

On the Distribution of the Number  
of Successes in Independent Trials

by

Leon Jay Gleser  
Purdue University

Department of Statistics  
Division of Mathematical Sciences  
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## SUMMARY

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Let  $S$  be the number of successes in  $n$  independent Bernoulli trials, where  $p_j$  is the probability of success on the  $j$ -th trial. Let  $\underline{p} = (p_1, p_2, \dots, p_n)$ , and for any integer  $c$ ,  $0 \leq c \leq n$ , let  $H(c|\underline{p}) = P\{S \leq c\}$ . Let  $\underline{p}^{(1)}$  be one possible choice of  $\underline{p}$  for which  $E(S) = \lambda$ . For any  $n \times n$  doubly stochastic matrix  $\Pi$ , let  $\underline{p}^{(2)} = \underline{p}^{(1)}\Pi$ . Then in the present paper it is shown that  $H(c|\underline{p}^{(1)}) \leq H(c|\underline{p}^{(2)})$  for  $0 \leq c \leq [\lambda - 2]$ , and  $H(c|\underline{p}^{(1)}) \geq H(c|\underline{p}^{(2)})$  for  $[\lambda + 2] \leq c \leq n$ . These results provide a refinement of inequalities for  $H(c|\underline{p})$  obtained by Hoeffding [2]. Their derivation is achieved by applying consequences of the partial ordering of majorization.

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1. Introduction and Summary.

Let  $S$  be the number of successes in  $n$  independent Bernoulli trials, where  $p_j$  is the probability of success on the  $j$ -th trial,  $0 \leq p_j \leq 1$ . Let

$$(1.1) \quad \underline{p} = (p_1, p_2, \dots, p_n),$$

and for any integer  $c$ ,  $0 \leq c \leq n$ , let

$$(1.2) \quad H(c|\underline{p}) = P\{S \leq c\}.$$

For fixed  $c$ , we are interested in the relationship between  $H(c|\underline{p}^{(1)})$  and  $H(c|\underline{p}^{(2)})$ , where  $\underline{p}^{(1)}$  and  $\underline{p}^{(2)}$  each belong to the region

$$(1.3) \quad D_\lambda = \{\underline{p}: 0 \leq p_i \leq 1, i = 1, 2, \dots, n; \sum_{i=1}^n p_i = \lambda\}.$$

That is,  $\underline{p}^{(1)}$  and  $\underline{p}^{(2)}$  are sequences of probabilities for the independent Bernoulli trials each of which result in an expected number of successes,  $E(S)$ , equal to  $\lambda$ .

Hoeffding [1; Theorem 4] has shown that for all  $\underline{p} \in D_\lambda$ ,

$$(1.4) \quad 0 \leq H(c|\underline{p}) \leq H(c|n^{-1}(\lambda, \lambda, \dots, \lambda)), \text{ if } 0 \leq c \leq [\lambda - 2],$$

$$(1.5) \quad H(c|n^{-1}(\lambda, \lambda, \dots, \lambda)) \leq H(c|\underline{p}) \leq 1, \text{ if } [\lambda + 2] \leq c \leq n,$$

where

$$(1.6) \quad H(c|n^{-1}(\lambda, \lambda, \dots, \lambda)) = \sum_{k=0}^c \binom{n}{k} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k},$$

and  $[x]$  denotes the greatest integer  $\leq x$ . Hoeffding [1; Theorem 4] also obtained bounds on  $H(c|\tilde{p})$  for  $c = [\lambda - 1]$ ,  $[\lambda]$ , and  $[\lambda + 1]$ . These will be discussed at the end of Section 3.

To motivate the major result of the present paper, let

$$(1.7) \quad \tilde{p}^*(\lambda) = n^{-1}(\lambda, \lambda, \dots, \lambda)$$

and

$$(1.8) \quad \hat{p}(\lambda) = (1, 1, \dots, 1, \lambda - [\lambda], 0, 0, \dots, 0),$$

where in  $\hat{p}(\lambda)$  there are  $[\lambda]$  ones and  $n - [\lambda] + 1$  zeroes. Note that both  $\tilde{p}^*(\lambda)$  and  $\hat{p}(\lambda)$  are elements of  $D_\lambda$ . We have already noted the role  $\tilde{p}^*(\lambda)$  plays in Hoeffding's bounds (1.4) and (1.5), giving the upper bound to  $H(c|\tilde{p})$  for  $0 \leq c \leq [\lambda - 2]$  in (1.4) and the lower bound to  $H(c|\tilde{p})$  for  $[\lambda + 2] \leq c \leq n$  in (1.5). On the other hand, we have  $H(c|\hat{p}(\lambda)) = 0$  for  $0 \leq c \leq [\lambda - 2]$  and  $H(c|\hat{p}(\lambda)) = 1$  for  $[\lambda + 2] \leq c \leq n$ ; these, of course, are the lower and upper bounds to  $H(c|\tilde{p})$  in (1.4) and (1.5) respectively.

Now note that for any  $\tilde{p} \in D_\lambda$ , we can write

$$(1.9) \quad \tilde{p}^*(\lambda) = \tilde{p} \Pi^*,$$

where  $\Pi^*$  is an  $n \times n$  doubly stochastic matrix, all of whose elements are  $n^{-1}$ .

Also, for any  $\tilde{p} \in D_\lambda$ , we can write

$$(1.10) \quad \tilde{p} = \hat{p}(\lambda) \Pi(\tilde{p}),$$

where  $\Pi(\tilde{p})$  is an  $n \times n$  doubly stochastic matrix whose first  $[\lambda + 1]$  rows are equal to  $\tilde{p}$ , and whose last  $n - [\lambda + 1]$  rows are equal to  $(n - [\lambda + 1])^{-1} (\mathbf{1} - \lambda \tilde{p})$ , where  $\mathbf{1} = (1, 1, 1, \dots, 1)$  is the  $1 \times n$  vector all of whose elements are ones. It is thus apparent that proof of the following theorem would yield Hoeffding's inequalities

(1.4) and (1.5) as corollaries, and would provide a more detailed picture of the behavior of  $H(c|p)$  as a function of  $p$ .

Theorem 1.1.

Let  $p^{(1)} \in D_\lambda$  and suppose that there exists a doubly stochastic  $n \times n$  matrix  $\Pi$  for which

$$(1.11) \quad p^{(2)} = p^{(1)} \Pi.$$

Then  $p^{(2)} \in D_\lambda$  and

$$(1.12) \quad H(c|p^{(1)}) \leq H(c|p^{(2)}), \quad \text{if } 0 \leq c \leq [\lambda - 2],$$

and

$$(1.13) \quad H(c|p^{(1)}) \geq H(c|p^{(2)}), \quad \text{if } [\lambda + 2] \leq c \leq n.$$

In Section 2, we apply Ostrowski's [1; Theorem 15] fundamental theorem on majorization to the problem of ordering, over various choices of  $p \in D_\lambda$ , the expected values  $Eg(S)$  of any function  $g(k)$  on  $0, 1, 2, \dots, n$ . The results obtained in Section 2 are then used in Section 3 to prove Theorem 1.1.

2. Majorization.

A  $1 \times n$  vector  $\tilde{x}$  is said to majorize a  $1 \times n$  vector  $\tilde{y}$  if  $x_{[1]} \geq y_{[1]}$ ,  $x_{[1]} + x_{[2]} \geq y_{[1]} + y_{[2]}, \dots, \sum_{i=1}^{n-1} x_{[i]} \geq \sum_{i=1}^{n-1} y_{[i]}$ , and  $\sum_{i=1}^n x_{[i]} = \sum_{i=1}^n y_{[i]}$ , where the  $x_{[i]}$ 's and  $y_{[i]}$ 's are the components of  $\tilde{x}$  and  $\tilde{y}$ , respectively, arranged in descending order ( $x_{[1]} \geq x_{[2]} \geq \dots \geq x_{[n]}$ , and similarly for the  $y_{[i]}$ 's). The relation of majorization to doubly stochastic matrices is given by the following result of Karamata [1; Theorem 14].

Lemma 2.1.

The vector  $\tilde{x}$  majorizes the vector  $\tilde{y}$  if and only if there exists an  $n \times n$  doubly stochastic matrix  $\Pi$  such that  $\tilde{y} = \tilde{x} \Pi$ .

The following result, originally due to Ostrowski (see [1, pp. 30-33]), relates majorization to the ordering of the values of functions  $F(\underline{z})$  over regions of  $n$ -dimensional Euclidean space.

Lemma 2.2.

Let  $F(\underline{z})$  be a function defined on  $n$ -dimensional vectors  $\underline{z} = (z_1, z_2, \dots, z_n)$ . For any  $i, j, i \neq j$ , and all  $\underline{z}$  in a region  $D$ , suppose that

$$(2.1) \quad (z_i - z_j) \left( \frac{\partial F}{\partial z_i} - \frac{\partial F}{\partial z_j} \right) \geq 0$$

whenever  $z_i \geq z_j$ . If  $\underline{x}, \underline{y} \in D$ , and if  $\underline{x}$  majorizes  $\underline{y}$ , then

$$(2.2) \quad F(\underline{x}) \geq F(\underline{y}).$$

A function satisfying (2.1) over a region  $D$  is said to satisfy a Schur condition on  $D$ .

Let  $g(k)$  be any function on  $0, 1, \dots, n$ , and let  $S$  be the number of successes in  $n$  independent Bernoulli trials, where  $p_j$  is the probability of success on the  $j$ -th trial. Let

$$(2.3) \quad h(\underline{p}) = E g(S),$$

for  $\underline{p} = (p_1, p_2, \dots, p_n)$ . Then  $h(\underline{p})$  is a function defined over the region

$$D = \{ \underline{z} : 0 \leq z_i \leq 1, i = 1, 2, \dots, n \}.$$

Lemma 2.3.

For any two components  $p_i$  and  $p_j, i < j$ , of  $\underline{p}$ ,

$$(2.4) \quad (p_i - p_j) \left( \frac{\partial h(\underline{p})}{\partial p_i} - \frac{\partial h(\underline{p})}{\partial p_j} \right) = - (p_i - p_j)^2 \sum_{k=0}^{n-2} f(k | \underline{p}^{ij}) \Delta g(k),$$

where for any function  $s(k)$  defined on the non-negative integers

$$(2.5) \quad \Delta s(k) \equiv s(k+2) - 2s(k+1) + s(k)$$

is the second difference of  $s(k)$ , where  $\tilde{p}^{ij}$  is the  $1 \times (n-2)$  vector formed by deteting the  $i$ -th and  $j$ -th components of  $\tilde{p}$ , and where

$$(2.6) \quad f(k|\tilde{p}^{ij}) = \text{probability of } k \text{ successes in the } n-2 \text{ trials other than trials } i \text{ and } j;$$

for  $k = 0, 1, \dots, n-2$ .

Proof.

We adopt the convention that  $f(k|\tilde{p}^{ij}) = 0$  for  $k < 0$  or  $k > n-2$ . Under this convention,

$$(2.7) \quad \begin{aligned} P\{S = k\} &= (1 - p_i)(1 - p_j) f(k|\tilde{p}^{ij}) + (p_i + p_j - 2p_i p_j) f(k-1|\tilde{p}^{ij}) \\ &+ p_i p_j f(k-2|\tilde{p}^{ij}). \end{aligned}$$

Using (2.7), we find that the left-hand-side of (2.4) is

$$(2.8) \quad (p_i - p_j) \left( \frac{\partial h(p)}{\partial p_i} - \frac{\partial h(p)}{\partial p_j} \right) = \sum_{k=0}^n g(k) \Delta f(k-2|\tilde{p}^{ij}).$$

It is now easily shown that the right-hand-sides of (2.4) and (2.8) are equal.

Q.E.D.

As a corollary of Lemma 2.3, we can prove a result earlier obtained by Karlin and Novikoff [4].

Corollary 2.1.

Suppose that  $g(k)$  is convex on  $0, 1, \dots, n-2$ , in the sense that  $\Delta g(k) \geq 0$ ,  $k = 0, 1, \dots, n-2$ . If  $p^{(1)} \in D_\lambda$  and if  $p^{(2)} = p^{(1)} \Pi$ , where  $\Pi$  is any  $n \times n$  doubly stochastic matrix, then  $p^{(2)} \in D_\lambda$  and

$$(2.9) \quad h(\tilde{p}^{(1)}) \leq h(\tilde{p}^{(2)}).$$

Proof.

Since  $\Delta g(k) \geq 0$ ,  $k = 0, 1, \dots, n - 2$ ,  $-h(\tilde{p})$  satisfies a Schur condition, as can be seen from (2.4). Hence, Lemmas 2.1 and 2.2 imply that  $-h(\tilde{p}^{(1)}) \geq -h(\tilde{p}^{(2)})$ , from which (2.9) immediately follows. Q.E.D.

Karlin and Novikoff [4] proved Corollary 2.1 in a somewhat different way. Their proof however, embodies the ideas underlying the usual proof of Lemma 2.2.

From Corollary 2.1 and the arguments in Section 1 relating any  $\tilde{p} \in D_\lambda$  by doubly stochastic matrices to  $\tilde{p}^*(\lambda)$  and  $\hat{p}(\lambda)$ , it follows that for any  $g(k)$  convex on  $0, 1, \dots, n - 2$ , and any  $\tilde{p} \in D_\lambda$ ,

$$(2.10) \quad \begin{aligned} & (1 - \delta)g([\lambda]) + \delta g([\lambda + 1]) \\ & \leq h(\tilde{p}) = E g(S) \\ & \leq \sum_{k=0}^n g(k) \binom{n}{k} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k}, \end{aligned}$$

where  $\delta = \lambda - [\lambda]$ .

The result (2.10) implies that  $E|S - b|^a$ , for any  $a > 0$  and any real number  $b$ , is highest over  $D_\lambda$  when  $S$  has a binomial distribution with parameters  $n$  and  $n^{-1} \lambda$  (i.e.,  $\tilde{p} = \tilde{p}^*(\lambda)$ ), and lowest when

$$S = \begin{cases} [\lambda + 1], & \text{with probability } \delta, \\ [\lambda], & \text{with probability } 1 - \delta \end{cases}$$

(i.e.,  $\tilde{p} = \hat{p}(\lambda)$ ). The upper bound in (2.10) was first obtained (using a different method) by Hoeffding [2]. The lower bound in (2.10) can also be obtained by the methods in Hoeffding's [2] paper.

### 3. Proof of Theorem 1.1.

For fixed integer  $c$ ,  $0 \leq c \leq n$ , let

$$(3.1) \quad g_c(k) = \begin{cases} 1, & \text{if } 0 \leq k \leq c \\ 0, & \text{if } c + 1 \leq k \leq n. \end{cases}$$



Then

$$(3.2) \quad H(c|\underline{p}) = E(g_c(S)).$$

Note that  $g_c(k)$  is not convex on  $0, 1, \dots, n-2$  when  $c \leq n-1$ , so that we cannot directly use Corollary 2.1 to prove Theorem 2.1. Instead, we make use of Lemmas 2.1 to 2.3.

First note that for  $c \leq n-1$

$$(3.3) \quad g_c(k) = \begin{cases} 1, & k = c, \\ -1, & k = c-1, \\ 0, & \text{otherwise.} \end{cases}$$

Thus, from Lemma 2.3,

$$(3.4) \quad (p_i - p_j) \left( \frac{\partial H(c|\underline{p})}{\partial p_i} - \frac{\partial H(c|\underline{p})}{\partial p_j} \right) = - (p_i - p_j)^2 (f(c|\underline{p}^{ij}) - f(c-1|\underline{p}^{ij})).$$

Now, Samuels [6] has shown (using a well-known inequality attributed to Newton) that if  $f(k)$  is the probability of  $k$  successes in  $m$  independent Bernoulli trials, and if  $\sum_{k=0}^m k f(k) = \tau$ , then  $f(k)$  is increasing in  $k$  for  $k \leq [\tau]$  and decreasing in  $k$  for  $k \geq [\tau + 1]$ . Hence, using the characterization of  $f(k|\underline{p}^{ij})$  given in (2.6), and noting that  $\sum_{k=0}^{n-2} kf(k|\underline{p}^{ij}) = \lambda - p_i - p_j$ , we have that (3.4) is non-negative for  $c \geq [\lambda - p_i - p_j + 2]$  and non-positive for  $c \leq [\lambda - p_i - p_j]$ . Since  $0 \leq p_i + p_j \leq 2$ , all  $i \neq j$ , this result means that for all  $\underline{p} \in D_\lambda$ , (3.4) is  $\leq 0$  for  $c \leq [\lambda - 2]$  and  $\geq 0$  for  $c \geq [\lambda + 2]$ . Thus, the bounds (1.12) and (1.13) in Theorem 1.1 follow by a direct application of Lemmas 2.1 and 2.2. Q.E.D.

Remark I.

Hoeffding [2; Theorem 4] also showed that for all  $\underline{p} \in D_\lambda$ ,

$$(3.5) \quad 0 \leq H([\lambda - 1]|\underline{p}) \leq H([\lambda - 1]|n^{-1}(\lambda, \lambda, \dots, \lambda)),$$

and

$$(3.6) \quad H([\lambda + 1]|n^{-1}(\lambda, \lambda, \dots, \lambda)) \leq H([\lambda + 1]|\underline{p}) \leq 1.$$

It might be thought that more detailed results for the cases  $c = [\lambda - 1]$ ,  $c = [\lambda + 1]$ , similar to the results in Theorem 1.1, can be obtained. That is, we might suspect that  $\tilde{p}^{(1)} \in D_\lambda$ ,  $\tilde{p}^{(2)} = \tilde{p}^{(1)}\Pi$  for doubly stochastic  $\Pi$ , implies that

$$(3.7) \quad H([\lambda - 1]|\tilde{p}^{(1)}) \leq H([\lambda - 1]|\tilde{p}^{(2)}),$$

$$(3.8) \quad H([\lambda + 1]|\tilde{p}^{(1)}) \geq H([\lambda + 1]|\tilde{p}^{(2)}).$$

The inequalities (3.7) and (3.8) do not, however, always hold. Inequality (3.7) holds if  $\tilde{p}^{(1)}$  is restricted to belong to the subset

$$D_\lambda^0 = \{\tilde{p}: \tilde{p} \in D_\lambda; [\lambda - 1] \leq [\lambda - p_i - p_j] \leq [\lambda + 1], \text{ all } i \neq j\}$$

of  $D_\lambda$ , as can be seen from the proof of Theorem 1.1. (Note: if  $\tilde{p}^{(1)} \in D_\lambda^0$  and  $\tilde{p}^{(2)} = \tilde{p}^{(1)}\Pi$ ,  $\Pi$  doubly stochastic, then  $\tilde{p}^{(2)} \in D_\lambda^0$ .) Similarly, Inequality (3.8) holds if  $\tilde{p}^{(1)}$  is restricted to the subset

$$D_\lambda^1 = \{\tilde{p}: \tilde{p} \in D_\lambda; [\lambda] \leq [\lambda - p_i - p_j] \leq [\lambda + 1], \text{ all } i \neq j\}$$

of  $D_\lambda$ .

That (3.7) does not hold in general can be seen by letting  $n = 4$ ,  $\tilde{p}^{(1)} = (1, 1/2, 1/4, 1/4)$ , and

$$(3.9) \quad \Pi = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}.$$

Here,  $\lambda = 2$ ,  $\tilde{p}^{(2)} = (3/4, 3/4, 1/4, 1/4)$ ,  $[\lambda - 1] = 1$ ,

and

$$H([\lambda - 1]|\tilde{p}^{(1)}) = \frac{9}{32} > \frac{69}{256} = H([\lambda - 1]|\tilde{p}^{(2)}).$$

That (3.8) does not hold in general can be seen by letting  $n = 4$ ,  $\tilde{p}^{(1)} = (1/4, 0, 3/4, 3/4)$ , and  $\Pi$  be as in (3.9). Here  $\lambda = 7/4$ ,  $\tilde{p}^{(2)} = (1/8, 1/8, 3/4, 3/4)$ ,  $[\lambda + 1] = 2$ , and

$$H([\lambda + 1]|\tilde{p}^{(1)}) = \frac{55}{64} < \frac{883}{1024} = H([\lambda + 1]|\tilde{p}^{(2)}).$$

Since Theorem 4 of [2] does not even show that  $H([\lambda]|\tilde{p})$  is bounded by the values of  $H([\lambda]|\tilde{p})$  for  $\tilde{p} = \tilde{p}^*(\lambda)$  and  $\tilde{p} = \hat{p}(\lambda)$ , it is unlikely that an ordering between  $H([\lambda]|\tilde{p}^{(1)})$  and  $H([\lambda]|\tilde{p}^{(2)})$ , for  $\tilde{p}^{(2)} = \tilde{p}^{(1)}\Pi$ , that always goes in the same direction for all  $\tilde{p}^{(1)} \in D_\lambda$ , all doubly stochastic  $\Pi$ , can be demonstrated. Indeed, it is easy to find examples in which  $H([\lambda]|\tilde{p}^{(1)}) < H([\lambda]|\tilde{p}^{(2)})$ , and examples in which  $H([\lambda]|\tilde{p}^{(1)}) > H([\lambda]|\tilde{p}^{(2)})$ .

Remark 2.

Hoeffding [2; Theorem 5] also showed that if  $0 \leq b \leq \lambda \leq c \leq n$ , then for all  $\tilde{p} \in D_\lambda$ ,

$$\begin{aligned} (3.10) \quad & H(c|\tilde{p}^*(\lambda)) - H(b-1|\tilde{p}^*(\lambda)) \\ & \leq P\{b \leq S \leq c\} = H(c|\tilde{p}) - H(b-1|\tilde{p}) \\ & \leq 1. \end{aligned}$$

Correspondingly, as a corollary to Theorem 1.1, we can establish the following result.

Theorem 3.1.

Suppose  $0 \leq b \leq [\lambda - 1]$  and  $[\lambda + 2] \leq c \leq n$ . Let  $\tilde{p}^{(1)} \in D_\lambda$  and let  $\tilde{p}^{(2)} = \tilde{p}^{(1)}\Pi$ , where  $\Pi$  is an  $n \times n$  doubly stochastic matrix. Then

$$(3.11) \quad H(c|\tilde{p}^{(2)}) - H(b-1|\tilde{p}^{(2)}) \leq H(c|\tilde{p}^{(1)}) - H(b-1|\tilde{p}^{(1)}).$$

Remark 3.

For the possible statistical applications of the results obtained in this paper, the reader is urged to read Section 5 of [2], and also the comments in [3] and [6].

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