

# Generalized Linear Models (GLMs)

Bruce A Craig

Department of Statistics  
Purdue University

# Outline

- Exponential family distributions
  - Examples of distributions
  - Expected value and variance
- Generalized linear models
  - Link functions
  - Estimation

Material covered in Chapter 8 of Faraway textbook

# LMs versus GLMs

- Linear model
  - Conceptual framework “DATA = FIT + ERROR”
  - FIT: Means of  $\mathbf{Y}$  are a linear function of  $\mathbf{X}$
  - ERROR: Deviations  $\varepsilon \sim N(0, \sigma^2 \mathbf{I})$
  - Overall model can be expressed  $\mathbf{Y} \sim N(\mathbf{X}\beta, \sigma^2 \mathbf{I})$
  - Normality most critical with prediction intervals
- Generalized linear model
  - The  $\mathbf{Y}$  are from an exponential family distribution
  - The means of  $\mathbf{Y}$  are linked to a linear function of  $\mathbf{X}$
  - Variance of each  $Y$  often a function of its mean
  - Best to move away from “DATA = FIT + ERROR” conceptual framework

# Exponential Family Distribution (EFD)

- Distribution can be discrete or continuous
- The probability mass/density function has general form

$$f(y_i; \theta_i, \phi) = \exp \left[ \frac{y_i \theta_i - b(\theta_i)}{a(\phi)} + c(y_i, \phi) \right]$$

- $\theta_i$  is the *canonical parameter* representing location (also called the *natural parameter*)
- $\phi$  is the *dispersion parameter* representing scale
- $a(\cdot)$ ,  $b(\cdot)$ ,  $c(\cdot)$  known functions
- Usually,  $a(\phi) = \phi/w$ 
  - $w$  is a known weight, can vary between observations
  - $\phi$  can be known (one-parameter distribution) or unknown (two-parameter distribution)

# Some Exponential Family Dists

- Discrete
  - Binomial (with known  $N$ )
  - Poisson
  - Negative binomial (with known  $r$ )
  - Multinomial
- Continuous
  - Exponential
  - Normal (and multivariate Normal)
  - Gamma
  - Beta
  - Weibull (with known shape)
  - Inverse gaussian

# Poisson Distribution $Y_i \sim Poi(\lambda)$

- The PMF is

$$\begin{aligned} f(y_i) &= \frac{\exp\{-\lambda\} \lambda^{y_i}}{y_i!} \\ &= \exp\{y_i \log(\lambda) - \lambda - \log(y_i!)\} \end{aligned}$$

- $\theta = \log(\lambda)$ ,  $\phi = 1$ , and  $w = 1$ .
- $a(\phi) = \phi/w$ ,  $b(\theta) = \exp\{\theta\}$ , and  $c(y_i, \phi) = -\log(y_i!)$

# Binomial Distribution $Y_i \sim \text{bin}(N, p)$

- The PMF is

$$\begin{aligned} f(y_i) &= \binom{N}{y_i} p^{y_i} (1-p)^{N-y_i} \\ &= \exp \left\{ y_i \log \frac{p}{1-p} + N \log(1-p) + \log \binom{N}{y_i} \right\} \end{aligned}$$

- $\theta = \log \frac{p}{1-p}$ ,  $\phi = 1$ , and  $w = 1$
- $a(\phi) = \phi/w$ ,  $b(\theta) = N \log(1 + \exp(\theta))$ ,  
 $c(y_i, \phi) = \log \binom{N}{y_i}$

# Binomial Distribution $Y_i \sim \text{bin}(N, p)$

- The PMF for  $\hat{p}_i = y_i/N$  is

$$\begin{aligned} f(\hat{p}_i) &= \binom{N}{N\hat{p}_i} p^{N\hat{p}_i} (1-p)^{N(1-\hat{p}_i)} \\ &= \exp \left\{ \frac{\hat{p}_i \log \frac{p}{1-p} + \log(1-p)}{1/N} + \log \binom{N}{N\hat{p}_i} \right\} \end{aligned}$$

- $\theta = \log \frac{p}{1-p}$ ,  $\phi = 1$ , and  $w = N$
- $a(\phi) = \phi/w$ ,  $b(\theta) = \log[1 + \exp(\theta)]$ ,  $c(\hat{p}_i, \phi) = \log \binom{N}{N\hat{p}_i}$

# Normal Distribution $Y_i \sim N(\mu, \sigma)$

- The PDF is

$$\begin{aligned} f(y_i) &= \frac{1}{\sigma} \sqrt{\frac{1}{2\pi}} \exp \left\{ -\frac{(y_i - \mu)^2}{2\sigma^2} \right\} \\ &= \exp \left\{ -\frac{y_i^2 - 2y_i\mu + \mu^2}{2\sigma^2} - \frac{1}{2} \log(2\pi\sigma^2) \right\} \\ &= \exp \left\{ \frac{y_i\mu - \mu^2/2}{\sigma^2} - \frac{1}{2} \left( \frac{y_i^2}{\sigma^2} + \log(2\pi\sigma^2) \right) \right\} \end{aligned}$$

- $\theta = \mu$ ,  $\phi = \sigma^2$ , and  $w = 1$
- $a(\phi) = \phi/w$ ,  $b(\theta) = \theta^2/2$ ,  $c(y_i, \phi) = -(y_i^2/\phi + \log(2\pi\phi)) / 2$

# Gamma Distribution $Y_i \sim \Gamma(\alpha, \beta)$

- The PDF is

$$\begin{aligned}f(y_i) &= \frac{\beta^\alpha}{\Gamma(\alpha)} y_i^{\alpha-1} \exp\{-\beta y_i\} \\&= \exp\{(\alpha - 1) \log y_i - \beta y_i + \alpha \log \beta - \log \Gamma(\alpha)\} \\&= \exp\left\{\frac{y_i \frac{\beta}{\alpha} - \log(\beta)}{-1/\alpha} + (\alpha - 1) \log y_i - \log \Gamma(\alpha)\right\}\end{aligned}$$

- $\theta = \beta/\alpha$ ,  $\phi = 1/\alpha$ , and  $w = 1$
- $a(\phi) = -\phi/w$ ,  $b(\theta_i) = \log(\theta)$ ,  
 $c(y_i, \phi) = (1/\phi - 1) \log y_i - \log \Gamma(1/\phi) - \log \phi/\phi$

# Expected Value of EFD

- Can use likelihood theory to determine mean and variance
- Given  $Y \sim EFD(\theta, \phi)$ , the log likelihood is

$$l(\theta, \phi; y) = \frac{y\theta - b(\theta)}{a(\phi)} + c(y, \phi)$$

- We know  $E(l'(\theta)) = 0$  at true value of  $\theta$ , so

$$\begin{aligned} l'(\theta, \phi; y) &= \frac{y - b'(\theta)}{a(\phi)} \\ &\downarrow \\ 0 &= \frac{E(y) - b'(\theta)}{a(\phi)} \end{aligned}$$

- Thus,  $E(Y) = b'(\theta)$

# Variance of EFD

- Similarly, we know

$$\begin{aligned} E(I''(\theta)) &= -E[(I'(\theta))^2] \\ &= -E\left(\frac{(y - b'(\theta))^2}{a^2(\phi)}\right) \\ &= -\text{Var}(Y)/a^2(\phi) \end{aligned}$$

- On the left-hand side

$$\begin{aligned} I''(\theta, \phi; y) &= \frac{-b''(\theta)}{a(\phi)} \\ &\downarrow \\ E(I''(\theta)) &= \frac{-b''(\theta)}{a(\phi)} \end{aligned}$$

- Thus,  $\text{Var}(Y) = b''(\theta)a(\phi)$

# Back to Examples

- Poisson

$$\begin{aligned}E(Y) &= b'(\theta) = \exp\{\log \lambda\} = \lambda \\ \text{Var}(Y) &= b''(\theta)a(\phi) = \exp\{\log \lambda\} \times 1 = \lambda\end{aligned}$$

- Binomial

$$\begin{aligned}E(Y) &= b'(\theta) = N \frac{\exp\{\theta\}}{1 + \exp\{\theta\}} = Np \\ \text{Var}(Y) &= b''(\theta)a(\phi) = N \frac{\exp\{\theta\}}{(1 + \exp\{\theta\})^2} \times 1 = Np(1 - p)\end{aligned}$$

- Binomial

$$\begin{aligned}E(\hat{p}) &= b'(\theta) = \frac{\exp\{\theta\}}{1 + \exp\{\theta\}} = p \\ \text{Var}(\hat{p}) &= b''(\theta)a(\phi) = \frac{\exp\{\theta\}}{(1 + \exp\{\theta\})^2} \times \frac{1}{N} = p(1 - p)/N\end{aligned}$$

# Back to Examples

- Normal

$$\begin{aligned}E(Y) &= b'(\theta) = \theta = \mu \\ \text{Var}(Y) &= b''(\theta)a(\phi) = 1 \times \sigma^2 = \sigma^2\end{aligned}$$

- Gamma

$$\begin{aligned}E(Y) &= b'(\theta) = \frac{1}{\theta} = \alpha/\beta \\ \text{Var}(Y) &= b''(\theta)a(\phi) = -\theta^{-2} \times \frac{-1}{\alpha} = \alpha/\beta^2\end{aligned}$$

# Generalized Linear Models

- Used to describe relationship between observations from an **EFD** and a set of predictors  $\mathbf{x}$
- Includes linear models so this is a broader more general model framework
- In addition to the specific distribution, need to specify a **link function** that describes how the mean of the response is related to a linear combination of predictors
- Do not confuse the link function with a transformation of the response variable
  - Analysis performed on the data, or  $Y$ , scale
  - Parameters  $\beta$ , however, are on the link scale

# Generalized Linear Models

- Data:  $(y_i, \mathbf{x}_i) = (y_i, x_{i1}, x_{i2}, \dots, x_{i,p-1})$ ,  $i = 1, 2, \dots, n$
- Random component:  $y_i \mid \mathbf{x}_i \stackrel{ind}{\sim} EFD(\theta_i, \phi)$
- Systematic component: Joint effects of  $\mathbf{x}_i$  determined through a linear combination

$$\eta_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_{p-1} x_{i,p-1} = \mathbf{x}_i \beta$$

- Link function  $g(\mu_i)$  relates  $\mu_i = E\{y_i\}$  with  $\eta_i = \mathbf{x}_i \beta$

$$g(\mu_i) = \eta_i$$

# Link Functions

- The link function  $g(\mu_i)$  completes a probability model
  - Logistic regression:  $g(\mu_i) = \text{logit}(\mu_i) = \mathbf{x}_i\beta$
  - Probit regression:  $g(\mu_i) = \Phi^{-1}(\mu_i) = \mathbf{x}_i\beta$
  - Poisson regression:  $g(\mu_i) = \log \mu_i = \mathbf{x}_i\beta$
  - Normal regression:  $g(\mu_i) = \mu_i = \mathbf{x}_i\beta$
- As it defines the mean function for the EFD
  - Logistic regression:  $\mu_i = \frac{1}{1 + \exp\{-\mathbf{x}_i\beta\}} = \frac{\exp\{\mathbf{x}_i\beta\}}{1 + \exp\{\mathbf{x}_i\beta\}}$
  - Probit regression:  $\mu_i = \Phi(\mathbf{x}_i\beta)$
  - Poisson regression:  $\mu_i = \exp\{\mathbf{x}_i\beta\}$
  - Normal regression:  $\mu_i = \mathbf{x}_i\beta$

# Canonical Link Functions

- Any monotone, continuous, and differentiable function can serve as the link
- When  $\eta = g(\mu) = \theta$ , then the canonical link is used
- This means that  $g(b'(\theta)) = \theta$  and  $\mathbf{X}'\mathbf{Y}$  is sufficient for  $\beta$ 

|                  |                |
|------------------|----------------|
| Normal: identity | Poisson: log   |
| Binomial: logit  | Gamma: inverse |
- While computationally convenient, context can compel using something other than the canonical link

# Fitting a GLM

- Estimation based on **maximum likelihood**
- Log-likelihood:

$$l(\beta) = \sum_{i=1}^n \frac{y_i \theta_i - b(\theta_i)}{a(\phi)} + \sum_{i=1}^n c(y_i, \phi)$$

↓

$$\frac{\partial l}{\partial \beta_j} = \frac{1}{\phi} \sum_{i=1}^n w_i \left( y_i \frac{\partial \theta_i}{\partial \beta_j} - b'(\theta_i) \frac{\partial \theta_i}{\partial \beta_j} \right)$$

- Using chain rule and fact that  $\partial \mu_i / \partial \theta_i = b''(\theta_i)$

$$\frac{\partial l}{\partial \beta_j} = \frac{1}{\phi} \sum_{i=1}^n \frac{(y_i - b'(\theta_i))}{b''(\theta_i)/w_i} \frac{\partial \mu_i}{\partial \beta_j}$$

# Fitting a GLM, continued

- Using mean and variance relationships, we set these partial derivatives to 0 to get the score equations

$$0 = \sum_{i=1}^n \frac{(y_i - \mu_i)}{V(\mu_i)} \frac{\partial \mu_i}{\partial \beta_j} = \sum_{i=1}^n \frac{(y_i - \mu_i)}{V(\mu_i)} x_{ij} / g'(\mu_i)$$

- Use iterative reweighted least-squares to obtain estimates

- Set  $k = 1$  and initial guess for  $\beta$  ( $\beta \rightarrow \hat{\eta} \xrightarrow{g^{-1}} \hat{\mu}$ )
- Approx  $g(y_i) = z_i$  using linearity around  $\hat{\mu}^{(k)}$   
$$z_i^{(k)} = g(\hat{\mu}_i^{(k)}) + (y_i - \hat{\mu}_i^{(k)})g'(\hat{\mu}_i^{(k)})$$
- Form weights using  $\text{Var}(z_i^{(k)}) = (g'(\hat{\mu}^{(k)}))^2 \text{V}(\hat{\mu}_i^{(k)})$
- Estimate  $\hat{\beta}^{(k+1)} = (\mathbf{X}'\mathbf{W}\mathbf{X})^{-1}\mathbf{X}'\mathbf{W}\mathbf{Z}$
- Repeat Steps 2-4 until convergence

# Maximum Likelihood

- Likelihood function
  - describes how probable data are for diff parameter values
- ML estimates
  - The parameter values that maximize the likelihood
  - Under regularity conditions,  $cov(\hat{\beta})$  is the inverse of the information matrix

$$\left[ -E \left( \frac{\partial^2 l(\beta)}{\partial \beta_i \partial \beta_j} \right) \right]^{-1}$$

- Properties
  - large-sample Normal distributions
  - asymptotically consistent (i.e. converge to the true parameter as sample size increases)
  - asymptotically efficient (i.e. large-sample SE no greater than other estimation methods)

# Fitting a GLM in R

- Utilize the `glm` and related functions

```
glm(formula, family = gaussian, data,...)
```

- Use `family` option to define EFD and link

```
family = binomial(link = "logit")  
family = gaussian(link = "identity")  
family = gamma(link = "inverse")  
family = inverse.gaussian(link = "1/mu^2")  
family = poisson(link = "log")
```