# Deep Kernel Posterior Learning under Infinite Variance Prior Weights

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### Overview

- Neal's (1996) foundational PhD thesis established the infinitely wide limit of a shallow Bayesian neural network is a Gaussian process (GP) under finite variance prior weights.
- Proof is an application of the classical central limit theorem (CLT).
- This talk considers infinite variance priors (examples: Cauchy, horseshoe, ...). Classical CLT breaks. Limit is not a GP.
- Questions/Goals: (1) Is this unbounded variance regime interesting?
  (2) If yes, provide an approach for posterior inference and UQ.
- Joint work with Jorge Loría (Aalto University).

### Wide limit of a shallow (one hidden layer) BNN

Define an L layer feedforward deep neural network (DNN) with L - 1 hidden layers by the recursion:

$$\begin{split} f_j^{(\ell+1)}(\mathbf{x}) &= g\left(b_j^{(\ell)} + \sum_{i=1}^{M_\ell} w_{ij}^{(\ell)} f_i^{(\ell)}(\mathbf{x})\right),\\ \psi(\mathbf{x}) &= \sum_{j=1}^{M_L} w_j^{(L)} f_j^{(L)}(\mathbf{x}), \end{split}$$

where  $g(\cdot)$  is a nonlinear activation and  $M_{\ell}$  is width of the  $\ell$ -th layer.

- Consider a shallow net (L = 2) and let  $w_j^{(2)} \stackrel{ind}{\sim} \mathcal{N}(0, 1/M_2)$ ,
- By "self-similarity" of normal (or more generally, by classical CLT), Neal (1996) established: the limit is a GP as  $M_2 \rightarrow \infty$ .

### Covariance of Neal's GP: kernel methods and deep BNNs

- The kernel of Neal's shallow GP depends on the activation  $g(\cdot)$ .
- Neal worked out two of these explicitly in his PhD thesis:
  - g(x) = 1(x > 0) leads to exponential (Matérn with v = 1/2) kernel (very rough GP).
  - g(x) = tanh(x) leads to squared exponential (Matérn with  $\nu \to \infty$ ) kernel (infinitely smooth GP).
- Cho and Saul (2009, NeurIPS) derived using the kernel trick a recursive kernel formula for deep GPs under ReLU and other activations, of the form:  $g_{\delta}(x) = x^{\delta} \mathbf{1}(x > 0)$ .

### Implications and extensions of Neal's (1996) result

- Neal's (1996) result is attractive, because posterior inference and uncertainty quantification are straightforward for a GP, unlike a finite-width BNN.
- Using Cho and Saul (2009), extensions to deep feedforward BNNs are by Lee et al. (2018, ICLR), de G. Matthews et al. (2018, ICLR).
- Also using Cho and Saul (2009), extensions to deep convolutional BNNs are by Garriga-Alonso et al. (2018, ICLR).
- In fact, the "Tensor Program" framework of Yang (2019, NeurIPS) establishes a GP limit under nearly arbitrary architectures.
- All of the above assume finite variance priors.

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### Infinite variance priors

• The main idea in all of the above is this: if  $w_j^{(L)} \stackrel{ind}{\sim} \mathcal{N}(0, 1/M_L)$ , then:

$$\psi(\mathbf{x}) = \sum_{j=1}^{M_L} w_j^{(L)} f_j^{(L)}(\mathbf{x}) \xrightarrow{D} \mathcal{N}(0, \mathbb{E}[f(\mathbf{x})f'(\mathbf{x})]).$$

- If w<sub>j</sub><sup>(2)</sup> has unbounded variance, the generalized CLT (Gnedenko and Kolmogorov, 1953) establishes an α-stable scaling limit under relatively mild conditions.
- Der and Lee (2005, NeurIPS) used the GCLT to work out an α-stable wide limit for shallow BNNs.

### Difficulties with $\alpha$ -stable limits

- Unfortunately, it is much harder to work with α-stable variables/processes for inferential purposes.
- Mean and covariance functions are in general not available.
- Define Z ~ S<sub>α</sub>(Σ) an symmetric α-stable with scale matrix Σ; all we have is the characteristic function:

$$\phi_{\mathsf{Z}}(t) = \mathbb{E}[\exp(\mathrm{i} t^{\mathsf{T}} \mathsf{Z})] = \exp\{-(t^{\mathsf{T}} \Sigma t)^{\alpha/2}\},\$$

not a closed form density.

- Some recent attempts: Favaro et al. (2023, Bernoulli), Peluchetti et al. (2020, AISTATS), Lee et al. (2023, JMLR) etc.
- Mostly concerned with the properties of the limiting process, rather than posterior inference, analogous to kriging in GP.

## A conditional GP representation

- Our essential idea is to write the (marginal) stable process as a (conditional) GP.
- We can exploit West (1987):

 $\mathbf{Z} \sim S_{lpha}(\mathbf{\Sigma}) \Longleftrightarrow \mathbf{Z} \stackrel{D}{=} S^{1/2}_{+} \mathbf{X}, \quad S_{+} \sim S^{+}_{lpha/2}, \ \mathbf{X} \sim \mathcal{N}(0, \mathbf{\Sigma}), \ S_{+} \perp \mathbf{X},$ 

where 
$$S^+_{\alpha/2}$$
 is positive  $\alpha/2$ -stable.

Instead doing inference based on Z, do inference on the augmented space (X, S<sub>+</sub>), where (X | S<sub>+</sub>) is a GP with random covariance S<sub>+</sub>Σ.

### A conditional GP representation

Define: 
$$z_j^{(\ell)}(\mathbf{x}_k) = \frac{1}{M_\ell^{1/2}} \sum_{i=1}^{M_\ell} w_{ij}^{(\ell)} f_i^{(\ell)}(\mathbf{x}_k)$$
, with  $w_{ij}^{(\ell)} = (s_+^{(\ell)})^{1/2} \tilde{w}_{ij}^{(\ell)}$   
where  $s_+^{(\ell)} \sim S_{\alpha/2}^+$  and  $\tilde{w}_{ij}^{(\ell)}$  are (wlog.) zero mean, unit variance.

- Then, marginally  $w_{ij}^{(\ell)}$  have infinite variance and  $z_j^{(\ell)}$  is  $\alpha$ -stable as  $M_{\ell} \to \infty$ .
- The marginally stable  $z_j^{(\ell)}$  admits the representation:

$$\mathbf{z}_{j}^{(\ell)} \mid s_{+}^{(\ell)}, \Sigma^{(\ell)} \sim \mathcal{N}(0, s_{+}^{(\ell)} \Sigma^{(\ell)}),$$

where  $s_+^{(\ell)} \sim S_{lpha/2}^+$ .

### Covariance kernel of the conditional GP

Leads to a Cho and Saul (2009) type recursive expression for the conditional covariance kernels (Prop. 1, Loría and Bhadra, 2024+):

$$\begin{split} \boldsymbol{\Sigma}_{k,h}^{(\ell)} &= \pi^{-1} \left[ \left( 1 + \boldsymbol{s}_{+}^{(\ell-1)} \boldsymbol{\Sigma}_{k,k}^{(\ell-1)} \right) \left( 1 + \boldsymbol{s}_{+}^{(\ell-1)} \boldsymbol{\Sigma}_{h,h}^{(\ell-1)} \right) \right]^{\delta/2} J_{\delta}(\boldsymbol{\theta}_{k,h}^{(\ell)}), \\ \boldsymbol{\theta}_{k,h}^{(\ell)} &= \cos^{-1} \left\{ \left[ 1 + \boldsymbol{s}_{+}^{(\ell-1)} \boldsymbol{\Sigma}_{k,h}^{(\ell-1)} \right] \left[ 1 + \boldsymbol{s}_{+}^{(\ell-1)} \boldsymbol{\Sigma}_{k,k}^{(\ell-1)} \right]^{-1/2} \left[ 1 + \boldsymbol{s}_{+}^{(\ell-1)} \boldsymbol{\Sigma}_{h,h}^{(\ell-1)} \right]^{-1/2} \right\}, \end{split}$$

where all the  $s_{+}^{(\ell)}$  are independent  $S_{\alpha/2}^{+}$  random variables.

- Sanity check: The  $\alpha \to 2$  limit is Gaussian. In this case,  $S^+_{\alpha/2} \to 1$ w.p. 1, and one recovers the Cho and Saul result, with deterministic kernels  $\Sigma^{(\ell)}$ .
- But for α < 2, the covariance kernel s<sup>(ℓ)</sup><sub>+</sub>Σ<sup>(ℓ)</sup> is random, although it is positive definite w.p. 1.

### Implications of a random kernel on feature learning

- The deterministic kernel under a GP limit can be thought of a degenerate random kernel, that puts all its prior mass on one point.
- Since posterior  $\propto$  likelihood  $\times$  prior, the posterior is also a degenerate point mass, at the same point.
- A data-dependent learning of the kernel posterior is thus not possible in a GP limit (Aitchison et al., ICML, 2020, 2021).
- However, the posterior of the random kernel s<sup>(ℓ)</sup><sub>+</sub>Σ<sup>(ℓ)</sup> is non-degenerate, when α < 2. Data-dependent learning is possible.</p>

### Posterior inference and prediction

Suppose one observes  $(\mathbf{y}, \mathbf{x}) = \{y_k, \mathbf{x}_k\}_{k=1}^n$  from the model:

$$y_k = \psi(\mathbf{x}_k) + \varepsilon_k, \ \varepsilon_k \overset{i.i.d.}{\sim} \mathcal{N}(\mathbf{0}, \sigma^2),$$

and goal is to find the posterior predictive of  $(y^* \mid y, x, x^*)$ .

■ We have (Prop. 2, Loría and Bhadra, 2024+):

$$\mathbf{y}^* \mid \mathbf{y}, \mathbf{x}, \mathbf{x}^*, \{s^{(\ell)}_+\}_{\ell=2}^L \sim \mathcal{N}_m(\boldsymbol{\mu}^*, \boldsymbol{\Lambda}^*),$$

• A kriging-like result, except  $\mu^*, \Lambda^*$  depend on the stable  $s_+^{(\ell)}$ .

We propose from the prior of s<sup>(ℓ)</sup><sub>+</sub> ~ S<sup>+</sup><sub>α/2</sub> to implement an independent samples Metropolis, so the above likelihood is what needs to be evaluated in MCMC.

#### Results: conditional mutual information

- In GP, one would look at decay of correlation over distance.
- But covariance does not exist for stable processes. Need an alternative.
- Cover and Thomas define conditional mutual information (CMI) as:

$$I(Y_1; Y_2 \mid S) = \int_{S} D_{KL}[p(Y_1, Y_2 \mid s) \mid\mid p(Y_1 \mid s)p(Y_2 \mid s)]p(s)ds,$$

Which in our case becomes:

$$I(Y_1; Y_2 \mid S) = -(1/2) \int_{S} \log(1 - \rho_{Y_1, Y_2}^2(s)) p(s) ds$$

### Results: conditional mutual information



Figure: Decay of the conditional mutual information for the deep  $\alpha$ -kernel process as a function of the distance between the inputs with  $L = 2, \delta = 1$ . The limiting Gaussian case ( $\alpha = 2$ ) is also included.

### Results: function fit and UQ

The true function in 1-d is:

 $f(\xi) = 5 imes \mathbf{1}_{\{\xi > 0\}}$ 

and generate observations as  $y(\xi) = f(\xi) + \varepsilon$ ;  $\varepsilon \sim \mathcal{N}(0, 0.5^2)$ .

The true function in 2-d is:

$$f(\xi_1,\xi_2) = 5 \times \mathbf{1}_{\{\xi_1 > 0\}} + 5 \times \mathbf{1}_{\{\xi_1 > 0\}}$$

and generate  $y(\xi_1,\xi_2) = f(\xi_1,\xi_2) + \varepsilon$ ;  $\varepsilon \sim \mathcal{N}(0,0.5^2)$ ,

The true function in 10-d is:

$$f(\boldsymbol{\xi}) = 6 \operatorname{sign}(\xi_1) + 8 \operatorname{sign}(\xi_2 + \xi_3) + 6 \operatorname{sign}(\xi_4 + \xi_5) + 6 \operatorname{sign}(\xi_6 + \xi_7) \\ + 6 \operatorname{sign}(\xi_8 + \xi_9) + 6 \operatorname{sign}(\xi_{10})$$

and generate  $y(\boldsymbol{\xi}) = f(\boldsymbol{\xi}) + \varepsilon$ , with  $\varepsilon \sim \mathcal{N}(0, 0.5^2)$ .

# Results: visualizations of function fit and UQ in 1-d



(a) Function fit for the different methods.



(b) 90% posterior predictive intervals for the Bayesian methods.

Figure: Function fit and uncertainty quantification for the competing methods for a 1-d function with a single jump.

- GP tends to oversmooth the jump discontinuity. Stable captures jumps better.
- Agapiou and Castillo (2024, AoS) give theoretical support for this behavior.

#### Results: effect of $\alpha$ in 1-d



Figure: Comparison of predictions (solid lines) and 25th to 75th percentile posterior predictive intervals (shaded regions) in one dimension for different values of  $\alpha$  for the D $\alpha$ KP.

Table: Out-of-sample errors in numerical examples, in twenty different splits. Best in **bold**. Stable not available for more than 2 dimensions.

	One Dimension		Two Dimensions		Ten Dimensions	
Method	RMSE (SD)	MAE (SD)	RMSE (SD)	MAE (SD)	RMSE (SD)	MAE (SD)
$D\alpha$ -KP DIWP GP Bayes GP MLE NNGP Stable	0.57 (0.05) 1.08 (0.04) 0.69 (0.05) 0.77 (0.06) 1.08 (0.04) 0.52 (0.03)	0.45 (0.04) 0.84 (0.03) 0.52 (0.04) 0.57 (0.04) 0.84 (0.03) <b>0.42</b> (0.03)	0.86 (0.09) 1.69 (0.05) 0.90 (0.06) 1.19 (0.08) 1.69 (0.05) 0.57 (0.08)	0.67 (0.07) 1.36 (0.05) 0.70 (0.06) 0.92 (0.07) 1.36 (0.04) <b>0.45</b> (0.04)	<b>8.08</b> (0.38) 8.92 (0.38) 10.39 (0.63) 8.32 (0.38) 8.92 (0.39) -	<b>6.48</b> (0.33) 7.21 (0.32) 8.32 (0.52) 6.68 (0.34) 7.21 (0.34) -

- The paper contains additional results on predictive performance in some benchmark UCI data sets.
- Evidence of non-Gaussian feature learning, timing, MCMC mixing, coverage of the posterior credible intervals are all available.

# Concluding remarks

- The conditional GP representation makes inference and prediction almost as easy as GPs.
- However, there are important distinctions with a GP regime in terms of representation learning, and function fit.
- Another GP regime is the neural tangent kernel or NTK (Jacot et al., 2018, NeurIPS), which arises due to Gaussian SGD noise.
- Non-Gaussian SGD noise (Simsekli et al., 2019, ICML) should give rise to analogous non-GP stable regime for the NTK.

## Main references

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