

An Alternative Proof of a Theorem of
Takacs on the GI|M|1 Queue*

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Mimeograph Series No. 47

June 1965

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An Alternative Proof of a Theorem of Takács on the GI|M|1 Queue*

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Marcel F. Neuts

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Abstract

An analytic proof is given of the fact that the stationary distribution for the imbedded Markov chain in a GI|M|1 queue is geometric. A generating function for the stationary transition probabilities is obtained as the unique solution to an integro-differential equation, which may be solved by reduction to a Wiener-Hopf equation.

We assume that customers arrive at a counter at the instants τ_0, τ_1, \dots and that the interarrival times $\tau_{n+1} - \tau_n$, ($n=0, 1, \dots$) are independent, identically distributed, positive random variables with distribution function $F(x)$. The service times are independent, identically distributed, negative exponential random variables with parameter μ .

This note supplies an analytic proof to the following theorem, due to Takács [2,3] :

Theorem:

The queue lengths $\xi_n, n = 0, 1, \dots$ immediately before arrivals form an irreducible, aperiodic Markov chain. This Markov chain is positive recurrent if and only if $\mu \alpha > 1$. The stationary probability distribution $\{\pi_k, k \geq 0\}$ is geometric with parameter δ , where δ is the root of the equation:

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$$(1) \quad z = \varphi \left[\mu(1-z) \right],$$

with smallest absolute value. $\varphi(\cdot)$ is the Laplace - Stieltjes transform of $F(\cdot)$.

Remark

This theorem was proved in [2] by Takács, using the backward equations for the Markov chain, together with a renewal argument. In [3] a combinatorial proof, together with a renewal argument was given, also by Takács.

We use the forward equations and obtain a recursive system of integro-differential equations for the generating functions of the n-step transition probabilities. The same method may be applied to the study of an $M|G|1$ queue in series with a queue with negative exponential service times.

Proof:

The queue-lengths $\{\xi_n, n \geq 0\}$ clearly form an irreducible, aperiodic Markov chain with transition probability matrix P , given by:

$$(2) \quad \begin{array}{cccccc} \left| \right. & & & & & & \left. \right| \\ & 1-p_0 & p_0 & 0 & 0 & 0 & \text{---} \\ & & & & & & \\ & 1-p_0-p_1 & p_1 & p_0 & 0 & 0 & \text{---} \\ & & & & & & \\ & 1-p_0-p_1-p_2 & p_2 & p_1 & p_0 & 0 & \text{---} \\ & & & & & & \\ & \vdots & \vdots & \vdots & \vdots & \vdots & \\ & & & & & & \end{array}$$

in which $p_j = \int_0^{\infty} e^{-\mu y} \frac{(\mu y)^j}{j!} dF(y),$

We denote the n -step transition probabilities by $P_y^{(n)}$. They satisfy the following recurrence relations:

$$(3) \quad P_{io}^{(n+1)} = \sum_{v=0}^{\infty} P_{iv}^{(n)} \left[1 - \sum_{\rho=0}^v p_{\rho} \right], \quad n \geq 0$$

and

$$(4) \quad P_{ij}^{(n+1)} = \sum_{v=j-1}^{\infty} P_{iv}^{(n)} P_{v-j+1}, \quad \text{for } j > 0, n \geq 0.$$

We introduce the generating functions:

$$(5) \quad U^{(n)}(z) = \sum_{j=0}^{\infty} P_{ij}^{(n)} \frac{z^j}{j!}, \quad |z| < \infty$$

and obtain the system of integro-differential equations:

$$(6) \quad \frac{d}{dz} U^{(n+1)}(z) = \int_0^{\infty} e^{-\mu y} U^{(n)}(z + \mu y) dF(y),$$

The initial conditions are obtained from equation (3), via the following manipulations:

$$(7) \quad U^{(n+1)}(0) = P_{io}^{(n+1)} = 1 - \sum_{v=0}^{\infty} P_{iv}^{(n)} \sum_{\rho=0}^v p_{\rho} =$$

$$1 - \sum_{\rho=0}^{\infty} \sum_{v=\rho}^{\infty} P_{iv}^{(n)} p_{v-\rho} = 1 - \sum_{\rho=0}^{\infty} \int_0^{\infty} e^{-\mu y} D_{\mu y}^{\rho} U^{(n)}(\mu y) dF(y)$$

$$= 1 - \int_0^{\infty} e^{-\mu y} (1 - D_{\mu y})^{-1} U^{(n)}(\mu y) dF(y) =$$

$$= \int_0^{\infty} dF(y) \int_0^{\mu y} e^{-\zeta} U^{(n)}(\zeta) d\zeta,$$

It is known from the theory of Markov chains that $\lim_{n \rightarrow \infty} U^{(n)}(z)$

= $U(z)$ exists. We have:

$$(8) \quad U(z) = \sum_{\nu=0}^{\infty} \frac{\mu^{\nu}}{\nu!} z^{\nu}, \quad |z| < \infty$$

The function $U(z)$ must satisfy the equations:

$$(9) \quad \frac{d}{dz} U(z) = \int_0^{\infty} e^{-\mu y} U(z+\mu y) dF(y), \quad \text{for all } |z| < \infty$$

and

$$(10) \quad U(0) = \int_0^{\infty} dF(y) \int_0^{\mu y} e^{-\zeta} U(\zeta) d\zeta$$

Equation (9) may be simplified through the following substitutions:

$$(11) \quad \begin{aligned} \text{a. } & \mu y = u \\ \text{b. } & U(z) = e^z R(z) \\ \text{c. } & F\left(\frac{u}{\mu}\right) - I(u) = K(u) \end{aligned}$$

$I(u)$ is a distribution with a unit-jump at zero.

We obtain:

$$(12) \quad R(z) = \int_0^{\infty} R(z+u) dK(u), \quad \text{for all } |z| < \infty$$

Lemma 1

If (12) holds for all $x > 0$, for some entire function $R(x)$, then it holds for all $|z| < \infty$.

Proof:

This may be verified by series expansion, using the fact that the distribution $K(\cdot)$ has moments of all order.

Partial integration in the right-hand side of (12) yields:

$$(13) \quad R'(x) = \int_0^{\infty} R'(x+u) K(u) du, \quad x > 0,$$

Setting $x + u = y$ and $K(-u) = H(u)$, we obtain:

$$(14) \quad R'(x) = \int_0^{\infty} R'(y) H(x-y) dy, \quad x > 0$$

which is a Wiener-Hopf equation for $R'(x)$.

It follows from the standard Wiener - Hopf theory [1,4] that equation (14) has a unique entire solution, which may be seen to be either an exponential function e^{-ax} for some $0 < a < 1$ or a constant.

Knowing this, we examine under which conditions there is an exponential solution

$$R'(x) = e^{-ax} C,$$

to equation (14). Substituting in (12) we obtain that a must be a solution to the equation:

$$(15) \quad 1 - a = \varphi(a\mu)$$

or if we set $1 - a = \delta$, then

$$(16) \quad \delta = \varphi \left[\mu (1-\delta) \right]$$

which is Takács' equation. The usual argument yields that (16) has a solution in $(0,1)$ if and only if $\alpha \mu > 1$.

So we obtain:

$$(17) \quad U(x) = C e^{\delta x}$$

The initial condition does not determine the constant C , since we have already included in the calculation of the initial condition that the $\{ \pi_v \}$ must sum to one.

This same condition applied to (17) yields $C = 1 - \delta$, which concludes the proof of the theorem.

Acknowledgement

The author is deeply indebted to Professor Harry Pollard for his many suggestions on the proof given above.

7.

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DOCUMENT CONTROL DATA - R&D

(Security classification of title, body of abstract and indexing annotation must be entered when the overall report is classified)

1. ORIGINATING ACTIVITY (Corporate author)		2a. REPORT SECURITY CLASSIFICATION	
Purdue University		UNCLASSIFIED	
		2b. GROUP	
3. REPORT TITLE			
An Alternative Proof of a Theorem of Takacs on the GI M 1 Queue			
4. DESCRIPTIVE NOTES (Type of report and inclusive dates)			
Technical Report, July 1965			
5. AUTHOR(S) (Last name, first name, initial)			
Neuts, Marcel F.			
6. REPORT DATE	7a. TOTAL NO. OF PAGES	7b. NO. OF REFS	
July, 1965	7	4	
8a. CONTRACT OR GRANT NO.	9a. ORIGINATOR'S REPORT NUMBER(S)		
NONR-1100(26)	Mimeograph Series Number 47		
b. PROJECT NO.	9b. OTHER REPORT NO(S) (Any other numbers that may be assigned this report)		
c.			
d.			
10. AVAILABILITY/LIMITATION NOTICES			
Releasable without limitations on dissemination			
11. SUPPLEMENTARY NOTES		12. SPONSORING MILITARY ACTIVITY	
		Logistics and Mathematical Sciences Branch Office of Naval Research Washington, D.C. 20360	
13. ABSTRACT			
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14.	KEY WORDS	LINK A		LINK B		LINK C	
		ROLE	WT	ROLE	WT	ROLE	WT
	GI M l queue Wiener-Hopf equation Ergodic distribution queue length						

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AN ALTERNATIVE PROOF OF A THEOREM OF TAKÁCS ON THE GI/M/1 QUEUE*

Marcel F. Neuts

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(Received July 23, 1965)

An analytic proof is given of the fact that the stationary distribution for the imbedded Markov chain in a GI/M/1 queue is geometric. A generating function for the stationary transition probabilities is obtained as the unique solution to an integro-differential equation, which may be solved by reduction to a Wiener-Hopf equation.

WE ASSUME that customers arrive at a counter at the instants τ_0, τ_1, \dots and that the interarrival times $\tau_{n+1} - \tau_n$, ($n=0, 1, \dots$) are independent, identically distributed, positive random variables with distribution function $F(x)$. The service times are independent, identically distributed, negative exponential random variables with parameter μ .

This note supplies an analytic proof to the following theorem, due to TAKÁCS:^[2,3]

THEOREM. *The queue lengths ξ_n , $n=0, 1, \dots$ immediately before arrivals form an irreducible, aperiodic Markov chain. This Markov chain is positive recurrent if and only if $\mu\alpha > 1$. The stationary probability distribution $\{\pi_k, k \geq 0\}$ is geometric with parameter δ , where δ is the root of the equation:*

$$z = \varphi[\mu(1-z)], \quad (1)$$

with smallest absolute value. $\varphi(\cdot)$ is the Laplace-Stieltjes transform of $F(\cdot)$

REMARK. This theorem was proved in reference 2 by Takács, using the backward equations for the Markov chain, together with a renewal argument. In reference 3 a combinatorial proof, together with a renewal argument was given, also by Takács.

We use the forward equations and obtain a recursive system of integro-differential equations for the generating functions of the n -step transition probabilities. The same method may be applied to the study of an M/G/1 queue in series with a queue with negative exponential service times.

Proof. The queue-lengths $\{\xi_n, n \geq 0\}$ clearly form an irreducible, aperiodic Markov chain with transition probability matrix P , given by:

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$$\left\| \begin{array}{cccccc} 1-p_0 & p_0 & 0 & 0 & 0 & \dots \\ 1-p_0-p_1 & p_1 & p_0 & 0 & 0 & \dots \\ 1-p_0-p_1-p_2 & p_1 & p_1 & p_0 & 0 & \dots \\ \dots & \dots & \dots & \dots & \dots & \dots \end{array} \right\|, \quad (2)$$

in which $p_j = \int_0^\infty e^{-\mu y} \frac{(\mu y)^j}{j!} dF(y)$.

We denote the n -step transition probabilities by $P_{ij}^{(n)}$. They satisfy the following recurrence relations:

$$P_{i0}^{(n+1)} = \sum_{\nu=0}^{\infty} P_{i\nu}^{(n)} [1 - \sum_{\rho=0}^{\nu} p_\rho], \quad (n \geq 0) \quad (3)$$

and
$$P_{ij}^{(n+1)} = \sum_{\nu=j-1}^{\infty} P_{i\nu}^{(n)} p_{\nu-j+1}, \quad \text{for } j > 0, n \geq 0. \quad (4)$$

We introduce the generating functions:

$$U^{(n)}(z) = \sum_{j=0}^{\infty} P_{ij}^{(n)} z^j / j!, \quad (|z| < \infty), \quad (5)$$

and obtain the system of integro-differential equations:

$$\frac{d}{dz} U^{(n+1)}(z) = \int_0^\infty \bar{e}^{\mu y} U^{(n)}(z + \mu y) dF(y). \quad (6)$$

The initial conditions are obtained from equation (3), via the following manipulations:

$$\begin{aligned} U^{(n+1)}(0) &= P_{i0}^{(n+1)} = 1 - \sum_{\nu=0}^{\infty} P_{i\nu}^{(n)} \sum_{\rho=0}^{\nu} p_\rho \\ &= 1 - \sum_{\rho=0}^{\infty} \sum_{\nu=\rho}^{\infty} P_{i\nu}^{(n)} p_{\nu-\rho} = 1 - \sum_{\rho=0}^{\infty} \int_0^\infty \bar{e}^{\mu y} D_{\mu y}^\rho U^{(n)}(\mu y) dF(y) \\ &= 1 - \int_0^\infty \bar{e}^{\mu y} (1 - D_{\mu y})^{-1} U^{(n)}(\mu y) dF(y) \\ &= \int_0^\infty dF(y) \int_0^{\mu y} \bar{e}^\xi U^{(n)}(\xi) d\xi. \end{aligned} \quad (7)$$

It is known from the theory of Markov chains that $\lim_{n \rightarrow \infty} U^{(n)}(z) = U(z)$ exists. We have:

$$U(z) = \sum_{\nu=0}^{\infty} (\pi_\nu / \nu!) z^\nu. \quad (|z| < \infty) \quad (8)$$

The function $U(z)$ must satisfy the equations:

$$\frac{d}{dz} U(z) = \int_0^\infty \bar{e}^{\mu y} U(z + \mu y) dF(y), \quad \text{for all } |z| < \infty, \quad (9)$$

and
$$U(0) = \int_0^\infty dF(y) \int_0^{\mu y} \bar{e}^\xi U(\xi) d\xi. \quad (10)$$

Equation (9) may be simplified through the following substitutions:

$$\begin{aligned} (a) \quad & \mu y = u, \\ (b) \quad & U(z) = e^z R(z), \\ (c) \quad & F(u/\mu) - I(u) = -K(u). \end{aligned} \tag{11}$$

$I(u)$ is a distribution with a unit-jump at zero. We obtain:

$$R(z) = \int_0^\infty R(z+u) dK(u), \text{ for all } |z| < \infty. \tag{12}$$

LEMMA 1. If (12) holds for all $x > 0$, for some entire function $R(x)$, then it holds for all $|z| < \infty$.

Proof. This may be verified by series expansion, using the fact that the distribution $K(\cdot)$ has moments of all order.

Partial integration in the right-hand side of (12) yields:

$$R'(x) = \int_0^\infty R'(x+u) K(u) du. \quad (x > 0) \tag{13}$$

Setting $x+u=y$ and $K(-u)=H(u)$, we obtain:

$$R'(x) = \int_0^\infty R'(y) H(x-y) dy, \quad (x > 0) \tag{14}$$

which is a Wiener-Hopf equation for $R'(x)$.

It follows from the standard Wiener-Hopf theory^[1,4] that equation (14) has a unique entire solution, which may be seen to be either an exponential function \bar{e}^{ax} for some $0 < a < 1$ or a constant.

Knowing this, we examine under which conditions there is an exponential solution

$$R'(x) = e^{-ax} C,$$

to equation (14). Substituting in (12) we obtain that a must be a solution to the equation:

$$1 - a = \varphi(a\mu), \tag{15}$$

or if we set $1 - a = \delta$, then

$$\delta = \varphi[\mu(1 - \delta)], \tag{16}$$

which is Takács' equation. The usual argument yields that (16) has a solution in $(0, 1)$ if and only if $\alpha\mu > 1$. So we obtain:

$$U(x) = Ce^{\delta x}. \tag{17}$$

The initial condition does not determine the constant C , since we have already included in the calculation of the initial condition that the $\{\pi_n\}$ must sum to one.

This same condition applied to (17) yields $C=1-\delta$, which concludes the proof of the theorem.

ACKNOWLEDGMENT

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1. R. E. A. C. PAYLEY AND N. WIENER, *Fourier Transforms in the Complex Domain*, Am. Math. Soc. Col. Publ., Vol. XIX, 1934.
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