{z}_{t+1}(l-1) = \hat{z}_t(l) + \psi_{l-1} a_{t+1}.$$

As soon as  $z_{t+1}$  becomes available, one may calculate

$a_{t+1} = z_{t+1} - \hat{z}_t(1)$  and update the forecast of  $z_{t+l}$  by  $\hat{z}_{t+1}(l-1)$ .

## Examples

Consider an ARI(1,1,0) model  $(1 - 1.8B + .8B^2)z_t = a_t$ . One has

$$\hat{z}_t(1) = 1.8z_t - .8z_{t-1},$$

$$\hat{z}_t(2) = 1.8\hat{z}_t(1) - .8z_t,$$

$$\hat{z}_t(l) = 1.8\hat{z}_t(l-1) - .8\hat{z}_t(l-2), \quad l > 2.$$

The  $\psi$  weights for updating are given by  $\psi_j = 1.8\psi_{j-1} - .8\psi_{j-2}$ ,  $j > 0$ , with  $\psi_0 = 1$ ,  $\psi_{-1} = 0$ .

Consider an IMA(0,2,2) model  $\nabla^2 z_t = (1 - .9B + .5B^2)a_t$ . One has

$$\hat{z}_t(1) = 2z_t - 1z_{t-1} - .9a_t + .5a_{t-1},$$

$$\hat{z}_t(2) = 2\hat{z}_t(1) - z_t + .5a_t,$$

$$\hat{z}_t(l) = 2\hat{z}_t(l-1) - \hat{z}_t(l-2), \quad l > 2.$$

The  $\psi$  weights are given by  $\psi_j = 2\psi_{j-1} - \psi_{j-2}$ ,  $j > 2$ , with  $\psi_0 = 1$ ,  $\psi_1 = 2 - .9 = 1.1$ ,  $\psi_2 = 2(1.1) - 1 + .5 = 1.7$ .

## Probability Limits of Forecasts

The forecasting error  $e_t(l) = \sum_{j=0}^{l-1} \psi_j a_{t+l-j}$  has variance

$$V(l) = (1 + \sum_{j=1}^{l-1} \psi_j^2) \sigma_a^2,$$

which naturally increases with  $l$ . The formula can be used to calculate “prediction intervals” for  $z_{t+l}$ ,

$$\hat{z}_t(l) \pm 1.96 \sqrt{(1 + \sum_{j=1}^{l-1} \psi_j^2) \sigma_a^2}.$$

For the examples above,  $(1 - 1.8B + .8B^2)z_t = a_t$  (model A) and  $\nabla^2 z_t = (1 - .9B + .5B^2)a_t$  (model B),  $\sqrt{V(l)/\sigma_a^2}$  at  $l = 1, \dots, 6$  are calculated and listed below.

$l$	1	2	3	4	5	6
A	1	2.06	3.19	4.35	5.50	6.62
B	1	1.49	2.26	3.22	4.34	5.57

## Forecast Function and Weights

For  $l > q$ ,  $\hat{z}_t(l) = \sum_{j=1}^{p+d} \varphi_j \hat{z}_t(l-j)$ , so the “eventual” forecasting function satisfies the equation  $\varphi(B)\hat{z}_t(l) = 0$ , hence are of the form

$$\hat{z}_t(l) = b_0^{(t)} f_0(l) + \cdots + b_{p+d-1}^{(t)} f_{p+d-1}(l),$$

where  $f_j(l)$  are determined by the roots of  $\varphi(B)$  and  $b_j^{(t)}$  by the initial values. Recall the form of complimentary function  $C_t(l)$ .

From  $\hat{z}_t(l) = \sum_{j=1}^{\infty} \pi_j \hat{z}_t(l-j)$ , one may express  $\hat{z}_t(l)$  directly in terms of  $z_t, z_{t-1}, \dots$ ,  $\hat{z}_t(l) = \sum_{j=1}^{\infty} \pi_j^{(l)} z_{t-j+1}$ .

$$\hat{z}_t(1) = \pi_1 z_t + \pi_2 z_{t-1} + \pi_3 z_{t-2} + \cdots$$

$$\hat{z}_t(2) = \pi_1 \hat{z}_t(1) + \pi_2 z_t + \pi_3 z_{t-1} + \cdots$$

$$= (\pi_1 \pi_1 + \pi_2) z_t + (\pi_1 \pi_2 + \pi_3) z_{t-1} + \cdots$$

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## General Form of Forecast Weights

From  $z_{t+l} = \sum_{k=0}^{l-1} \psi_k a_{t+l-k} + \hat{z}_t(l)$ , one has

$$\begin{aligned} \hat{z}_t(l) &= \hat{z}_{t+l-1}(1) - \sum_{k=1}^{l-1} \psi_k a_{t-l-k} \\ &= \pi_1 z_{t+l-1} + \pi_2 z_{t+l-2} + \cdots + \pi_{l-1} z_{t+1} + \pi_l z_t + \pi_{l+1} z_{t-1} + \cdots \\ &\quad + \psi_1 (-z_{t+l-1} + \pi_1 z_{t+l-2} + \cdots + \pi_{l-2} z_{t+1} + \pi_{l-1} z_t + \pi_l z_{t-1} + \cdots) \\ &\quad + \cdots + \psi_{l-1} (-z_{t+1} + \pi_1 z_t + \pi_2 z_{t-1} + \cdots). \end{aligned}$$

Adding up the coefficients of  $z_t, z_{t-1}, \dots$ , one has

$$\pi_j^{(l)} = \pi_{l+j-1} + \psi_1 \pi_{l+j-2} + \cdots + \psi_{l-1} \pi_j = \pi_{j+1}^{(l-1)} + \psi_{l-1} \pi_j.$$

For example,  $\pi_j^{(2)} = \pi_{j+1} + \psi_1 \pi_j$ ,

$$\pi_j^{(3)} = \pi_{j+2} + \psi_1 \pi_{j+1} + \psi_2 \pi_j = \pi_{j+1}^{(2)} + \psi_2 \pi_j.$$

The coefficients of  $z_{t+l-1}, \dots, z_{t+1}$  vanish as  $\sum_j \pi_j \psi_{k-j} = 0$  for  $k > 0$ , where  $\psi_0 = -\pi_0 = 1$ ,  $\psi_j = \pi_j = 0$ ,  $j < 0$ .

### Example: IMA(0,1,1)

Consider the model  $\nabla z_t = a_t - \theta a_{t-1}$ . One has

$$\hat{z}_t(1) = z_t - \theta a_t, \quad \hat{z}_t(l) = \hat{z}_t(l-1) = \hat{z}_t(1), \quad l > 1,$$

which give a constant forecast function. Since  $\psi_j = 1 - \theta$ ,  $j > 0$ , the forecast function can be updated through

$$\hat{z}_{t+1}(l) = \hat{z}_{t+1}(l-1) = \hat{z}_t(l) + (1 - \theta)a_{t+1}, \quad l > 1.$$

The  $\pi$  weights are  $\pi_j = (1 - \theta)\theta^{j-1}$ . Note that

$$\pi_j^{(2)} = \pi_{j+1} + \psi_1 \pi_j = (1 - \theta)\theta^j + (1 - \theta)^2 \theta^{j-1} = \pi_j,$$

so there is no surprise here. The calculation applies recursively to  $\pi_j^{(3)}, \pi_j^{(4)}, \dots$ . The variance of  $e_t(l)$  is easily seen to be

$$V(l) = \sigma_a^2(1 + (l-1)(1-\theta)^2).$$

### Example: IMA(0,2,2)

Consider the model  $\nabla^2 z_t = a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2}$ . One has

$$\hat{z}_t(1) = 2z_t - 1z_{t-1} - \theta_1 a_t - \theta_2 a_{t-1},$$

$$\hat{z}_t(2) = 2\hat{z}_t(1) - z_t - \theta_2 a_t,$$

$$\hat{z}_t(l) = 2\hat{z}_t(l-1) - \hat{z}_t(l-2), \quad l > 2.$$

Since  $\psi_j = \lambda_0 + \lambda_1 j$ , where  $\lambda_0 = 1 + \theta_2$  and  $\lambda_1 = 1 - \theta_1 - \theta_2$ ,

$$\hat{z}_{t+1}(l-1) = \hat{z}_t(l) + (\lambda_0 + \lambda_1(l-1))a_{t+1}, \quad l > 1.$$

In the form of the complimentary function,  $\hat{z}_t(l) = b_0^{(t)} + b_1^{(t)}l$ , and

$$b_0^{(t+1)} = b_0^{(t)} + b_1^{(t)} + \lambda_0 a_{t+1}, \quad b_1^{(t+1)} = b_1^{(t)} + \lambda_1 a_{t+1}.$$

The variance of  $e_t(l)$  is given by

$$V(l) = \sigma_a^2(1 + (l-1)\lambda_0^2 + \frac{1}{6}l(l-1)(2l-1)\lambda_1^2 + \lambda_0\lambda_1l(l-1)).$$



## Examples: AR(1) and ARI(1,1,0)

Consider the model  $z_t = \phi z_{t-1} + a_t$ . One has  $\hat{z}_t(l) = z_t \phi^l$ . Since  $\psi_j = \phi^j$ , the variance of  $e_t(l)$  is given by

$$V(l) = \sigma_a^2 \sum_{j=0}^{l-1} \phi^{2j} = \sigma_a^2 (1 - \phi^{2l}) / (1 - \phi^2).$$

Consider the model  $(1 - \phi B)(1 - B)z_t = a_t$ . One has

$$\hat{z}_t(l) - \hat{z}_t(l-1) = \phi^l (z_t - z_{t-1}).$$

It follows that  $\hat{z}_t(l) = z_t + (\sum_{j=1}^l \phi^j)(z_t - z_{t-1})$ ,  $l > 0$ , or

$$\hat{z}_t(l) = z_t + (z_t - z_{t-1})\phi(1 - \phi^l)/(1 - \phi),$$

which “converges” to  $z_t + (z_t - z_{t-1})\phi/(1 - \phi)$ . It can be shown that  $\psi_j = (1 - \phi^{j+1})/(1 - \phi)$ , so the variance of  $e_t(l)$  is given by

$$V(l) = \frac{\sigma_a^2}{(1 - \phi)^2} \left\{ l + \frac{\phi^2(1 - \phi^{2l})}{1 - \phi^2} - 2 \frac{\phi(1 - \phi^l)}{1 - \phi} \right\}.$$

## Example: ARMA(1,1)

Consider the model  $z_t = \phi z_{t-1} + a_t - \theta a_{t-1}$ . One has

$$\hat{z}_t(1) = \phi z_t - \theta a_t,$$

$$\hat{z}_t(l) = \phi \hat{z}_t(l-1) = \phi^{l-1} \hat{z}_t(1), \quad l > 1.$$

Since  $\psi_j = (\phi - \theta)\phi^{j-1}$ ,  $j > 0$ , so for  $l > 1$ ,

$$\hat{z}_{t+1}(l-1) = \hat{z}_t(l) + (\phi - \theta)\phi^{l-2} a_{t+1}.$$

In particular, one has the updating formula

$$\hat{z}_{t+1}(1) = \hat{z}_t(2) + (\phi - \theta)a_{t+1} = \phi \hat{z}_t(1) + (\phi - \theta)a_{t+1}.$$

The  $\pi$  weights are  $\pi_j = (\phi - \theta)\theta^{j-1}$ , so  $\pi_j^{(l)} = \phi^{l-1}(\phi - \theta)\theta^{j-1}$ . The variance of  $e_t(l)$  is seen to be

$$V(l) = \sigma_a^2 \{1 + (\phi - \theta)^2 (1 - \phi^{2(l-1)}) / (1 - \phi^2)\}.$$

## Example: ARIMA(1,1,1)

Consider the model  $(1 - \phi B)(1 - B)z_t = a_t - \theta a_{t-1}$ . One has

$$\hat{z}_t(1) = (1 + \phi)z_t - \phi z_{t-1} - \theta a_t,$$

$$\hat{z}_t(l) = (1 + \phi)\hat{z}_t(l-1) - \phi\hat{z}_t(l-2), \quad l > 1.$$

Since  $(1 - \phi B)(1 - B)\hat{z}_t(l) = 0$ ,  $l > 1$ ,  $\hat{z}_t(l) = b_0^{(t)} + b_1^{(t)}\phi^l$ , where

$$b_0^{(t)} = (z_t - \phi z_{t-1} - \theta a_t)/(1 - \phi) = z_t - b_1^{(t)},$$

$$b_1^{(t)} = (\theta a_t - \phi(z_t - z_{t-1}))/ (1 - \phi),$$

yielding  $\hat{z}_t(l) = z_t + \phi \frac{1-\phi^l}{1-\phi} (z_t - z_{t-1}) - \theta \frac{1-\phi^l}{1-\phi} a_t$ . As  $a_t = z_t - \sum_{j=1}^{\infty} \pi_j z_{t-j}$ , where  $\pi_1 = 1 + \phi - \theta$ ,  $\pi_j = (1 - \theta)(\theta - \phi)\theta^{j-2}$ ,  $j > 1$ , some algebra yields

$$\hat{z}_t(l) = (1 - \alpha_l)z_t + \alpha_l \left\{ (1 - \theta) \sum_{j=1}^{\infty} \theta^{j-1} z_{t-j} \right\},$$

where  $\alpha_l = (\theta - \phi)(1 - \phi^l)/(1 - \phi)$ . Now  $\psi_j = [(1 - \theta) + \phi^l(\theta - \phi)]/(1 - \phi)$ , so

$$V(l) = \frac{\sigma_a^2}{(1 - \phi)^2} \left\{ l(1 - \theta)^2 + (\theta - \phi)^2 \frac{1 - \phi^{2l}}{1 - \phi^2} + 2(1 - \theta)(\theta - \phi) \frac{1 - \phi^l}{1 - \phi} \right\}.$$

## Forecasting with Finite Samples

With finite samples  $(z_t, z_{t-1}, \dots, z_1)$ , the procedure developed above works without a problem for  $q = 0$ . For  $q > 0$ , however,  $a_t, \dots, a_{t-q+1}$  appearing in  $\hat{z}_t(l)$  also depend on  $z_{-1}, z_{-2}, \dots$ , so modifications are needed.

For invertible models, the  $\pi$  weights decay exponentially, so it is reasonable to set  $z_{-1} = z_{-2} = \dots = 0$  when calculating  $a_k$  from  $z_k, z_{k-1}, \dots$ .

Using the innovations algorithm, one can calculate the exact one step forecast  $\hat{z}_{t+1} = E[z_{t+1} | z_t, \dots, z_1]$  with error variance  $v_t$ . The exact multiple step forecast with finite samples will be discussed along with the state space models.