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On the Moments of Elementary Symmetric Functions of the Roots of Two Matrices and Approximations to Their Distributions

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On the Moments of Elementary Symmetric Functions of the Roots of Two Matrices and Approximations to Their Distributions

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- 1. Summary. The moments of the $(s-i)^{th}$ elementary symmetric function (esf) in s non-null characteristic roots, λ_i (i=1,2,...,s), of a matrix in multivariate analysis have been shown to be derivable from those of the 1th esf. The first four moments of the $(s-1)^{th}$ esf have thus been obtained from those of the first esf and an approximation to the distribution of the $(s-1)^{th}$ esf suggested. Further, the moments of the $(s-1)^{th}$ esf in the s characteristic mots, $\theta_1 = \lambda_1/(1+\lambda_1)$, have been derived from those of the first esf in the λ 's and an approximation to the distribution of the former also obtained. Similar results have been given for the second and $(s-2)^{th}$ esf in the λ 's but general expressions have been presented only for the first two moments of these esf's. In addition, the first moment has been obtained for the 1th esf in the λ 's and in the θ 's, and based on all these studies, approximations to the distribution of the 1th esf in each case suggested.
- 2. Introduction. Most of the distribution problems in multivariate analysis are based on the distribution of the non-null characteristic roots of a matrix derived from sample observations taken from multivariate normal populations. This distribution, given by Roy [8], is of the form

$$(2.1) \quad f(\lambda_1, \lambda_2, \dots, \lambda_s; m, n) = C(s, m, n) \prod_{i=1}^{s} \left\{ \lambda_i^{m_i} / (1 + \lambda_i)^{m+n+s+1} \right\} \left\{ \lambda_i^{(\lambda_i - \lambda_j)} \right\}$$

$$0 < \lambda_1 \le \lambda_2 \le \dots \le \lambda_s < \infty,$$



where

(2.2)
$$C(s,m,n) = \pi^{s/2} \frac{s}{\pi} \left\{ \left(\frac{2m+2n+s+1+2}{2} \right) / \left(\frac{2m+1+1}{2} \right) / \left(\frac{2n+1+1}{2} \right) \right\}$$

and m and n are defined differently for various situations described in (3), (5).

The studies on the first esf in the λ 's have been carried out by Pillai (2), (4), (5) and Fillai and Samson (7). In this paper, moments and approximations to the distributions of some of the other esf's are considered and in particular those of the (s-1)th, (s-2)th and the second esf's. Further, the esf's in the characteristic roots, $\Theta_1 = \lambda_1/(1+\lambda_1)$, are studied in a similar manner.

3. Moments of the $(s-1)^{th}$ esf in the λ 's. First consider the following lemma.

Lemma 1. If $U_{i,m,n}^{(s)}$ and $\mu_r \left\{ U_{i,m,n}^{(s)} \right\}$ denote the ith esf in the s λ 's following the distribution (2.1), and its rth moment respectively, then

(3.1)
$$\mu_{\mathbf{r}}^{i} \left\{ \mathbf{U}_{s-i,m,n}^{(s)} \right\} = \frac{\mathbf{c}(\underline{s},\underline{m},\underline{n})}{\mathbf{c}(\underline{s},\underline{n-r},\underline{m+r})} \quad \mu_{\mathbf{r}}^{i} \left\{ \mathbf{U}_{i,n-r}^{(s)} \right\}.$$

Froof. Let $\gamma_i = 1/\lambda_i$ (i=1,2,...,s). Then the ith esf in the γ 's equals the quotient of the (s-i)th esf and the sth esf in the λ 's or, reciprocally, the (s-i)th esf in the λ 's equals the quotient of the ith esf and the sth esf in the γ 's. Now note that the distribution of the γ 's follows the distribution (2.1) with m and n interchanged. Hence the lemma follows. From lemma 1, putting i=1 in (3.1), the rth moment of the (s-1)th esf in the λ 's is obtained as

(3.2)
$$\mu_{\mathbf{r}}^{'}\left\{U_{\mathbf{s}-\mathbf{l},\mathbf{m},\mathbf{n}}^{(\mathbf{s})}\right\} = \frac{\mathbf{c}(\mathbf{s},\mathbf{m},\mathbf{n})}{\mathbf{c}(\mathbf{s},\mathbf{n}-\mathbf{r},\mathbf{m}+\mathbf{r})} \quad \mu_{\mathbf{r}}^{'}\left\{U_{\mathbf{l},\mathbf{n}-\mathbf{r},\mathbf{m}+\mathbf{r}}^{(\mathbf{s})}\right\}.$$

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The first three central moments of $U_{1,m,n}^{(s)}$ are available in (2), (7) and the fourth central moment in (5). Using these results, the following first four moments of the (s-1)th esf in the λ 's are obtained:

(33)
$$\mu_{1}^{r} \left\{ v_{s-1,m,n}^{(s)} \right\} = \frac{s(2n+s-1)}{2(m+1)} \quad \frac{s}{n} \quad \frac{(2m+1+1)}{(2n+1-1)} ,$$

$$(3.4) \quad \mu_{2}^{2} \quad \left\{ u_{s-1,m,n}^{(s)} \right\} = \frac{s(2n+s-3)}{4(m+2)^{2}} \quad \left\{ \frac{(2m+2n+s+1)(2m+s+4)}{(m+1)(2m+3)} + s(2n+s-3) \right\} \land$$
where $\Lambda = \frac{s}{n} \quad \left(\frac{2m+1+3}{2m+1-3} \right) \left(\frac{2m+1+3}{2m+1-3} \right)$,

$$(3.5) \quad \mu_3^* \left\{ \begin{matrix} u(s) \\ s-1, m, n \end{matrix} \right\} = \left\{ \frac{s(2n+s-5)}{8(n+3)^3} \right\} \quad \left\{ \frac{(2m+2n+s+1)(2m+s+6)}{(m+2)(2m+7)} \right\}$$

$$\left(\frac{4(m+2n+s-2)(2m+2s+6)}{(m+1)(2m+8)} + 3s(2n+s-5)\right) + s^{2}(2n+s-5)^{2}$$
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where
$$B = \int_{1}^{S} \frac{(2m+i+5)(2m+i+3)(2m+i+1)}{(2n+i-1)(2n+i-3)(2n+i-5)}$$

and

$$(3.6) \quad \mu_{\downarrow} \left\{ v_{s-1,m,n}^{(s)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{48(2m+2n+s+1)(c+D(n-4)(m+n+s+1))}{(m+1)(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{48(2m+2n+s+1)(c+D(n-4)(m+n+s+1))}{(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{48(2m+2n+s+1)(c+D(n-4)(m+n+s+1))}{(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{48(2m+2n+s+1)(c+D(n-4)(m+n+s+1))}{(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{s(2n+s-7)}{(m+2)(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} \quad \left\{ \frac{s(2n+s-7)}{(m+2)(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+2)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+3)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+4)^4} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+3)(m+3)(m+5)(2m+7)(2m+9)(2m+11)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+3)(m+3)(m+5)(2m+7)(2m+7)(2m+7)(2m+7)(2m+7)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+3)(m+5)(2m+7)(2m+7)(2m+7)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+2)(m+7)(2m+7)(2m+7)(2m+7)(2m+7)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+2)(m+7)(2m+7)(2m+7)(2m+7)(2m+7)(2m+7)(2m+7)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+7)(2m+7)(2m+7)(2m+7)(2m+7)(2m+7)(2m+7)} \right\} = \left\{ \frac{s(2n+s-7)}{16(m+7)(2m+7)$$

where

$$c = \frac{6}{\frac{5}{100}} (-1)^{j} f_{j} p^{6-j},$$

$$D = \frac{5}{100} (-1)^{j} g_{j} p^{5-j}$$

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$$p = (m + s + 5)$$

$$E = s(2m+s-7) \left\{ \frac{2(2m+2n+s+1)(2m+s+8)}{(m+3)(2m+9)} \left(\frac{4(2n+m+s-3)(2m+2s+8)}{(m+2)(m+3)} + 3s(2m+s-7) \right) + s^2(2m+s-7)^2 \right\},$$

$$F = \frac{s}{4\pi} \frac{(2m+i+7)(2m+i+5)(2m+i+3)(2m+i+1)}{(2m+i+3)(2m+i+3)(2m+i+1)}$$

and f's and g's are polynomials in s given in [5].

4. Moments of the $(s-1)^{th}$ esf in the 9's. As in the previous section a lemma may be stated.

Lemma 2. If $v_{i,m,n}^{(s)}$ and $u_r^{'}$ $\left\{v_{i,m,n}^{(s)}\right\}$ denote the ith esf in the s θ 's and its r^{th} moment respectively, then

(4.1)
$$\mu_{\mathbf{r}}^{'} \left\{ \mathbf{v}_{s-1,m,n}^{(s)} \right\} = \frac{c(s,m,n)}{c(s,n,m+r)} \sum_{j=0}^{r} {r \choose j} s^{r-j} \mu_{j}^{'} \left\{ \mathbf{v}_{1,n,m+r}^{(s)} \right\}.$$

Proof. For proving (4.1), it is enough if we note that

(4.2)
$$v_{s-1,m,n}^{(s)} = (s + \sum_{i=1}^{s} \gamma_i) / \frac{s}{n} (1 + \gamma_i)$$
 where γ 's,

as stated above, follow the distribution (2.1) with m and n interchanged.

The rest easily follows.

The first four moments of the (s-1)th esf in the 0's thus obtained are given below:

(4.3)
$$\mu_{1}^{\circ} \left\{ v_{s-1,m,n}^{(s)} \right\} = \frac{s(2m+2n+s+3)}{2(m+1)} \quad \frac{s}{i=1} \quad \frac{(2m+i+1)}{(2m+2n+s+i+2)} ,$$

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$$(4.4) \mu_{2}^{\circ} \left\{ v_{s-1,m,n}^{(s)} \right\} = \frac{s(2m+2n+s+5)}{h(m+2)^{2}} \left\{ \frac{(2n+s+1)(2m+s+4)}{(m+1)(2m+5)} + s(2m+2n+s+5) \right\} . \Lambda_{1}$$

where
$$A_1 = \frac{s}{s} \frac{(2m+i+3)(2m+i+1)}{(2m+2n+s+i+4)(2m+2n+s+i+2)}$$
,

and

$$(4.6) \quad \mu_{4} \left\{ V_{s-1,m,n}^{(s)} \right\} = \left\{ \frac{s(2m+2n+s+9)}{16(m+4)^4} \right\} \quad \left\{ \frac{h8(2n+s+1)\left(C+Dn(m+n+s+5)\right)}{(m+1)(m+2)(m+3)(m+5)\left(2m+7\right)\left(2m+9\right)\left(2m+11\right)} + E_{1} \right\} \cdot F_{1}$$

where C and D are given in (3.6),

$$E_{1} = s(2m+2n+s+9) \left\{ \frac{2(2n+s+1)(2m+s+8)}{(m+3)(2m+9)} \left(\frac{4(2n+m+s+5)(2m+2s+8)}{(m+2)(m+5)} + 3s(2m+2n+s+9) \right) + s^{2}(2m+2n+s+9)^{2} \right\}$$

and
$$F_1 = \frac{s}{s} \frac{(2m+i+7)(2m+i+5)(2m+i+3)(2m+i+1)}{(2m+2n+s+i+6)(2m+2n+s+i+4)(2m+2n+s+i+4)(2m+2n+s+i+4)}$$

It may be noted that $\mu_r \left\{ v_{s-1,m,n}^{(s)} \right\}$ can be obtained from $\mu_r \left\{ v_{s-1,m,n}^{(s)} \right\}$ by attaching negative signs to all terms except that in n in each linear

 factor involving n and further changing n to m+n+s+1. This method was noted by Pillai earlier (2) in connection with the moments of the first esf.

5. Moments of other esf's in the Θ 's and λ 's.

It has been shown (2), (4), (6) that the distribution of 0's obtainable from (2.1) by the transformation $\lambda_1 = \theta_1/(1-\theta_1)(i=1,2,...,s)$ can be expressed in a determinantal form and when integrated within the limits, $0 < \theta_1 \le ... \le \theta_s < 1$, is given by

(5.1)
$$C(s,m,n) \ V(s-1,s-2,...,1,0) = C(s,m,n)$$

$$\int_{0}^{\infty} e_{s}^{m+s-1} (1-e_{s})^{n} de_{s} \int_{0}^{\infty} e_{s}^{m} (1-e_{s})^{n} de_{s}$$

Further,

(5.2)
$$\mu_1 \left\{ V_{1,m,n}^{(s)} \right\} = V(s,s-1,\ldots,s-1+1,s-1-1,\ldots,1,0).$$

The determinant in (5.2) can be shown to be equal to

(5.3)
$$\binom{s}{i} = \frac{1}{\pi} \frac{(2m+s-j+2)}{(2m+2n+2s-j+3)}$$

based on some particular cases evaluated in [6]. From this result, using the method given at the end of the last section or otherwise, the first moment of the i^{th} esf in the λ 's is given by

(5.4)
$$\mu_{1}^{i} \left\{ U_{i,m,n}^{(s)} \right\} = {i \choose i} \frac{i}{\pi} \frac{(2m+s-j+2)}{(2n+j-1)}.$$

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Now consider
$$\mu_2\left\{v_{2,m,n}^{(8)}\right\}$$
. It can be shown that [1]

$$(5.5) \mu_{2}^{\circ} \left\{ v_{2,m,n}^{(s)} \right\} = V(s+1,s,s-3,...,l,0) + V(s+1,s-l,s-2,s-4,...,l,0) + V(s,s-1,s-2,s-3,s-5,s-6,...,l,0)$$

Now substituting the values of the determinants (6) in (5.5),

(5.6)
$$\mu_{2}^{1}\left\{v_{2,m,n}^{(s)}\right\} = \frac{s(s-1)(2m+s)(2m+s+1)}{3!} \quad c_{1}$$

where
$$G_1 = 6n^2 \left\{ \frac{1}{4} s(s-1)m^2 + 2(s-1)(2s^2+s+8)m + s^4 + 7s^2 - 8s + 12 \right\}$$

 $+ 3n \left\{ 16s(s-1)m^3 + \frac{1}{4}(s-1)(8s^2+5s+8)m^2 + 2(s-1)(10s^3+12s^2+27s+24)m + 4s^5 + 3s^4 + 12s^3 + 5s^2 - 24s + 36 \right\} + s(s+1)(2m+s+1)(2m+s+2)(m+s)(2m+2s+1)$
 $+ (s-2)(2m+2s+3)(2m+s-1) \left\{ \frac{1}{4}sm^2 + 2s(3s+2)m + 2s^3 + 3s^2 + s + 6 \right\}$.

It may be pointed out that (5.6) equals (4.4) when s=3. The second moment of the second esf in the λ 's can be obtained from (5.6) using the method given at the end of section 4. Hence we get

(5.7)
$$\mu_{2}^{'}\left\{U_{2,m,n}^{(s)}\right\} = \frac{s(s-1)}{3!} \frac{(2m+s)(2m+s+1)}{n (2n+j-3)} \cdot c$$

where G is obtained from G_1 by attaching negative sign to the first degree term in n and then changing n to m+n+s+1 in all the terms.

Further, applying (3.1) with 1=2 and r=2, we get from (5.7)

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(5.8)
$$\mu_{2}^{\circ}\left\{U_{s-2,m,n}^{(s)}\right\} = \frac{C(s,m,n)}{C(s,n-2,m+2)} \quad \mu_{2}^{\circ}\left\{U_{2,n-2,m+2}^{(s)}\right\}.$$

6. Approximations to the distributions of $U_{s-1,m,n}^{(s)}$ and $V_{s-1,m,n}^{(s)}$. The following approximation to the distribution of $U_{s-1,m,n}^{(s)}$ may be suggested: (writing $U_{s-1}^{(s)}$ for $U_{s-1,m,n}^{(s)}$)

$$(6.1) \ r(v_{s-1}^{(s)}) = \frac{\left\{v_{s-1}^{(s)}\right\} \left\{(2m+3)/(s-1)\right\} - 1}{\left\{\beta(2m+3,2m+s-1)\right\} (s-1)_{s}(2m+3)/(s-1) \left\{1 + \left(\frac{v_{s-1}^{(s)}}{s}\right)^{1/(s-1)}\right\}} \frac{2m+2m+s+2}{0 < v_{s-1}^{(s)} < \infty}.$$

Similarly, an approximation to the distribution of $v_{s-1}^{(s)}$ may be suggested as follows:

(6.2)
$$r(v_{s-1}^{(s)}) = \frac{\left\{v_{s-1}^{(s)}\right\}^{\left\{(2m+3)/(s-1)\right\}^{-1}\left\{1-\left(\frac{s-1}{s-1}\right)^{1/(s-1)}\right\}^{-2m+s}}{\left\{\beta(2m+3,2m+s+1)\right\}^{\left\{(s-1)s^{(2m+3)/(s-1)}\right\}^{-1}} \circ \langle v_{s-1}^{(s)} \rangle \langle s|. }$$

The first moments are the same for the respective approximate and exact distributions. The other (three) approximate and exact moments tend to be the same in each case when m and n are large.

7. Approximations to the distributions of $U_1^{(s)}$ and $V_1^{(s)}$. From the study of the first moment of $U_1^{(s)}$ as well as $V_1^{(s)}$ and their second moments in particular cases in section 5, one might suggest the following approximations:

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These again have their first mement from the respective approximate and exact distributions the same.

8. Accuracy of the approximations. For a comparison of the exact and approximate moments, first consider $U_{i,m,n}^{(s)}$ with i=s-1. While it may be noted that the first moment of the approximate distribution is the same as that of the exact given in (3.3), the next three moments from the approximate distribution are as follows:

(8.1)
$$(App.) \mu_2 \left\{ U_{s-1,m,n}^{(s)} \right\} = s^2 \frac{2(s-1)}{\pi} \frac{(2m+1+2)}{(2n-s+1)},$$

(8.2)
$$(App.) \mu_3 \left\{ U_{s-1,m,n}^{(s)} \right\} = s^3 \frac{3(s-1)}{\pi} \frac{(2m+1+2)}{(2n-2s+1+1)}$$

(8.3) (App.)
$$\mu_{\downarrow}^{*} \left\{ U_{s-1,m,n}^{(s)} \right\} = s^{\downarrow} \frac{\mu_{(s-1)}}{\pi} \frac{(2m+1+2)}{(2n-3s+1+2)}$$

For large n, the excess over unity of the ratio of approximate to exact moments in the case of second, third and fourth moments respectively can be

shown to be approximately

(8.4)
$$\delta_2 = ((s-3)(s-1) + 2)/2(m+1), \delta_3 = 3\delta_2 \text{ and } \delta_4 = 6\delta_2$$
.

When s=3, n=70 and m=5, the exact first four moments are respectively 0.02766, 0.00009381, 0.00003844 and 0.000001878 and $\delta_2 = 0.11$, $\delta_3 = 0.33$ and $\delta_4 = 0.72$. These values of δ 's are smaller than the corresponding values in (8.4). Where s=3, the approximate distribution could be used with reasonable accuracy for values of m, $n \geq 50$. For s=4, similar accuracy can be attained when m, $n \geq 100$ and for s=5 when m, $n \geq 150$. It may, however, be pointed out that the value of m could be smaller than indicated above to give adequate accuracy provided m is fairly large, and in most of the applications m is connected with sample sizes and, therefore, is often large. For s=2, $U_{1,m,n}^{(s)}$ is the sum of the roots and this case has been treated extensively in $\{2\}$, $\{4\}$ and $\{7\}$.

The approximate distribution $\sqrt[4]{V_{1,m,n}^{(s)}}$ gives better accuracy than discussed above when i=s-l for the reason that the linear factors in the moments involving min tend to make the moments from the approximate and exact distributions closer than in the case of $U_{1,m,n}^{(s)}$ for larger values of m and n. In general, this is true of other values of i also.

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REFERENCES

- (1) T. A. MIJARES, ''The moments of elementary symmetric functions of the roots of a matrix in multivariate analysis,'' Ann. Math. Stat., Vol. 32 (1961), pp. 1152-1160.
- (2) K. C. S. PILLAI, "'On some distribution problems in multivariete analysis," Mimeograph Series No. 88, Institute of Statistics, University of North Carolina, 1954.
- (3) K. C. S. PILLAL; ''Some new test criteria in multivariate analysis,''
 Ann. Math. Stat., Vol. 26 (1955) pp. 117-121.
- [4] K. C. S. FILLAI, "Some results useful in multivariate analysis,"

 Ann. Math. Stat., Vol. 27 (1956), pp. 1106-1114.
- (5) K. C. S. PILLAI, Statistical Tables for Tests of Multiveriate Hypotheses, The Statistical Center, University of the Philippines, Manila, 1960.
- [6] K. C. S. PILIAI and T. A. MIJARES, ''On the moments of the trace of a matrix and approximations to its distribution,'' Ann. Math. Stat. Vol. 30 (1959)., pp. 1135-1140.
- [7] K. C. S. PILIAI and PABLO SAMSON, Jr., ''On Hotelling's generalization of T2,'' Biometrika, Vol. 46 (1959), pp. 160-168.
- [8] S. N. ROY, ''p-statistics or some generalizations in analysis of variance appropriate to multivariate problems,'' Sankhya vol. 4 (1939), pp. 381-396.



ON THE MOMENTS OF ELEMENTARY SYMMETRIC FUNCTIONS OF THE ROOTS OF TWO MATRICES

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- 1. Summary. A lemma is given first which provides an easy method of expressing the product of an sth order Vandermonde type determinant and the kth and lth $(k, l \ge 0)$ powers of the rth and hth $(r, h \le s)$ elementary symmetric functions (esf's) respectively as a linear compound of determinants. The lemma extends itself readily to the product of powers of any number of esf's up to the sth. Using this lemma and some reduction formulae for certain special types of Vandermonde type determinants, a second lemma has been proved to show that certain formulas for the moments of esf's in s non-null characteristic roots $\lambda_i(0 < \lambda_1 \le \lambda_2 \le \cdots \le \lambda_s < \infty)$ of a matrix can be easily derived from corresponding formulas for the moments of corresponding esf's in s non-null roots, θ_i , $(0 < \theta_1 \le \cdots \le \theta_s < 1)$ of another matrix and vice versa. Illustrations are given explaining both lemmas.
- 2. Introduction. Many of the distribution problems in multivariate analysis are based on the distribution of the non-null characteristic roots of a matrix derived from sample observations taken from multivariate normal populations. This distribution given by Fisher [1], Girshick [2], Hsu [3] and Roy [12] is of the form

$$(2.1) f(\theta_1, \theta_2, \dots, \theta_s) = C(s, m, n) \prod_{i=1}^s \theta_i^m (1 - \theta_i)^n \prod_{i>j} (\theta_i - \theta_j) \\ 0 < \theta_1 \leq \dots \leq \theta_s < 1,$$

where

(2.2)
$$C(s, m, n) = \pi^{s/2} \prod_{i=1}^{s} \Gamma[(2m + 2n + s + i + 2)/2]/ \{\Gamma[(2m + i + 1)/2]\Gamma[(2n + i + 1)/2]\Gamma(i/2)\}$$

and m and n are defined differently for various situations described in [7], [9]. Now, if $\lambda_i = \theta_i/(1-\theta_i)$, $(i=1,2,\cdots,s)$, the joint distribution of the λ 's is obtained from (2.1) as

$$(2.3) f_1(\lambda_1, \lambda_2, \dots, \lambda_s) = C(s, m, n) \left[\prod_{i=1}^s \lambda_i^m / (1 + \lambda_i)^{m+n+s+1} \right] \prod_{i>j} (\lambda_i - \lambda_j)$$

$$0 < \lambda_1 \leq \dots \leq \lambda_s < \infty.$$

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The studies on the first esf in θ 's as well as the λ 's have been carried out by Pillai [6], [8], [9], Pillai and Mijares [10] and Pillai and Samson [11]. Mijares [5] has carried out some studies on esf's in general. In this paper, a lemma is proved which enables one to write down easily the moments of $U_{i,m,n}^{(s)}$ from the respective moments of $V_{i,m,n}^{(s)}$ and vice versa, where $U_{i,m,n}^{(s)}$ and $V_{i,m,n}^{(s)}$ denote the *i*th esf's in the s λ 's and s θ 's respectively. But first, a lemma is given (see next section) on which will be based the proof of the main lemma showing the easy derivation of the moments of $U_{i,m,n}^{(s)}$ from the respective moments of $V_{i,m,n}^{(s)}$.

3. Product of a Vandermonde determinant and powers of esf's. In this section we introduce the following lemma:

LEMMA 1. Let $D(g_s, g_{s-1}, \dots, g_1)$, $(g_j \ge 0, j = 1, 2, \dots, s)$, denote the determinant

(3.1)
$$D(g_s, g_{s-1}, \cdots, g_1) = \begin{vmatrix} x_s^{g_s} & x_s^{g_{s-1}} & x_s^{g_1} \\ \vdots & & & \\ x_1^{g_s} & x_1^{g_{s-1}} & x_1^{g_1} \end{vmatrix}.$$

If $a_r(r \leq s)$ denotes the rth esf in s x's, then (i)

$$(3.2) a_{\tau}D(g_{s}, g_{s-1}, \cdots, g_{1}) = \sum' D(g'_{s}, g'_{s-1}, \cdots, g'_{1}),$$

where $g'_j = g_j + \delta$, $j = 1, 2, \dots, s$, $\delta = 0, 1$ and \sum' denotes the sum over the $\binom{s}{r}$ combinations of s g's taken r at a time for which r indices $g'_j = g_j + 1$ such that $\delta = 1$ while for other indices $g'_j = g_j$ such that $\delta = 0$.

$$(3.3) a_r a_h D(g_s, g_{s-1}, \dots, g_1) = \sum_{s=1}^{n} D(g_s'', g_{s-1}'', \dots, g_1''),$$

where $h \leq s$, $g''_i = g''_i + \delta$, $j = 1, 2, \dots, s$, $\delta = 0, 1$ and \sum'' denotes summation over the $\binom{s}{r}\binom{s}{k}$ terms obtained by taking h at a time of the s g's in each D in \sum' in (3.2) for which h indices $g''_i = g'_j + 1$ while for other indices $g''_i = g'_j$.

(iii) $(a_r)^k(a_h)^l D(g_s, g_{s-1}, \dots, g_1)$, $(k, l \ge 0)$ can be expressed as a sum of $\binom{s}{r}^k\binom{s}{h}^l$ determinants obtained by performing on $D(g_s, g_{s-1}, \dots, g_1)$ in any order (i) k times and (i) l times with r = h.

However, if at least two of the indices in any determinant are equal, the corresponding term in the summation vanishes.

Before indicating a proof of the lemma, let us consider an illustration. Let us note first that [4]

$$(3.4) \quad a_w = \sum_{i=1}^{w} (-1)^{w+\sum_{i=1}^{w} p_i} [s_1^{p_1} s_2^{p_2} \cdots s_w^{p_w} / (1^{p_1} 2^{p_2} \cdots w^{p_w} p_1 ! p_2 ! \cdots p_w !)],$$

where \sum extends over all non-negative values of p_1, \dots, p_w such that $p_1 + 2p_2 + \dots + wp_w = w$, and $s_k = \sum_{j=1}^s x_j^k$. Also note that if we multiply the right side of (3.1) by e^{ts_u} , differentiate with respect to t once and put t = 0 we get,

$$(3.5) \quad s_u D(g_s, g_{s-1}, \dots, g_1) = \sum_{j=1}^s D(g_s, g_{s-1}, \dots, g_{j+1}, g_j + u, g_{j-1}, \dots, g_1).$$

Now consider the special case, w = 4. We get from (3.4)

$$(3.6) a_4 = s_1^4/4! + s_2^2/8 - s_2s_1^2/4 + s_1s_3/3 - s_4/4.$$

Using the right side of (3.6) and by repeatedly applying (3.5) with varying values of u (from 1 to 4) we get

$$a_4 D(g_s, g_{s-1}, \cdots, g_1)$$

$$= b_0 \sum_{j=1}^s D(g_s, g_{s-1}, \cdots, g_{j+1}, g_j + 4, g_{j-1}, \cdots, g_1)$$

$$+ b_1 \sum_{j \neq j' = 1}^s D(g_s, g_{s-1}, \cdots, g_j + 3, g_{j-1}, \cdots, g_{j'} + 1, \cdots, g_1)$$

$$+ b_2 \sum_{j \neq j' \neq j'' = 1}^s D(g_s, g_{s-1}, \cdots, g_j + 2, \cdots, g_{j'} + 1, \cdots, g_{j''} + 1, \cdots, g_1)$$

$$+ b_3 \sum_{j \neq j'' = 1}^s D(g_s, g_{s-1}, \cdots, g_j + 2, \cdots, g_{j'} + 2, \cdots, g_1)$$

$$+ b_4 \sum_{j \neq j'' \neq j'' \neq j''' \neq j''' = 1}^s D(g_s, g_{s-1}, \cdots, g_j + 1, \cdots, g_{j'} + 1, \cdots, g_{j''} + 1, \cdots, g_{j'''} + 1, \cdots, g_{j''} + 1, \cdots, g_{j''} + 1, \cdots, g_{j'''} + 1,$$

and where the indices j, j', j'', j''' are the only ones which have been increased. Now since in the last sum on the right side of (3.7) there are only $\binom{s}{4}$ distinguishable terms, it is obvious that $a_4 D(g_s, g_{s-1}, \dots, g_1)$ is obtained from (3.7) as a sum of $\binom{s}{4}$ determinants whose indices are obtained by selecting 4 out of s g's at a time and increasing each of the 4 selected g's by unity.

Now consider the general case (i). Apply (3.4) to (3.1) with w = r. We can show that the coefficient of the determinant of a specified set of indices obtained in this operation such that at least one g on the left side of (3.2) has been increased by more than unity, is equal to zero. For instance, the coefficient of the determinant with one g index increased by r - j > 1 while any other j g's increased each by unity is given by

(3.8)
$$(1/j!) \sum_{i=1}^{r-j+2} (-1)^{r-j+2} [1^{p_1} 2^{p_2} \cdots (r-j)^{p_{r-j}} p_1! p_2! \cdots p_{r-j}!]^{-1} = 0,$$

where $p_1 + 2p_2 + \cdots + (r-j)p_{r-j} = r-j.$

In a similar manner, coefficients of all other determinants with at least one index increased by more than unity can be shown to be equal to zero. There remain, therefore, only determinants in which r out of s indices have been increased just by unity. It may be observed that this last set of determinants is obtained from the term $s_1^r/r!$ in (3.4), (w=r), while all other terms arise from more than one term in (3.4), (w=r), and their coefficients are obtained as sums of positive and negative values where each sum (coefficient) equals zero.

Now it may be seen that the truth of (ii) in Lemma 1 can be observed easily by an application of (3.4) to the right side of (3.2) with w = h.

Similarly (iii) further follows easily by repeated application of (i), as stated in the lemma, k + l times, k times using (3.4), (w = r), and l times using (3.4) with w = h. In addition, it may be pointed out that the method of proof extends itself to generalize (iii) further to include powers of any number of esf's up to the sth.

4. Derivation of moments of $U_{i,m,n}^{(s)}$ from those of $V_{i,m,n}^{(s)}$. In this section we prove the main lemma, stated below. Let

$$(4.1) V(m+s-1+q_s, \dots, m+q_1; n)$$

$$= \begin{vmatrix} \int_0^1 \theta_s^{m+s-1+q_s} (1-\theta_s)^n d\theta_s & \dots & \int_0^1 \theta_s^{m+q_1} (1-\theta_s)^n d\theta_s \\ & \dots & \\ \int_0^{\theta_2} \theta_1^{m+s-1+q_s} (1-\theta_1)^n d\theta_1 & \dots & \int_0^{\theta_2} \theta_1^{m+q_1} (1-\theta_1)^n d\theta_1 \end{vmatrix}$$

and let

$$(4.2) U(m+s-1+q_s, \cdots, m+q_1; p) = \begin{vmatrix} \int_0^\infty \frac{\lambda_s^{m+s-1+q_s}}{(1+\lambda_s)^p} d\lambda_s \cdots \int_0^\infty \frac{\lambda_s^{m+q_1}}{(1+\lambda_s)^p} d\lambda_s \\ & \cdots \\ \int_0^{\lambda_2} \frac{\lambda_1^{m+s-1+q_s}}{(1+\lambda_1)^p} d\lambda_1 \cdots \int_0^{\lambda_2} \frac{\lambda_1^{m+q_1}}{(1+\lambda_1)^p} d\lambda_1 \end{vmatrix}$$

$$q_j \ge 0 j = 1, 2, \cdots, s, \text{ and } p = m+n+s+1.$$

Now, from Lemma 1 and (2.1), the kth moment $\mu'_k\{V^{(s)}_{i,m,n}\}$, of $V^{(s)}_{i,m,n}$, can be expressed as a linear compound of determinants of the V type in (4.1) where q_s , q_{s-1} , \cdots , q_1 may take different sets of values in different terms. Further, the coefficients of the linear compound would involve as a common factor C(s, m, n) but otherwise would be independent of m and n.

Similarly, $\mu'_{k}\{U_{i,m,n}^{(s)}\}$ can be shown to be a linear compound of the determinants of the *U*-type in (4.2) with the coefficients of corresponding terms in this com-

pound the same as in the previous compound, the correspondence of terms being marked by the equality of the vector $(q_s, q_{s-1}, \dots, q_1)$ in the two compounds.

Now we state the lemma.

LEMMA 2. $\mu'_{k}\{U_{i,m,n}^{(s)}\}\$ is derivable from $\mu'_{k}\{V_{i,m,n}^{(s)}\}\$ by making the following changes in the expression for the latter (obtained by evaluating the linear compound of V-type determinants): (a) Multiply by -1 all terms except the term in n in each linear factor involving n and (b) change n to m+n+s+1 after performing (a).

Before proving the lemma let us illustrate it by a couple of special cases. Using (i) of Lemma 1 we get

(4.3)
$$\mu_1'\{V_{i,m,n}^{(s)}\} = C(s, m, n)V(m+s, m+s-1, \cdots, m+s-i+1, m+s-i-1, \cdots, m+1, m; n).$$

The right side of (4.3) can be shown to be equal to

(4.4)
$${s \choose i} \prod_{j=1}^{i} \frac{(2m+s-j+2)}{(2m+2n+2s-j+3)}$$

based on some particular cases of determinants evaluated in [10]. From this result using Lemma 2, the first moment of the *i*th esf in the λ 's is given by

(4.5)
$$\mu_1' \left\{ U_{i,m,n}^{(s)} \right\} = {s \choose i} \prod_{j=1}^i \frac{(2m+s-j+2)}{(2n+j-1)}.$$

Now consider $\mu_2'\{V_{2,m,n}^{(s)}\}$. Using (ii) of Lemma 1 with h=r we get

$$\mu_2' \{ V_{2,m,n}^{(s)} \}$$

$$= C(s, m, n)\{V(m + s + 1, m + s, m + s - 3, \dots, m + 1, m; n)$$

$$(4.6) + V(m+s+1, m+s-1, m+s-2, m+s-4, \cdots, m+1, m; n) + V(m+s, m+s-1, m+s-2, m+s-3, m+s-5, \cdots, m+1, m; n)\},$$

....

Now substituting the values of the determinants [10] in (4.6)

(4.7)
$$\mu_2' \left\{ V_{2,m,n}^{(s)} \right\} = \frac{s(s-1)(2m+s)(2m+s+1)}{3! \prod_{j=1}^6 (2m+2n+2s-j+5)} G_1,$$

where

$$G_{1} = 6n^{2} \{4s(s-1)m^{2} + 2(s-1)(2s^{2} + s + 8)m + s^{4} + 7s^{2} - 8s + 12\}$$

$$+ 3n \{16s(s-1)m^{3} + 4(s-1)(8s^{2} + 5s + 8)m^{2} + 2(s-1)$$

$$\cdot (10s^{3} + 12s^{2} + 27s + 24)m + 4s^{5} + 3s^{4} + 12s^{3} + 5s^{2} - 24s + 36\}$$

$$+ s(s+1)(2m+s+1)(2m+s+2)(m+s)(2m+2s+1)$$

$$+ (s-2)(2m+2s+3)(2m+s-1)$$

$$\cdot \{4sm^{2} + 2s(3s+2)m + 2s^{3} + 3s^{2} + s + 6\}.$$

Using Lemma 2 we get from (4.7),

(4.8)
$$\mu_2' \left\{ U_{2,m,n}^{(s)} \right\} = \frac{s(s-1)}{3!} \frac{(2m+s)(2m+s+1)}{\prod_{j=1}^6 (2n+j-3)} G,$$

where G is obtained from G_1 by attaching negative sign to the first degree term in n and then changing n to m + n + s + 1 in all the terms.

Proof. Apply Theorem 3 of [8] to the V-determinant in (4.1). We get

$$(4.9) V(m+s-1+q_s, \cdots, m+q_1; n) = (m+s+q_s+n)^{-1}(B^{(s)}+(m+s-1+q_s)C^{(s)}),$$

where

$$B^{(s)} = 2\sum_{j=s-1}^{1} (-1)^{s-j-1} V(2m+s+j-2+q_s+q_j; 2n+1)$$

$$(4.10) \qquad \times V(m+s-2+q_{s-1}, \cdots, m+j+q_{j+1}, \cdots, m+j-2+q_{j+1}, \cdots, m+q_1; n)$$

and

$$(4.11) \quad C^{(s)} = V(m+s-2+q_s, m+s-2+q_{s-1}, \cdots, m+q_1; n).$$

Again, applying Theorem 4 of [8] to the U-determinant in (4.2) we get

$$(4.12) U(m+s-1+q_s, \cdots, m+q_1; p) = [p-(m+s+q_s)]^{-1}(Q^{(s)}+(m+s-1+q_s)R^{(s)}),$$

where

$$Q^{(s)} = 2\sum_{j=s-1}^{1} (-1)^{s-j-1} U(2m+s+j-2+q_s+q_j; 2p-1)$$

$$\times U(m+s-2+q_{s-1}, \cdots, m+j+q_{j+1},$$

$$m+j-2+q_{j-1}, \cdots, m+q_1; p)$$

and

$$(4.14) \quad R^{(s)} = U(m+s-2+q_s, m+s-2+q_{s-1}, \cdots, m+q_1; p).$$

First, it may be observed that the factor $m + s + q_s + n$ in (4.9) becomes the factor $p - (m + s + q_s)$ in (4.12) by changes (a) and (b) of the lemma. Further, repeated application of Theorem 3 of [8] to the right side of (4.9) would reduce it to a linear compound of terms each of which is a product of s/2 simple beta functions of Type I (V-type) if s is even and (s + 1)/2 beta functions if s is odd. The coefficients of this linear compound would involve products of functions of m and n of the type $(m + j + q_j + n)^{-1}$ and the type $(m + j - 1 + q_j)$ as in (4.9). Similarly, repeated application of Theorem 4 of [8] to the right side of (4.12) would reduce it to a linear compound as above with the

exception that simple beta functions involved will be of Type II (U-type) instead of Type I and $[p - (m + j + q_j)]^{-1}$ will replace $(m + j + q_j + n)^{-1}$. Now it may be observed that changes (a) and (b) of the lemma will make the corresponding coefficients the same in the two linear compounds which are obtained after repeated applications of Theorems 3 and 4 of [8] to (4.9) and (4.12) respectively. It remains, therefore, to show that C(s, m, n) times each term of the linear compound involving products of beta functions of Type I reduces to C(s, m, n) times the corresponding term in the second linear compound involving the product of beta functions of Type II using (a) and (b) of the lemma. Now note that, if s is even,

$$C(s, m, n) = 2^{-s(s+6)/8}$$

$$(4.15) \times \frac{\prod_{i=1}^{s/2} \Gamma(2m+2n+s+2i+1)}{\prod_{i=1}^{s/2} \{\Gamma(2m+2i)\Gamma(2n+2i)\Gamma(i)\}(1\cdot3)(1\cdot3\cdot5)\cdots(1\cdot3\cdot5\cdots(s-3))}$$

and if s is odd

$$C(s, m, n) = 2^{-(s-1)(s+5)/8}$$

(4.16)
$$\times \frac{\prod_{i=1}^{(s-1)/2} \Gamma(2m+2n+s+2i+1)\Gamma(m+n+s+1)}{\prod_{i=1}^{(s-1)/2} \{\Gamma(2m+2i)\Gamma(2n+2i)\Gamma(i)\}\Gamma[(2m+s+1)/2]\Gamma[(2n+s+1)$$

Now, for s even, consider the term

Substitute in (4.17) the value of C(s, m, n) from (4.15) and those of the Type I beta functions and we get

$$g(m,s)[(2m+2n+2s+q_s+q_{s-1}-1)\cdots(2m+2n+2s+1)]$$

$$\cdot[(2m+2n+2s+q_{s-2}+q_{s-3}-3)\cdots(2m+2n+2s-1)]$$

$$(4.18) \cdots[(2m+2n+q_2+q_1+3)\cdots(2m+2n+s+3)]$$

$$\cdot[(2n+3)(2n+2)][(2n+5)\cdots(2n+2)]$$

$$\cdots[(2n+s-1)\cdots(2n+2)],$$

where g(m, s) is a function of m and s.

Similarly consider

$$(4.19) C(s, m, n)U(2m + 2s - 3 + q_s + q_{s-1}; 2p - 1)U(2m + 2s - 7 + q_{s-2} + q_{s-3}; 2p - 1) \cdots U(2m + 1 + q_2 + q_1; 2p - 1).$$

After substitution of values of C(s, m, n) and U's in (4.19) we get

$$g(m,s)[(2n-q_{s-1}+3)\cdots(2n+1)][(2n-q_{s-2}-q_{s-3}+7)\cdots(2n+3)]\cdots[(2n+2s-q_{2}-q_{1}-1)\cdots(2n+s-1)]$$

$$(4.20)$$

$$(4.20)$$

$$(2m+2n+s+3)\cdots(2m+2n+2s)][(2m+2n+s+5)\cdots(2m+2n+2s)].$$

$$(2m+2n+2s)]\cdots[(2m+2n+2s-1)(2m+2n+2s)].$$

Now it may be noted that (4.20) can be obtained from (4.18) by (a) and (b) of the lemma. In a similar manner, when s is even, other corresponding terms of the linear compounds in the two cases can be shown to satisfy (a) and (b) of the lemma.

If s is odd, we may consider the terms like

(4.21)
$$C(s, m, n)V(2m + 2s - 3 + q_s + q_{s-1}; 2n + 1) \\ \cdots V(2m + 3 + q_3 + q_2; 2n + 1)V(m + q_1; n).$$

Using (4.16) and the values of the V's and performing (a) and (b) in (4.21) we will get

$$C(s, m, n)U(2m + 2s - 3 + q_s + q_{s-1}; 2p - 1)$$

 $\cdots U(2m + 3 + q_3 + q_2; 2p - 1)U(m + q_1; p).$

Similarly, if s is odd, we can show that other corresponding terms of the linear compounds in the two cases satisfy (a) and (b).

Hence the lemma.

It may be noted that $\mu'_{r}\{V_{i,m,n}\}$ may similarly be derived from $\mu'_{r}\{U_{i,m,n}\}$ by inverse operations of (b) and (a) of Lemma 2. Further, Lemma 2 readily extends to the case of product moments (say of the rth and hth esf's) in view of (ii) of Lemma 1.

REFERENCES

- [1] Fisher, R. A. (1939). The sampling distribution of some statistics obtained from non-linear equations. Ann. Eugenics 9 238-249.
- [2] GIRSHICK, M. A. (1939). On the sampling theory of roots of determinantal equations.

 Ann. Eugenics 10 203-224.
- [3] Hsu, P. L. (1939). On the distribution of the roots of certain determinantal equations.

 Ann. Eugenics 9 250-258.
- [4] MacMahon, P. A. (1960). Combinatory Analysis. Chelsea, New York.
- [5] MIJARES, T. A. (1961). The moments of elementary symmetric functions of the roots of a matrix in multivariate analysis. Ann. Math. Statist. 32 1152-1160.
- [6] PILLAI, K. C. S. (1954). On some distribution problems in multivariate analysis.

 Mimeograph Ser. No. 88, Institute of Statistics, Univ. of North Carolina.

- [7] PILLAI, K. C. S. (1955). Some new test criteria in multivariate analysis. Ann. Math. Statist. 26 117-121.
- [8] PILLAI, K. C. S. (1956). Some results useful in multivariate analysis. Ann. Math. Statist. 27 1106-1114.
- [9] PILLAI, K. C. S. (1960). Statistical Tables for Tests of Multivariate Hypotheses. The Statistical Center, Manila.
- [10] PILLAI, K. C. S. and MIJARES, T. A. (1959). On the moments of the trace of a matrix and approximations to its distribution. Ann. Math. Statist. 30 1135-1140.
- [11] PILLAI, K. C. S. and Samson, Pablo, Jr. (1959). On Hotelling's generalization of T².

 Biometrika 46 160-168.
- [12] Roy, S. N. (1939). p-statistics or some generalizations in analysis of variance appropriate to multivariate problems. Sankhyā 4 381-396.