## RANDOMLY STARTED SIGNALS WITH WHITE NOISE

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## Abstract

It is shown that if B(t),  $t \ge 0$ , is a Wiener process, U is an independent random variable uniformly distributed on (0,1), and  $\varepsilon$  is a constant, then the distribution of B(t) +  $\varepsilon \sqrt{(t-U)^+}$ ,  $0 \le t \le 1$ , is absolutely continuous with respect to Wiener measure on C[0,1] if  $0 < \varepsilon < 2$ , and singular with respect to this measure if  $\varepsilon > \sqrt{8}$ .

1. INTRODUCTION. Let  $C[0,\infty)$  be the space of continuous functions on  $[0,\infty)$ , let  $\mathfrak F$  be the Borel subsets of  $C[0,\infty)$  for the topology of uniform convergence on compact sets, and let  $\mathfrak F$  be Wiener measure on  $\mathfrak F$ . For  $t\geq 0$  define the random variable B(t) on  $(C[0,\infty), \mathfrak F, \mathfrak F)$  by B(t)(f)=f(t), so that  $B(t), t\geq 0$ , is a standard Wiener process. Let U be a random variable independent of  $B(t), t\geq 0$ . (Formally, we must enlarge our probability space to permit such a U.) For a positive constant  $\delta$  define  $W_{\delta}(t), t\geq 0$ , by

$$W_{\delta}(t) = B(t) + \int_{0}^{t} \delta 2^{-1} (s-U)^{-\frac{1}{2}} I(U \le s \le U+1) ds,$$

where I denotes the indicator function, and let  $\gamma_\delta$  be the distribution of  $\textbf{W}_\delta.$  We prove

THEOREM 1. If  $0 < \delta < 2$ ,  $\gamma_{\delta}$  is absolutely continuous with respect to  $\mu$ . If  $\delta > \sqrt{8}$ ,  $\gamma_{\delta}$  is singular with respect to  $\mu$ .

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We do not know what happens for  $\delta \in [2, \sqrt{8}]$ . We remark that Theorem 1 is essentially equivalent to the statement that the distribution of  $B(t) + \delta \sqrt{(t-U)^+}$ ,  $0 \le t \le 1$ , is absolutely continuous with respect to Wiener measure on C[0,1] if  $0 < \delta < 2$ , and singular with respect to this measure if  $\delta > \sqrt{8}$ . Also, notice that it is easy to show that, for a fixed number a and any constant  $\varepsilon > 0$ , the distribution  $\eta$  of the process

$$Y(t) = B(t) + \int_{0}^{t} \epsilon 2^{-1} (s-a)^{-\frac{1}{2}} I(a \le s \le a+1) ds$$

is singular with respect to  $\mu$ . This can be done either using Girsanov's formula, which will be stated in Section 3, or by showing that if

$$F = \{f \in C[0,\infty): \lim_{n\to\infty} n^{-1} \sum_{k=0}^{n-1} (f(a+2^{-k})-f(a+2^{-(k+1)}))2^{k/2} > 0\},$$

then  $\mu(F) = 0$  while  $\eta(F) = 1$ , both statements holding by the strong law of large numbers for iid random variables.

The result ([1]) that, for constant  $\varepsilon$ , the probability

$$P_{\varepsilon} = P(\exists t: B(t+h)-B(t) > \varepsilon \sqrt{h} \text{ for all } h \in (0,1))$$

equals zero for  $\varepsilon > 1$  and equals one if  $\varepsilon < 1$  has somewhat the same flavor as Theorem 1, although the proofs of these results are only related in that both the proof that  $P_{\varepsilon} = 0$  if  $\varepsilon > 1$ , and the proof that  $\gamma_{\delta}$  is singular with respect to  $\mu$  if  $\delta > \sqrt{8}$ , have a common ancestor in Dvoretsky's argument in [2].

2. SINGULARITY. Let  $\epsilon > \sqrt{8}$ . The measure  $\gamma_{\epsilon}$  will be shown to be singular with respect to  $\mu$  by exhibiting a set  $A_{\epsilon} \in \mathcal{F}$  such that  $\gamma_{\epsilon}(A_{\epsilon}) = 1$  and  $\mu(A_{\epsilon}) = 0$ . Put  $\varphi(s) = [2(s-1)\ln s]^{\frac{1}{2}}/(s^{\frac{1}{2}}-1)$ . Then  $\varphi(s)$  decreases to  $\sqrt{8}$  as s decreases to 1. Let  $r(\epsilon) = r > 1$  satisfy  $\sqrt{8} < \varphi(r) < \epsilon$ , put  $\beta = \epsilon^2/\varphi^2(r) > 1$  and  $\alpha = (\beta+1)/2$ . For integers  $n \ge 1$  and  $0 \le k \le [r^n]$ , where [ ] is the greatest integer function, define the functions  $Q_{k,n}$  on  $C[0,\infty)$  by

$$Q_{k,n}(f) = n^{-\frac{1}{2}} \sum_{m=1}^{n} (r^{-m+1} - r^{-m})^{-\frac{1}{2}} (f(kr^{-n} + r^{-m+1}) - f(kr^{-n} + r^{-m})),$$

and put

$$S_n(f) = I(\max_{0 \le k \le [r^n]} Q_{k,n}(f) \ge (2n\alpha \ln r)^{\frac{1}{2}}).$$

The set  $\mathbf{A}_{\epsilon}$  is defined by

$$A_{\epsilon} = \{f: \lim \sup_{n\to\infty} S_n(f) = 1\}.$$

To show  $\mu(A_\epsilon)=0$ , we note that, considered as a random variable on  $(C[0,\infty),\,\mathfrak{F},\,\mu)$ ,  $Q_{k,n}$  is  $n^{-\frac{1}{2}}$  times the sum of n independent standard normal random variables, so that  $Q_{k,n}$  itself has a standard normal distribution. Thus if  $\Phi(x)=(2\pi)^{-\frac{1}{2}}\int\limits_{-\frac{1}{2}}^{x}e^{-t^2/2}dt$ ,

$$\mu(S_{n}(f) = 1) \leq ([r^{n}]+1)(1-\Phi[(2n\alpha \ln r)^{\frac{1}{2}}])$$

$$\leq 2r^{n}e^{-[(2n\alpha \ln r)^{\frac{1}{2}}]^{2}/2}$$

$$= 2r^{n(1-2\alpha)}.$$

Since 
$$\alpha > 1$$
,  $\sum_{n=1}^{\infty} \mu(S_n(f) = 1) < \infty$ , so  $\mu(A_{\epsilon}) = 0$ .

Now let k(U,n)=k be that integer satisfying  $kr^n \leq U < (k+1)r^n$ . The conditional distribution of

$$(r^{-m+1}-r^{-m})^{-\frac{1}{2}}[W_{\epsilon}(kr^{-n}+r^{-m+1}) - W_{\epsilon}(kr^{-n}+r^{-m})]$$

given U = u is normal with variance 1 and mean equal to

$$(r^{-m+1}-r^{-m})^{-\frac{1}{2}} \int_{0}^{kr^{-n}+r^{-m+1}} \varepsilon 2^{-1} (s-u)^{-\frac{1}{2}} ds$$

$$kr^{-n}+r^{-m}$$

$$\geq (r^{-m+1}-r^{-m})^{-\frac{1}{2}} \int_{0}^{kr^{-n}+r^{-m+1}} \varepsilon 2^{-1} (s-kr^{-n})^{-\frac{1}{2}} ds$$

$$= \varepsilon (r-1)^{-\frac{1}{2}} (r^{\frac{1}{2}}-1)$$

$$= (2\beta \ln r)^{\frac{1}{2}},$$

so that conditioned on U = u

$$Y = n^{-\frac{1}{2}} \sum_{m=1}^{n} (r^{-m+1} - r^{-m})^{-\frac{1}{2}} (W_{\epsilon}(kr^{-n} + r^{-m+1}) - W_{\epsilon}(kr^{-n} + r^{-m}))$$

is normal with variance 1 and mean exceeding  $(2n\beta \ln r)^{\frac{1}{2}}$ . In particular,  $P(Y > (2n\alpha \ln r)^{\frac{1}{2}} | U = u) \ge \Phi[(2n\beta \ln r)^{\frac{1}{2}} - (2n\alpha \ln r)^{\frac{1}{2}}] = q_n, \text{ so }$   $\gamma_{\epsilon} \{ f \in C[0,\infty) \colon S_n(f) = 1 \} \ge q_n. \text{ Since } q_n \to 1 \text{ as } n \to \infty \text{ we get } \gamma_{\epsilon}(A_{\epsilon}) = 1.$ 

3. ABSOLUTE CONTINUITY. If f(s),  $s \ge 0$ , is a measurable function such that  $\int\limits_0^\infty f^2(s) ds < \infty$ , Girsanov's formula (see [3]) gives that if  $\rho$  is the distribution of the process  $B(t) + \int\limits_0^t f(s) ds$ ,  $t \ge 0$ , then the Radon Nikodym derivative of  $\rho$  with respect to  $\mu$  is

$$\frac{d\rho}{d\mu} = \exp(\int_{0}^{\infty} f(s) dB(s) - \frac{1}{2} \int_{0}^{\infty} f^{2}(s) ds).$$

We let EX stand for  $\int_{C[0,\infty)} Xd\mu$ . Of course,  $E \frac{d\rho}{d\mu} = 1$ .

For an integer n>1 and a constant  $\delta>0$  put  $\alpha_n(v,t,\delta)=\alpha_n(v,t)=\delta 2^{-1}(v-t)^{-\frac{1}{2}}I(t+n^{-1}\leq v\leq t+1)$ . Let

$$W_{\delta}^{n}(t) = B(t) + \int_{0}^{t} \alpha_{n}(s,U)ds,$$

and let  $\gamma^n_\delta$  be the distribution of  $\textbf{W}^n_\delta.$  We will show that, for 0 <  $\delta$  < 2,

$$E(\frac{d\gamma_{\delta}^{n}}{du})^{2} \leq M_{\delta} < \infty,$$

which gives that the random variables  $\frac{d\gamma_\delta^n}{d\mu}$  are uniformly absolutely continuous with respect to  $\mu$ . Since  $W_\delta^n(t) - W_\delta(t) \le \delta/\sqrt{n} \to 0$  as  $n \to \infty$ , this implies that  $\gamma_\delta$  is absolutely continuous with respect to  $\mu$  if  $0 < \delta < 2$ .

We have

$$\begin{split} E(\frac{d\gamma_{\delta}^{n}}{d\mu})^{2} &= E[(\int_{0}^{1} exp(\int_{0}^{\infty} \alpha_{n}(v,t)dB(v) - \frac{1}{2} \int_{0}^{\infty} \alpha_{n}^{2}(v,t)dv)dt)^{2}] \\ &= E \int_{0}^{1} \int_{0}^{1} exp(\int_{0}^{\infty} (\alpha_{n}(v,t) + \alpha_{n}(v,s))dB(v) - \frac{1}{2} \int_{0}^{\infty} (\alpha_{n}^{2}(v,t) + \alpha_{n}^{2}(v,s))dv)ds dt \\ &= \int_{0}^{1} \int_{0}^{1} E \exp(\int_{0}^{\infty} (\alpha_{n}(v,t) + \alpha_{n}(v,s))dB(v) - \frac{1}{2} \int_{0}^{\infty} (\alpha_{n}^{2}(v,t) + \alpha_{n}^{2}(v,s))dv)ds dt \\ &= \int_{0}^{1} \int_{0}^{1} exp(\int_{0}^{\infty} \alpha_{n}(v,t)\alpha_{n}(v,s)dvE \exp(\int_{0}^{\infty} (\alpha_{n}(v,t) + \alpha_{n}(v,s))dB(v) - \frac{1}{2} \int_{0}^{\infty} (\alpha_{n}(v,t) + \alpha_{n}(v,s))^{2}dv)ds dt \\ &= \int_{0}^{1} \int_{0}^{1} exp(\int_{0}^{\infty} \alpha_{n}(v,t)\alpha_{n}(v,s)dv)ds dt \\ &= 2\int_{0}^{1} \int_{0}^{1} exp((\epsilon^{2}/4) \int_{0}^{s+1} [(v-t)(v-s)]^{-\frac{1}{2}}dv)ds dt. \end{split}$$

Now if s < t < s+1,

$$\begin{array}{c} s+1 \\ \int \\ t+n^{-1} \end{array} \left[ (v-t)(v-s) \right]^{-\frac{1}{2}} dv < \int \limits_{t}^{s+1} \left[ (v-t)(v-s) \right]^{-\frac{1}{2}} dv \\ = \ell n \left[ (2-(t-s) + 2\sqrt{1-(t-s)})/(t-s) \right] \\ \leq \ell n \left[ 4/(t-s) \right], \end{array}$$

so that 
$$E(\frac{d\gamma_{\delta}^n}{d\mu})^2 \leq 2\int\limits_0^1 \int\limits_s^1 (4/(t-s))^{\delta^2/4} dt \ ds < \infty \ \text{if } 0 < \delta < 2.$$

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