

## **Lecture 14: Blocking and Confounding in $2^k$ design**

Montgomery: Chapter 7

## Randomized Complete Block $2^k$ Design

- There are  $n$  blocks
- Within each block, all treatments (level combinations) are conducted.
- Run order in each block must be randomized
- Analysis follows general block factorial design
- When  $k$  is large, cannot afford to conduct all the treatments within each block. Other blocking strategy should be considered.

### Filtration Rate Experiment (revisited)

| factor   |          |          |          | original response |
|----------|----------|----------|----------|-------------------|
| <i>A</i> | <i>B</i> | <i>C</i> | <i>D</i> |                   |
| -        | -        | -        | -        | 45                |
| +        | -        | -        | -        | 71                |
| -        | +        | -        | -        | 48                |
| +        | +        | -        | -        | 65                |
| -        | -        | +        | -        | 68                |
| +        | -        | +        | -        | 60                |
| -        | +        | +        | -        | 80                |
| +        | +        | +        | -        | 65                |
| -        | -        | -        | +        | 43                |
| +        | -        | -        | +        | 100               |
| -        | +        | -        | +        | 45                |
| +        | +        | -        | +        | 104               |
| -        | -        | +        | +        | 75                |
| +        | -        | +        | +        | 86                |
| -        | +        | +        | +        | 70                |
| +        | +        | +        | +        | 96                |

- Suppose there are two batches of raw material. Each batch can be used for only 8 runs. It is known these two batches are very different. Blocking should be employed to eliminate this variability.
- How to select 8 treatments (level combinations, or runs) for each block?

## $2^2$ Design with Two Blocks

Suppose there are two factors ( $A, B$ ) each with 2 levels, and two blocks ( $b_1, b_2$ ) each containing two runs (treatments). Since  $b_1$  and  $b_2$  are interchangeable, there are three possible blocking scheme:

| $A$ | $B$ | response | blocking scheme |       |       |
|-----|-----|----------|-----------------|-------|-------|
|     |     |          | 1               | 2     | 3     |
| -   | -   | $y_{--}$ | $b_1$           | $b_1$ | $b_2$ |
| +   | -   | $y_{+-}$ | $b_1$           | $b_2$ | $b_1$ |
| -   | +   | $y_{-+}$ | $b_2$           | $b_1$ | $b_1$ |
| +   | +   | $y_{++}$ | $b_2$           | $b_2$ | $b_2$ |

Comparing blocking schemes:

Scheme 1:

- block effect:  $b = \bar{y}_{b_2} - \bar{y}_{b_1} = \frac{1}{2}(-y_{--} - y_{+-} + y_{-+} + y_{++})$
- main effect:  $B = \frac{1}{2}(-y_{--} - y_{+-} + y_{-+} + y_{++})$
- $B$  and  $b$  are not distinguishable, or, confounded.

## Comparing Blocking Schemes (continued)

Scheme 2:

$$\text{block effect: } b = \bar{y}_{b_2} - \bar{y}_{b_1} = \frac{1}{2}(-y_{--} + y_{+-} - y_{-+} + y_{++})$$

$$\text{main effect: } A = \frac{1}{2}(-y_{--} + y_{+-} - y_{-+} + y_{++})$$

$A$  and  $b$  are not distinguishable, or confounded.

Scheme 3:

$$\text{block effect: } b = \bar{y}_{b_2} - \bar{y}_{b_1} = \frac{1}{2}(y_{--} - y_{+-} - y_{-+} + y_{++})$$

$$\text{interaction: } AB = \frac{1}{2}(y_{--} - y_{+-} - y_{-+} + y_{++})$$

$AB$  and  $b$  become indistinguishable, or confounded.

The reason for confounding: the block arrangement matches the contrast of some factorial

effect.

Confounding makes the effect **Inestimable**.

**Question: which scheme is the best (or causes the least damage)?**

## $2^k$ Design with Two Blocks via Confounding

Confound blocks with the effect (contrast) of the highest order

Block 1 consists of all treatments with the contrast coefficient equal to -1

Block 2 consists of all treatments with the contrast coefficient equal to 1

Example 1. Block  $2^3$  Design

| factorial effects (contrasts) |    |    |    |    |    |    |     |
|-------------------------------|----|----|----|----|----|----|-----|
| I                             | A  | B  | C  | AB | AC | BC | ABC |
| 1                             | -1 | -1 | -1 | 1  | 1  | 1  | -1  |
| 1                             | 1  | -1 | -1 | -1 | -1 | 1  | 1   |
| 1                             | -1 | 1  | -1 | -1 | 1  | -1 | 1   |
| 1                             | 1  | 1  | -1 | 1  | -1 | -1 | -1  |
| 1                             | -1 | -1 | 1  | 1  | -1 | -1 | 1   |
| 1                             | 1  | -1 | 1  | -1 | 1  | -1 | -1  |
| 1                             | -1 | 1  | 1  | -1 | -1 | 1  | -1  |
| 1                             | 1  | 1  | 1  | 1  | 1  | 1  | 1   |

Defining relation:  $b = ABC$ :

Block 1:  $(- - -), (+ + -), (+ - +), (- + +)$

Block 2:  $(+ - -), (- + -), (- - +), (+ + +)$

Example 2: For  $2^4$  design with factors:  $A, B, C, D$ , the defining contrast

(optimal) for blocking factor ( $b$ ) is

$$b = ABCD$$

In general, the optimal blocking scheme for  $2^k$  design with two blocks is given by  $b = AB \dots K$ , where  $A, B, \dots, K$  are the factors.

## Analyze Unreplicated Block $2^k$ Experiment

Filtration Experiment (four factors:  $A, B, C, D$ ):

- Use defining relation:  $b = ABCD$ , i.e., if a treatment satisfies  $ABCD = -1$ , it is allocated to block 1 ( $b_1$ ); if  $ABCD = 1$ , it is allocated to block 2 ( $b_2$ ).
- (Assume that, all the observations in block 2 will be reduced by 20 because of the poor quality of the second batch of material, i.e. the true block effect=-20).

| factor   |          |          |          | blocks     | response    |
|----------|----------|----------|----------|------------|-------------|
| <i>A</i> | <i>B</i> | <i>C</i> | <i>D</i> | $b = ABCD$ |             |
| -        | -        | -        | -        | $1=b_2$    | $45-20=25$  |
| +        | -        | -        | -        | $-1=b_1$   | 71          |
| -        | +        | -        | -        | $-1=b_1$   | 48          |
| +        | +        | -        | -        | $1=b_2$    | $65-20=45$  |
| -        | -        | +        | -        | $-1=b_1$   | 68          |
| +        | -        | +        | -        | $1=b_2$    | $60-20=40$  |
| -        | +        | +        | -        | $1=b_2$    | $80-20=60$  |
| +        | +        | +        | -        | $-1=b_1$   | 65          |
| -        | -        | -        | +        | $-1=b_1$   | 43          |
| +        | -        | -        | +        | $1=b_2$    | $100-20=80$ |
| -        | +        | -        | +        | $1=b_2$    | $45-20=25$  |
| +        | +        | -        | +        | $-1=b_1$   | 104         |
| -        | -        | +        | +        | $1=b_2$    | $75-20=55$  |
| +        | -        | +        | +        | $-1=b_1$   | 86          |
| -        | +        | +        | +        | $-1=b_1$   | 70          |
| +        | +        | +        | +        | $1=b_2$    | $96-20=76$  |

## SAS File for Block Filtration Experiment

```
goption colors=(none);
data filter;
  do D = -1 to 1 by 2;do C = -1 to 1 by 2;
  do B = -1 to 1 by 2;do A = -1 to 1 by 2;
  input y @@;  output;
  end; end; end; end;
cards;
25 71 48 45 68 40 60 65 43 80 25 104 55 86 70 76
;

data inter;
set filter; AB=A*B; AC=A*C; AD=A*D; BC=B*C; BD=B*D; CD=C*D; ABC=AB*C;
ABD=AB*D; ACD=AC*D; BCD=BC*D; block=ABC*D;

proc glm data=inter;
class A B C D AB AC AD BC BD CD ABC ABD ACD BCD block;
model y=block A B C D AB AC AD BC BD CD ABC ABD ACD BCD; run;

proc reg outest=effects data=inter;
```

```
model y=A B C D AB AC AD BC BD CD ABC ABD ACD BCD block;
data effect2; set effects; drop y intercept _RMSE_;
proc transpose data=effect2 out=effect3;
data effect4; set effect3; effect=col1*2;
proc sort data=effect4; by effect;
proc print data=effect4;

data effect5; set effect4; where _NAME_ ^= 'block';
proc print data=effect5; run;

proc rank data=effect5 normal=blom;
var effect; ranks neff;

symbol1 v=circle;
proc gplot; plot effect*neff=_NAME_; run;
```

### SAS output: ANOVA Table

| Source   | DF | Squares     | Mean Square | F Value | Pr > F |
|----------|----|-------------|-------------|---------|--------|
| Model    | 15 | 7110.937500 | 474.062500  | .       | .      |
| Error    |    | 0           | 0.000000    | .       | .      |
| Co Total | 15 | 7110.937500 |             |         |        |

| Source | DF | Type I SS   | Mean Square | F Value | Pr > F |
|--------|----|-------------|-------------|---------|--------|
| block  | 1  | 1387.562500 | 1387.562500 | .       | .      |
| A      | 1  | 1870.562500 | 1870.562500 | .       | .      |
| B      | 1  | 39.062500   | 39.062500   | .       | .      |
| C      | 1  | 390.062500  | 390.062500  | .       | .      |
| D      | 1  | 855.562500  | 855.562500  | .       | .      |
| AB     | 1  | 0.062500    | 0.062500    | .       | .      |
| AC     | 1  | 1314.062500 | 1314.062500 | .       | .      |
| AD     | 1  | 1105.562500 | 1105.562500 | .       | .      |
| BC     | 1  | 22.562500   | 22.562500   | .       | .      |
| BD     | 1  | 0.562500    | 0.562500    | .       | .      |
| CD     | 1  | 5.062500    | 5.062500    | .       | .      |
| ABC    | 1  | 14.062500   | 14.062500   | .       | .      |

|     |   |           |           |   |   |
|-----|---|-----------|-----------|---|---|
| ABD | 1 | 68.062500 | 68.062500 | . | . |
| ACD | 1 | 10.562500 | 10.562500 | . | . |
| BCD | 1 | 27.562500 | 27.562500 |   |   |

proportion of variance explained by blocks

$$\frac{1387.5625}{7110.9375} = 19.5\%$$

Similarly proportion of variance can be calculated for other effects.

### SAS output: factorial effects and block effect

| Obs | _NAME_ | COL1    | effect  |
|-----|--------|---------|---------|
| 1   | block  | -9.3125 | -18.625 |
| 2   | AC     | -9.0625 | -18.125 |
| 3   | BCD    | -1.3125 | -2.625  |
| 4   | ACD    | -0.8125 | -1.625  |
| 5   | CD     | -0.5625 | -1.125  |
| 6   | BD     | -0.1875 | -0.375  |
| 7   | AB     | 0.0625  | 0.125   |
| 8   | ABC    | 0.9375  | 1.875   |
| 9   | BC     | 1.1875  | 2.375   |
| 10  | B      | 1.5625  | 3.125   |
| 11  | ABD    | 2.0625  | 4.125   |
| 12  | C      | 4.9375  | 9.875   |
| 13  | D      | 7.3125  | 14.625  |
| 14  | AD     | 8.3125  | 16.625  |
| 15  | A      | 10.8125 | 21.625  |

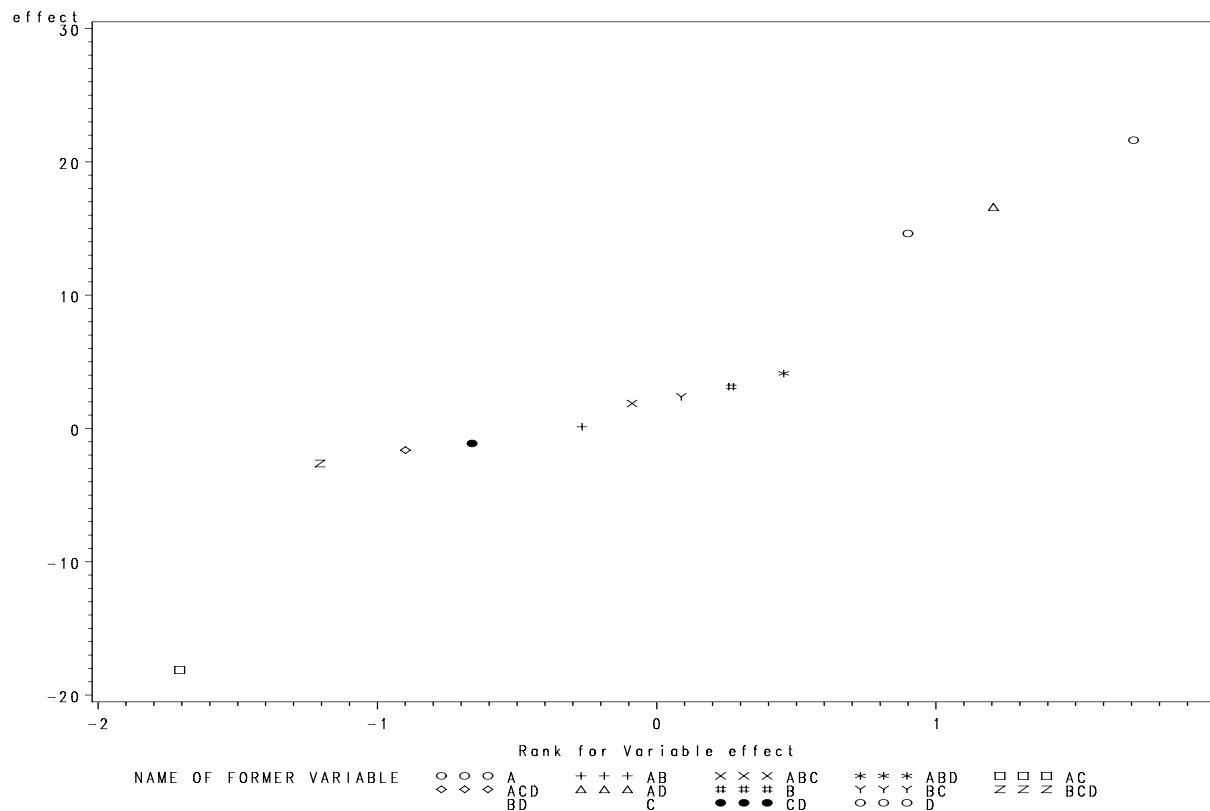
Factorial effects are exactly the same as those from the original data (why?)

blocking effect:  $-18.625 = \bar{y}_{b_2} - \bar{y}_{b_1}$ , is in fact

$$-20(\text{true blocking effect}) + 1.375(\text{some interaction of } ABCD)$$

This is caused by confounding between  $b$  and  $ABCD$ .

## SAS output: QQ plot without Blocking Effect



significant effects are:

*A, C, D, AC, AD*

## $2^k$ Design with Four Blocks

Need two 2-level blocking factors to generate 4 different blocks.  
 Confound each blocking factors with a high order factorial effect.  
 The interaction between these two blocking factors matters.  
 The interaction will be confounded with another factorial effect.

Optimal blocking scheme has least confounding severity.

$2^4$  design with four blocks: factors are  $A, B, C, D$  and the blocking factors are  $b1$  and  $b2$

| A     | B  | C  | D  | AB | AC | .....CD | ABC | ABD | ACD | BCD | ABCD |    |    |        |
|-------|----|----|----|----|----|---------|-----|-----|-----|-----|------|----|----|--------|
| -1    | -1 | -1 | -1 | 1  | 1  | 1       | -1  | -1  | -1  | -1  | 1    |    |    |        |
| 1     | -1 | -1 | -1 | -1 | -1 | 1       | 1   | 1   | 1   | -1  | -1   | b1 | b2 | blocks |
| -1    | 1  | -1 | -1 | -1 | 1  | 1       | 1   | 1   | -1  | 1   | -1   | -1 | -1 | 1      |
| 1     | 1  | -1 | -1 | 1  | -1 | 1       | -1  | -1  | 1   | 1   | 1    | 1  | -1 | 2      |
|       |    |    |    |    |    |         |     |     |     |     |      | -1 | 1  | 3      |
|       |    |    |    |    |    |         |     |     |     |     |      | 1  | 1  | 4      |
| ..... |    |    |    |    |    |         |     |     |     |     |      |    |    |        |
| -1    | -1 | 1  | 1  | 1  | -1 | 1       | 1   | 1   | -1  | -1  | 1    |    |    |        |
| 1     | 1  | 1  | 1  | -1 | 1  | 1       | -1  | -1  | 1   | -1  | -1   |    |    |        |
| -1    | -1 | 1  | 1  | -1 | -1 | 1       | -1  | -1  | -1  | 1   | -1   |    |    |        |
| 1     | 1  | 1  | 1  | 1  | 1  | 1       | 1   | 1   | 1   | 1   | 1    |    |    |        |

possible blocking schemes:

Scheme 1:

defining relations:  $b_1 = ABC$ ,  $b_2 = ACD$ ; induce confounding

$$b_1 b_2 = ABC * ACD = A^2 BC^2 D = BD$$

Scheme 2:

Defining relations:  $b_1 = ABCD$ ,  $b_2 = ABC$ , induce confounding

$$b_1 b_2 = ABCD * ABC = D$$

**Which is better?**

## $2^k$ Design with $2^p$ Blocks

- $k$  factors:  $A, B, \dots, K$ , and  $p$  is usually much less than  $k$ .
- $p$  blocking factors:  $b_1, b_2, \dots, b_p$  with levels -1 and 1
- confound blocking factors with  $k$  chosen high-order factorial effects, i.e.,  $b_1 = \text{effect}_1$ ,  $b_2 = \text{effect}_2$ , etc. ( $p$  defining relations)
- These  $p$  defining relations induce another  $2^p - p - 1$  confounding.
- treatment combinations with the same values of  $b_1, \dots, b_p$  are allocated to the same block. Within each block.
- each block consists of  $2^{k-p}$  treatment combinations (runs)
- Given  $k$  and  $p$ , optimal schemes are tabulated, e.g., Montgomery Table 7.9, or Wu&Hamada Appendix 3A