

7. THE LAW OF LARGE NUMBERS AND THE CENTRAL LIMIT THEOREM

The law of large numbers and the central limit theorem play very important roles in statistics.

7.1. The Weak Law of Large Numbers (WLLN). Two probability inequalities are needed in the proof of WLLN.

Theorem 7.1. (*Markov's Inequality*) Let X be a non-negative random variable and $\mathbb{E}(X) < \infty$. For any $t > 0$,

$$\mathbb{P}(X \geq t) \leq \frac{\mathbb{E}(X)}{t}.$$

Proof. We only prove for the continuous case.

$$\mathbb{E}(X) = \int_0^\infty xf(x)dx = \int_0^t xf(x)dx + \int_t^\infty xf(x)dx \geq \int_t^\infty xf(x)dx \geq t\mathbb{P}(X > t).$$

□

We can easily prove the result using the concept of indicator function. Note that $I(X \geq t) \leq (X/t)$ is always true, and the result is obtained by taking expectation on both sides.

Theorem 7.2. (*Chebyshev's Inequality*) Let X be a random variable with finite mean $\mu = \mathbb{E}(X)$ and variance $\sigma^2 = \mathbb{V}(X)$. Then for $t > 0$ we have

$$\mathbb{P}(|X - \mu| \geq t) \leq \frac{\sigma^2}{t^2}.$$

Proof. We apply Markov's inequality to the random variable $|X - \mu|^2$:

$$\mathbb{P}(|X - \mu| \geq t) = \mathbb{P}(|X - \mu|^2 \geq t^2) \leq \frac{\mathbb{E}(|X - \mu|^2)}{t^2} = \frac{\sigma^2}{t^2}.$$

□

Similarly we can use indicator function to prove Chebyshev's inequality by taking expectation over both sides of $I(|X - \mu| \geq t) \leq \frac{|X - \mu|^2}{t^2}$.

Theorem 7.3. (*The Weak Law of Large Numbers (WLLN)*) If X_1, X_2, \dots is a sequence of iid random variables with finite expectation μ , then for any $\epsilon > 0$ we have

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\left| \frac{1}{n} \sum_{i=1}^n X_i - \mu \right| \geq \epsilon \right) = 0.$$

Proof. We only prove a simple case where X_1, X_2, \dots have finite variance $\sigma^2 < \infty$. Let $\bar{X}_n = \sum_{i=1}^n X_i/n$ and we have $\mathbb{E}(\bar{X}_n) = \mu$ and $\mathbb{V}(\bar{X}_n) = \sigma^2/n$. Using Chebyshev's inequality,

$$0 \leq \lim_{n \rightarrow \infty} \mathbb{P} \left(\left| \frac{1}{n} \sum_{i=1}^n X_i - \mu \right| \geq \epsilon \right) \leq \lim_{n \rightarrow \infty} \frac{\sigma^2}{n\epsilon^2} = 0.$$

□

The WLLN says that the distribution of the average of iid random variables \bar{X}_n will be more and more concentrated on its mean μ as n gets large. In other words, the observed sample mean $\sum_{i=1}^n x_i/n$ will be a very good approximation to μ as n gets large.

7.2. The Central Limit Theorem (CLT).

Theorem 7.4. (The Central Limit Theorem) Let X_1, X_2, \dots be a sequence of iid random variables with finite mean μ and variance σ^2 . Define

$$Z_n = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}$$

where $\bar{X}_n = \sum_{i=1}^n X_i/n$, then the distribution of Z_n converges to a standard normal distribution as $n \rightarrow \infty$. What it means is that, for any z we have

$$\lim_{n \rightarrow \infty} \mathbb{P}(Z_n \leq z) = \Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx.$$

The CLT says that the mean of a random sample X_1, \dots, X_n from **any** distribution can be approximated by a normal distribution as long as μ and σ^2 are finite and n gets large. It is a little bit surprising and has tremendous applications.

Example 7.5. Let $X_i \sim \text{Bernoulli}(p)$ and $\bar{X}_n = \sum_{i=1}^n X_i/n$. Notice that the exact distribution of $n\bar{X}_n$ is $\text{Binomial}(n, p)$. In Figure 7.1 we illustrate that the distribution of $n\bar{X}_n$ becomes increasingly similar to a normal pdf as $n \rightarrow \infty$, which verifies the CLT.

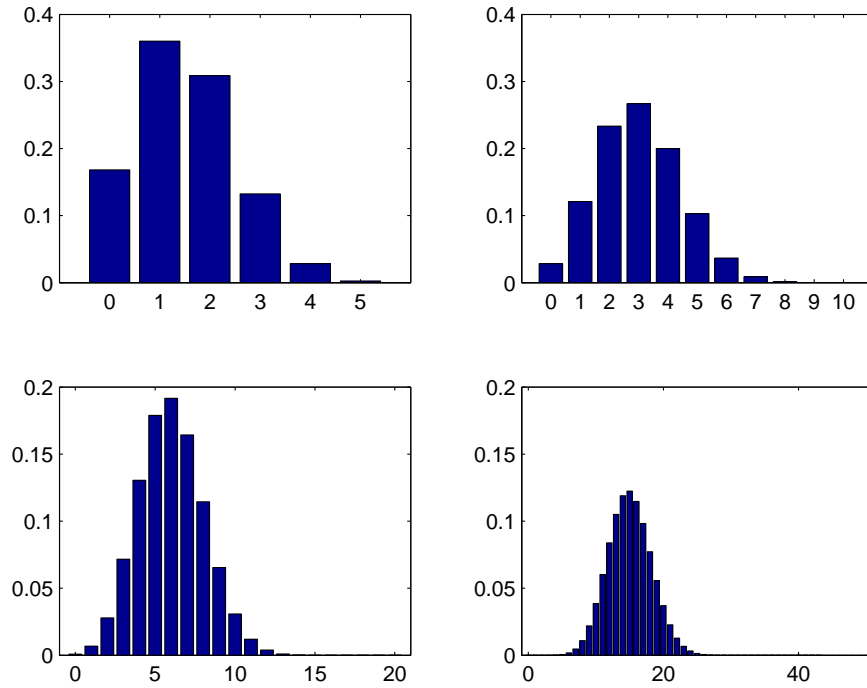


FIGURE 7.1. The distribution of $\text{Binomial}(n, p = 0.3)$ for $n = 5, 10, 20$ and 50 .

Example 7.6. (Example 7.8 in [1]) Let Y_i be the service time for the i th customer, and we have $\mu = 1.5$, $\sigma^2 = 1$. So $\mathbb{P}(\sum_{i=1}^{100} Y_i \leq 120) = \mathbb{P}(\bar{Y} \leq 1.2)$. Since $n = 100$ is large, CLT tells that \bar{Y} is approximately $N(\mu, \sigma^2/n) = N(1.5, 0.01)$. So $\mathbb{P}(\bar{Y} \leq 1.2) = \mathbb{P}(\frac{\bar{Y}-1.5}{0.1} \leq \frac{1.2-1.5}{0.1}) \approx \mathbb{P}(Z \leq -3) = \Phi(-3) = 0.00135$. Here Z is a standard normal and $\Phi(z)$ can be found from Table 4, Appendix III in [1].